

Article

Cadmium-Induced Tubular Dysfunction in Type 2 Diabetes: A Population-Based Cross-Sectional Study

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Abstract: The global prevalence of diabetes, and its major complication, diabetic nephropathy, have reached epidemic proportions. The toxic metal cadmium (Cd) also induces nephropathy, indicated by a sustained reduction in the estimated glomerular filtration rate (eGFR) and the excretion of β_2 -microglobulin (β_2M) above 300 $\mu\text{g}/\text{day}$, which reflects kidney tubular dysfunction. However, little is known about the nephrotoxicity of Cd in the diabetic population. Here, we compared Cd exposure, eGFR, and tubular dysfunction in both diabetics ($n = 81$) and non-diabetics ($n = 593$) who were residents in low- and high-Cd exposure areas of Thailand. We normalized the Cd and β_2M excretion rates (E_{Cd} and E_{β_2M}) to creatinine clearance (C_{cr}) as E_{Cd}/C_{cr} and E_{β_2M}/C_{cr} . Tubular dysfunction and a reduced eGFR were, respectively, 8.7-fold ($p < 0.001$) and 3-fold ($p = 0.012$) more prevalent in the diabetic than the non-diabetic groups. The doubling of E_{Cd}/C_{cr} increased the prevalence odds ratios for a reduced eGFR and tubular dysfunction by 50% ($p < 0.001$) and 15% ($p = 0.002$), respectively. In a regression model analysis of diabetics from the low-exposure locality, E_{β_2M}/C_{cr} was associated with E_{Cd}/C_{cr} ($\beta = 0.375$, $p = 0.001$) and obesity ($\beta = 0.273$, $p = 0.015$). In the non-diabetic group, E_{β_2M}/C_{cr} was associated with age ($\beta = 0.458$, $p < 0.001$) and E_{Cd}/C_{cr} ($\beta = 0.269$, $p < 0.001$). However, after adjustment for age, and body mass index (BMI), E_{β_2M}/C_{cr} was higher in the diabetics than non-diabetics of similar E_{Cd}/C_{cr} ranges. Thus, tubular dysfunction was more severe in diabetics than non-diabetics of similar age, BMI, and Cd body burden.

Keywords: β_2 -microglobulin; cadmium; diabetes; diabetic nephropathy; GFR; tubular proteinuria



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1. Introduction

Type 2 diabetes is a metabolic disorder resulting in a rise of fasting plasma glucose ≥ 126 mg/dL. The worldwide prevalence of diabetes, which is often linked to obesity, has now reached epidemic proportions. However, it is increasingly apparent that exposure to various diabetogenic pollutants, such as cadmium (Cd), is an important risk factor [1–6]. Strong evidence supporting the diabetogenicity of Cd comes from the Wuhan–Zhuhai prospective cohort study. This study, which measured the fasting plasma glucose levels and urinary Cd over a 3-year period, reported that for each tenfold increase in urinary Cd, the prevalence of prediabetes increased by 42% [4].

Although Cd exposure has only recently been recognized as a risk factor for diabetes [5,6], the increased susceptibility to Cd-induced nephrotoxicity in people with diabetes was first noted in the 1990 Belgian population study (Cadmibel) [7]. Similar observations were then made in studies from Sweden [8,9], Australia [10], the U.S. [11], and Korea [12]. The defective tubular reabsorption of proteins, indicated by an increase in the excretion of proteins of low molecular weight, such as β_2 -microglobulin (β_2M), is the most frequently reported sign of Cd nephrotoxicity [13,14]. In early studies, the

effect of Cd on the glomerular filtration rate (GFR) involved workplace exposure, where workers were exposed mainly through the inhalation of relatively high levels of Cd in fumes and dust [15,16]. Evidence that low environmental Cd exposure through the diet also affects the GFR has now been reported in adult populations in many countries, including the U.S. [11,17–19], Thailand [20], Guatemala [21], Myanmar [22], Taiwan [23], and Korea [24,25].

Epidemiologic studies implicating low environmental exposure to Cd in the pathogenesis of tubulopathy and GFR deterioration are abundant. In comparison, studies on these Cd-induced pathologies in the diabetic population are limited. Therefore, the present study aimed to quantify Cd exposure levels, kidney tubular dysfunction, and the reduction of eGFR experienced by diabetics and non-diabetics who live in low- and high-Cd-exposure areas of Thailand.

We used the excretion of Cd as a measure of long-term exposure (body burden) and a sign of nephrotoxicity. The utility of Cd excretion as an indicator of its toxicity to kidneys is based on a study showing that Cd excreted in complexes with metallothionein (MT) emanated from injured or dying kidney tubular epithelial cells [26,27]. Tubular proteinuria was indicated by a rise of β_2 M excretion ≥ 300 $\mu\text{g/L}$ of filtrate. We employed established equations of the Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) to compute the estimated GFR (eGFR) [28].

For the accurate quantification of the kidney burden of Cd and its effects, we normalized the excretion of Cd and β_2 M (E_{Cd} and $E_{\beta_2\text{M}}$) to creatinine clearance (C_{cr}), denoted as $E_{\text{Cd}}/C_{\text{cr}}$ and $E_{\beta_2\text{M}}/C_{\text{cr}}$, respectively [29]. This C_{cr} normalization depicts an amount of a given chemical excreted per volume of filtrate, which is at least roughly related to the amount of the chemical excreted per nephron. In effect, C_{cr} normalization corrects for differences in the number of surviving nephrons among study subjects [29].

2. Materials and Methods

2.1. Cohort Participants

Participants came from three population-based studies undertaken in a Cd pollution area in the Mae Sot District, Tak Province ($n = 211$), two low-Cd-exposure locations in Bangkok ($n = 322$), and the Pakpoo municipality of Nakhon Si Thammarat Province ($n = 141$). The Institutional Ethical Committees of Chulalongkorn University, Chiang Mai University, and the Mae Sot Hospital approved the study protocol for the Mae Sot and Bangkok groups [30]. The Office of the Human Research Ethics Committee of Walailak University in Thailand approved the study protocol for the Pakpoo group [31]. All participants gave informed consent, and all had resided at their current addresses for at least 30 years. Exclusion criteria included pregnancy, breastfeeding, a history of metal work, and a hospital record or physician's diagnosis of an advanced chronic disease. Smoking, diabetes, hypertension, regular use of medications, educational level, occupation, and family health history were ascertained by questionnaire. Prediabetes and diabetes were indicated by fasting plasma glucose levels ≥ 110 and ≥ 126 mg/dL, respectively (<https://www.cdc.gov/diabetes/basics/getting-tested.html>) (accessed on 17 February 2023) or a physician's prescription of anti-diabetic medications. Hypertension was defined as systolic blood pressure ≥ 140 mmHg, diastolic blood pressure ≥ 90 mmHg, a physician's diagnosis, or prescription of anti-hypertensive medications.

As workplace exposure was an exclusion criterion, participants presumably acquired Cd from the diet and/or smoking. Based on the measured levels of Cd in duplicate diets [32] and a nationwide survey of Cd levels in soils and food crops [33], environmental exposure to Cd in Bangkok and Nakhon Si Thammarat were low, and the Mae Sot district was high. From a previous study conducted in the Mae Sot district, the Cd content of the paddy soil samples exceeded the standard of 0.15 mg/kg, and rice samples collected from households contained 4 times the amount of the permissible Cd level of 0.1 mg/kg [34]. A health survey reported that the prevalence of CKD in the Mae Sot was 16.1%, while the prevalence of tubular proteinuria was 36.1% [35].

2.2. Urine and Blood Sampling and Analysis

Second morning urine samples were collected after an overnight fast. Samples of whole blood were obtained within 3 h of urine sampling. Aliquots of urine, whole blood, and plasma were stored at $-20\text{ }^{\circ}\text{C}$ or $-80\text{ }^{\circ}\text{C}$ for later analysis. The assay for the urine and plasma concentrations of creatinine ($[\text{cr}]_{\text{u}}$ and $[\text{cr}]_{\text{p}}$) was based on the Jaffe reaction. The assay of $\beta_2\text{M}$ in the urine ($[\beta_2\text{M}]_{\text{u}}$) was based on the latex immunoagglutination method (LX test, Eiken 2MGII; Eiken and Shionogi Co., Tokyo, Japan).

For the Bangkok group, the urinary concentration of Cd ($[\text{Cd}]_{\text{u}}$) was determined by inductively coupled plasma mass spectrometry (ICP/MS, Agilent 7500, Agilent Technologies, Santa Clara, CA, USA). Multi-element standards (EM Science, EM Industries, Inc., Newark, NJ, USA) were used to calibrate the Cd analyses. The accuracy and precision of those analyses were ascertained with reference urine (Lyphochek[®], Bio-Rad, Sydney, Australia). The $[\text{Cd}]_{\text{u}}$ assigned to samples with Cd below the detection limit was $0.05\text{ }\mu\text{g/L}$ divided by the square root of 2 [36].

For the Pakpoo group, the $[\text{Cd}]_{\text{u}}$ was determined with the GBC System 5000 graphite furnace atomic absorption spectrophotometer (AAS) (GBC Scientific Equipment, Hampshire, IL, USA). Multielement standards were used to calibrate the metal analysis (Merck KGaA, Darmstadt, Germany). Reference urine levels 1, 2, and 3 (Lyphochek, Bio-Rad, Hercules, CA, USA) were used for quality control, analytical accuracy, and precision assurance. When a $[\text{Cd}]_{\text{u}}$ level was less than its detection limit, the concentration assigned was $0.1\text{ }\mu\text{g/L}$ divided by the square root of 2 [36].

For the Mae Sot group, the $[\text{Cd}]_{\text{u}}$ was determined with AAS (Shimadzu Model AA-6300, Kyoto, Japan). Urine standard reference material No. 2670 (National Institute of Standards, Washington, DC, USA) was used for quality assurance and control purposes. The limit of detection of the $[\text{Cd}]_{\text{u}}$ was $0.06\text{ }\mu\text{g/L}$. None of the urine samples contained a $[\text{Cd}]_{\text{u}}$ below the detection limit.

The comparability of the $[\text{Cd}]_{\text{u}}$ was ascertained by simultaneous quantification of Cd in the reference urine samples where the coefficient of variation was within acceptable clinical chemistry standards.

2.3. Normalization of E_{Cd} to C_{cr} and E_{cr}

E_x was normalized to C_{cr} as $E_x/C_{\text{cr}} = [x]_{\text{u}}[\text{cr}]_{\text{p}}/[\text{cr}]_{\text{u}}$, where $x = \text{Cd}$ or $\beta_2\text{M}$; $[x]_{\text{u}}$ = urine concentration of x (mass/volume); $[\text{cr}]_{\text{p}}$ = plasma creatinine concentration (mg/dL); and $[\text{cr}]_{\text{u}}$ = urine creatinine concentration (mg/dL). E_x/C_{cr} was expressed as an amount of x excreted per volume of filtrate [29].

E_x was normalized to E_{cr} as $[x]_{\text{u}}/[\text{cr}]_{\text{u}}$, where $x = \text{Cd}$ or $\beta_2\text{M}$; $[x]_{\text{u}}$ = urine concentration of x (mass/volume); and $[\text{cr}]_{\text{u}}$ = urine creatinine concentration (mg/dL). The ratio $[x]_{\text{u}}/[\text{cr}]_{\text{u}}$ was expressed in $\mu\text{g/g}$ of creatinine.

2.4. Estimated Glomerular Filtration Rate

The GFR is the product of the nephron number and the mean single nephron GFR, and, in theory, the GFR is indicative of nephron function [28,37]. In practice, the GFR is estimated from established CKD-EPI equations and reported as eGFR [28].

The male eGFR = $141 \times [\text{plasma creatinine}/0.9]^Y \times 0.993^{\text{age}}$, where $Y = -0.411$ if $[\text{cr}]_{\text{p}} \leq 0.9\text{ mg/dL}$ and $Y = -1.209$ if $[\text{cr}]_{\text{p}} > 0.9\text{ mg/dL}$. The female eGFR = $144 \times [\text{plasma creatinine}/0.7]^Y \times 0.993^{\text{age}}$, where $Y = -0.329$ if $[\text{cr}]_{\text{p}} \leq 0.7\text{ mg/dL}$ and $Y = -1.209$ if $[\text{cr}]_{\text{p}} > 0.7\text{ mg/dL}$.

2.5. Statistical Analysis

The data were analyzed with IBM SPSS Statistics 21 (IBM Inc., New York, NY, USA). We used the Mann–Whitney U-test to assess differences in the means between the two groups. The Pearson chi-squared test was used to assess differences in the percentages. To identify the departures of continuous variables from a normal distribution, the one-sample Kolmogorov–Smirnov test was used, and logarithmic transformation was applied

to variables that showed rightward skewing before they were subjected to parametric statistical analysis. Logistic regression analysis was used to determine the prevalence odds ratio (POR) for tubular proteinuria and reduced eGFR. Tubular proteinuria was defined as $E_{\beta_2M}/C_{cr} \geq 300 \mu\text{g/L}$ of filtrate. Reduced eGFR was assigned when $e\text{GFR} \leq 60 \text{ mL/min/1.73 m}^2$. Univariate/covariance analyses with Bonferroni correction in multiple comparisons were used to obtain the mean E_{β_2M}/C_{cr} adjusted for age and BMI, and interactions among groups of diabetics and non-diabetics stratified by three ranges of E_{Cd}/C_{cr} . For all tests, p -values ≤ 0.05 were considered to indicate statistical significance.

3. Results

3.1. Cohort Composition and Characteristics

The Thai cohort of 674 participants consisted of 463 drawn from two low-exposure localities, and 211 from an area where environmental Cd pollution is endemic (Table 1).

Table 1. Characterization of cohort participants.

Parameters	All Participants, $n = 674$	Diabetics, $n = 81$	Non-Diabetics, $n = 593$	p
Low-exposure locality (%)	67.1	86.4	64.4	<0.001
Female, %	57.6	74.1	55.3	<0.001
Smoking, %	31.9	14.8	34.2	<0.001
Hypertension ^c %	25.8	55.6	21.8	<0.001
Age, years	47.1 \pm 16.3	58.5 \pm 9.7	45.6 \pm 16.4	<0.001
BMI, kg/m^2	22.8 \pm 4.2	26.0 \pm 4.8	22.3 \pm 3.9	<0.001
Obese ^a (%)	3.9		2.4	<0.001
eGFR ^b , mL/min/1.73 m^2	87.4 \pm 23.1	76.2 \pm 21.0	89.0 \pm 23.0	<0.001
Reduced eGFR ^c (%)	11.4	22.2	9.9	<0.001
Plasma creatinine, mg/dL	0.94 \pm 0.28	0.96 \pm 0.36	0.93 \pm 0.27	0.911
Urine creatinine, mg/dL	98.8 \pm 68.4	90.1 \pm 58.2	99.9 \pm 69.7	0.453
Plasma-to-urine creatinine ratio	0.016 \pm 0.015	0.016 \pm 0.012	0.016 \pm 0.015	0.391
Urine Cd, $\mu\text{g/L}$	4.59 \pm 11.92	2.75 \pm 6.94	4.84 \pm 12.44	0.003
Urine β_2M , $\mu\text{g/L}$	2327 \pm 11,517	4105 \pm 21,667	2010 \pm 8532	<0.001
Normalized to C_{cr} (E_x/C_{cr}) ^d				
(E_{Cd}/C_{cr}) \times 100, $\mu\text{g/L}$ filtrate	4.48 \pm 10.38	3.32 \pm 8.40	4.64 \pm 10.62	<0.001
(E_{β_2M}/C_{cr}) \times 100, $\mu\text{g/L}$ filtrate	3858 \pm 20,142	6694 \pm 33,617	3352 \pm 16,642	<0.001
(E_{β_2M}/C_{cr}) \times 100, $\mu\text{g/L}$ filtrate (%)				
≥ 300	36.3	75.3	29.3	<0.001
≥ 1000	20.4	49.4	14.8	<0.001
Normalized to E_{cr} (E_x/E_{cr}) ^e				
E_{Cd}/E_{cr} , $\mu\text{g/g}$ creatinine	4.12 \pm 7.80	2.75 \pm 5.52	4.30 \pm 8.04	<0.001
E_{β_2M}/E_{cr} , $\mu\text{g/g}$ creatinine	2540 \pm 11,420	4004 \pm 13,978	2279 \pm 10,899	<0.001
E_{β_2M}/E_{cr} , $\mu\text{g/g}$ creatinine (%)				
≥ 300	37.6	79.0	30.2	<0.001
≥ 1000	21.3	53.1	15.6	<0.001

Abbreviations: n , number of subjects; BMI, body mass index; β_2M , β_2 -microglobulin; eGFR, estimated glomerular filtration rate; E_x , excretion of x ; cr , creatinine; C_{cr} , creatinine clearance; Cd, cadmium; ^a Obese was defined as BMI $> 30 \text{ kg/m}^2$; ^b eGFR, was determined by CKD-EPI equations [28]; ^c reduced eGFR corresponds to eGFR $\leq 60 \text{ mL/min/1.73m}^2$; ^d $E_x/E_{cr} = [x]_u/[cr]_u$; ^e $E_x/C_{cr} = [x]_u[cr]_p/[cr]_u$, where $x = \beta_2M$ or Cd [29]. Data for all continuous variables are arithmetic means \pm standard deviation (SD). Data for β_2M were from 535 subjects. Data for all other continuous variables were from 674 subjects. For all tests, $p \leq 0.05$ identifies statistical significance, determined by Pearson chi-square test for % differences and by the Mann-Whitney U-test for mean differences between diabetes and non-diabetes.

Females constituted 57.6% of the cohort, and they formed 74.1% of the diabetic group ($n = 81$). Seventy subjects with diagnosed diabetes (86.4%) were residents of the low-exposure location, Pakpoo, and 11 diabetic cases were from a high-exposure area. The diabetic group was older (58.5 years) than the non-diabetic group (45.6 years), and the overall mean age was 47.2 years. The percentage (%) of smoking was lower in diabetics

(14.9%) than in non-diabetics (34.2%). The % of obese participants was higher in the diabetics than in the non-diabetics (14.8% vs. 2.4%).

The diabetic group had a lower mean eGFR and a higher % of reduced eGFR, defined as $eGFR \leq 60 \text{ mL/min/1.73 m}^2$, compared to non-diabetics (22.2% vs. 9.9%). The mean E_{Cd}/C_{cr} and mean E_{Cd}/E_{cr} in diabetics were all lower than in non-diabetics. Conversely, the mean $E_{\beta_{2M}}/C_{cr}$ and mean $E_{\beta_{2M}}/E_{cr}$ in diabetics were all higher than in non-diabetics. Nearly half (49.4%) of the diabetic group and 14.8% of the non-diabetic group had severe tubular proteinuria ($E_{\beta_{2M}}/C_{cr}$ values $\geq 1000 \mu\text{g/L}$ filtrate).

3.2. Predictors of Tubular Proteinuria and Reduced eGFR

We employed logistic regression analysis to screen factors that may increase the risk of tubular proteinuria and/or reduced eGFR. In this analysis, the independent variables were age, BMI, $\log_2[(E_{Cd}/C_{cr}) \times 10^5]$, hypertension, smoking, gender, and diabetes (Table 2).

Table 2. Risk factors for tubular dysfunction and reduced eGFR.

Parameters	Number of Participants	Tubular Dysfunction ^a		Reduced eGFR ^b	
		POR (95% CI)	<i>p</i>	POR (95% CI)	<i>p</i>
Age, year	535 (674)	1.100 (1.076, 1.126)	<0.001	1.146 (1.107, 1.188)	<0.001
BMI, kg/m ²	535 (674)	0.963 (0.908, 1.021)	0.202	1.070 (0.993, 1.153)	0.075
$\log_2[(E_{Cd}/C_{cr}) \times 10^5]$	535 (674)	1.149 (1.054, 1.253)	0.002	1.500 (1.304, 1.724)	<0.001
Hypertension	162 (174)	0.724 (0.444, 1.181)	0.195	0.744 (0.387, 1.428)	0.374
Smoking	195 (215)	1.563 (0.863, 2.830)	0.140	1.124 (0.511, 2.474)	0.771
Gender (male)	223 (286)	0.819 (0.480, 1.398)	0.464	1.284 (0.615, 2.680)	0.506
Diabetes	81 (81)	8.690 (4.421, 17.08)	<0.001	2.973 (1.274, 6.937)	0.012

Abbreviations: POR, prevalence odds ratio; CI, confidence interval. Coding: female = 1, male = 2, normotension = 1, hypertension = 2, non-smoker = 1, smoker = 2. ^a Tubular dysfunction was defined as $(E_{\beta_{2M}}/C_{cr}) \times 100 \geq 300 \mu\text{g/L}$ filtrate; ^b reduced eGFR was defined as estimated GFR $\leq 60 \text{ mL/min/1.73 m}^2$. Data were generated from logistic regression analyses relating POR to tubular proteinuria and reduced eGFR to a set of seven independent variables (first column). First number and number in parenthesis in second column correspond to number of participants in analyses of tubular dysfunction and reduced eGFR, respectively. For all tests, *p*-values ≤ 0.05 indicate a statistically significant association of POR with a given independent variable.

BMI, hypertension, smoking, and gender did not show significant associations with tubular dysfunction or a reduced eGFR. Three other independent variables, namely, age, diabetes, and E_{Cd}/C_{cr} , were all associated with both tubular proteinuria and reduced eGFR.

Tubular dysfunction and reduced eGFR were more prevalent in the diabetic than the non-diabetic groups by 8.7-fold ($p < 0.001$) and by 3-fold ($p = 0.012$), respectively. For a one-year increase in age, the POR for tubular dysfunction and reduced eGFR rose by 10% ($p < 0.001$) and 15% ($p < 0.001$), respectively. The doubling of E_{Cd}/C_{cr} was associated with a 15% increase in the POR for tubular dysfunction ($p = 0.002$) and a 50% increase in the POR for reduced eGFR ($p < 0.001$).

3.3. Effects of Cadmium and Diabetes on β_{2M} Excretion

We further evaluated the effects of diabetes and Cd exposure on $E_{\beta_{2M}}$ by scatterplots and covariance analyses (Figure 1).

A direct relationship was seen between $E_{\beta_{2M}}/C_{cr}$ and E_{Cd}/C_{cr} in the non-diabetic group ($R^2 0.136$, $p < 0.001$) (Figure 1a). After controlling for interactions and differences in age and BMI, the mean $\log[(E_{\beta_{2M}}/C_{cr}) \times 10^4]$ was, respectively, the highest, in the middle, and lowest in non-diabetics who had high, moderate, and low E_{Cd}/C_{cr} ranges ($F = 24.08$, $p < 0.001$) (Figure 1b).

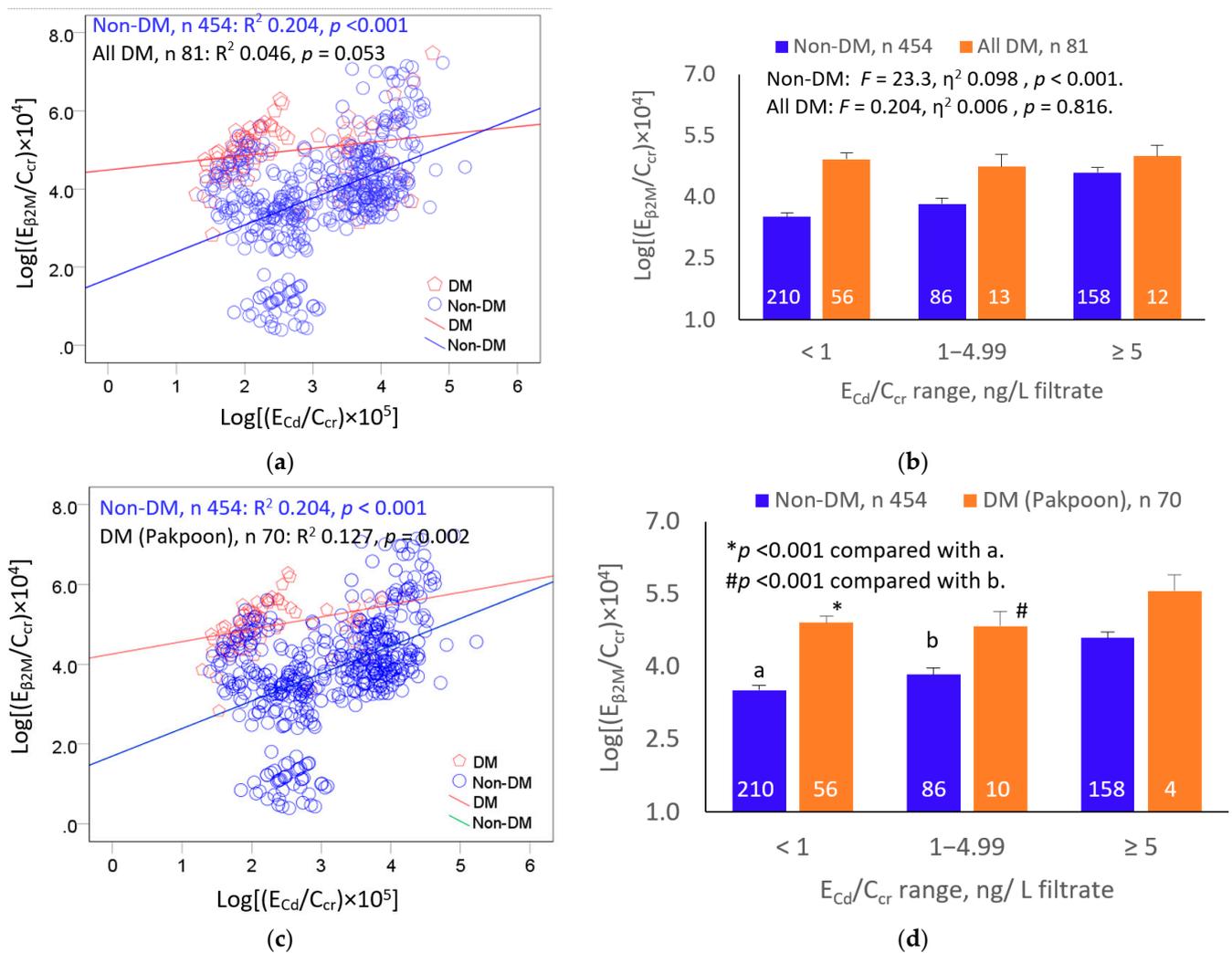


Figure 1. Effects of diabetes and cadmium exposure on β_2M excretion. Scatterplot (a) relates $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ to $\log[(E_{Cd}/C_{cr}) \times 10^5]$ in all diabetic and all non-diabetic participants. Bar graph (b) depicts mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ in all diabetic and non-diabetics grouped by ranges of $\log[(E_{Cd}/C_{cr}) \times 10^5]$. Scatterplot (c) relates $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ to $\log[(E_{Cd}/C_{cr}) \times 10^5]$ in diabetics from Pakpoon and all non-diabetics. Bar graph (d) depicts mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ in diabetics from Pakpoon and all non-diabetics grouped by ranges of $\log[(E_{Cd}/C_{cr}) \times 10^5]$. Coefficients of determination (R^2) and p -values are provided for all scatterplots. Mean values were adjusted for age, BMI, and interactions. Units of $E_{\beta 2M}/C_{cr}$ and E_{Cd}/C_{cr} are ng/L of filtrate.

In the analysis, including all 81 diabetics (Figure 1a), the relationship between $E_{\beta 2M}/C_{cr}$ and E_{Cd}/C_{cr} did not reach a statistically significant level ($p = 0.053$), and the variation in the mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ across the three E_{Cd}/C_{cr} ranges was insignificant ($F = 0.204$, $p = 0.816$) (Figure 1b). However, the relationship between $E_{\beta 2M}$ and E_{Cd} was significant when 11 diabetic cases from a high-Cd exposure area were excluded (R^2 0.127, $p = 0.002$) (Figure 1c). After adjustment for age and BMI, the mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ was higher in diabetic subjects of low and moderate ranges of E_{Cd}/C_{cr} than in non-diabetics of similar E_{Cd}/C_{cr} ranges (Figure 1d). For the highest E_{Cd}/C_{cr} subsets, the mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ in the diabetics and non-diabetics was similar.

The results of the regression analyses of β_2M excretion are provided separately for the diabetics, diabetics from a low-exposure area, and non-diabetics (Table 3).

Table 3. Association of β_2M excretion with cadmium excretion and other independent variables.

Independent Variables/Factors	Excretion Rate of β_2M ^a					
	All Diabetics, <i>n</i> = 81		Diabetics, <i>n</i> = 70		Non-Diabetics, <i>n</i> = 454	
	β ^b	<i>p</i>	β	<i>p</i>	β	<i>p</i>
Age, years	0.202	0.072	0.159	0.159	0.458	<0.001
Log ₂ [(E _{Cd} /C _{cr}) × 10 ⁵], µg/L filtrate	0.181	0.113	0.375	0.001	0.269	<0.001
Smoking	0.206	0.155	0.253	0.055	0.036	0.443
Obesity	0.145	0.194	0.273	0.015	−0.049	0.193
Gender	−0.136	0.328	−0.209	0.122	0.025	0.572
Hypertension	0.105	0.346	−0.032	0.767	0.024	0.532
Adjusted R ²	0.089	0.043	0.244	0.001	0.386	<0.001

Abbreviations: *n*, number of participants. ^a Excretion rate of β_2M as log[(E _{β_2M} /C_{cr}) × 10⁴]; ^b β , standardized regression coefficients. Coding: female = 1, male = 2, normotension = 1, hypertension = 2, non-smoker = 1, smoker = 2, obese = 1, non-obese = 2. Data were generated from multiple regression model analyses relating E _{β_2M} /C_{cr} to six independent variables (first column) in all participants, diabetes, and non-diabetes. For all tests, *p*-values < 0.05 indicate a statistically significant association. β coefficients indicate the strength of the association of E _{β_2M} /C_{cr} and independent variables. Adjusted R² indicates the proportion of the variation in E _{β_2M} /C_{cr} attributable to all six independent variables.

In a model including 70 diabetic cases from the low-exposure locality (Pakpoon), 24.4% of the variation in E _{β_2M} /C_{cr} was explained by all six independent variables. E _{β_2M} /C_{cr} varied directly with E_{Cd}/C_{cr} (β = 0.375, *p* = 0.001) and obesity (β = 0.273, *p* = 0.015). However, when all 81 diabetics were included in the analysis, only 8.9% of the E _{β_2M} /C_{cr} variation was explained by age, log₂[(E_{Cd}/C_{cr}) × 10⁵], smoking, obesity, gender, and hypertension. In effect, none of these six variables showed a significant association with E _{β_2M} /C_{cr} in the diabetic group.

In the non-diabetic group, the six independent variables explained 38.6% of the E _{β_2M} /C_{cr} variability, where E _{β_2M} /C_{cr} varied directly with age (β = 0.458, *p* < 0.001) and with E_{Cd}/C_{cr} (β = 0.269, *p* < 0.001).

3.4. Inverse Association of eGFR and Cadmium

Similarly, we used multiple regression analyses to compare the strength of the association of the eGFR and E_{Cd}/C_{cr} in three subsets (Table 4).

Table 4. Associations of eGFR with cadmium excretion and other independent variables.

Independent Variables/Factors	eGFR ^a , mL/min/1.73m ²					
	All Diabetics, <i>n</i> = 81		Diabetics, <i>n</i> = 70		Non-Diabetics, <i>n</i> = 593	
	β ^b	<i>p</i>	β	<i>p</i>	β	<i>p</i>
Age, year	−0.444	<0.001	−0.472	<0.001	−0.574	<0.001
Log ₂ [(E _{Cd} /C _{cr}) × 10 ⁵], µg/L filtrate	−0.244	0.014	−0.145	0.167	−0.263	<0.001
Smoking	−0.208	0.095	−0.175	0.167	0.033	0.352
Obesity	−0.151	0.116	−0.146	0.170	0.037	0.190
Gender	0.163	0.173	0.162	0.214	−0.024	0.448
Hypertension	−0.045	0.634	−0.066	0.530	−0.014	0.633
Adjusted R ²	0.330	<0.001	0.293	<0.001	0.534	<0.001

Abbreviations: *n*, number of participants. ^a eGFR was determined by CKD-EPI equations [28]; ^b β , standardized regression coefficients. Coding: female = 1, male = 2, normotension = 1, hypertension = 2, non-smoker = 1, smoker = 2. Data were generated from multiple regression model analyses relating eGFR to six independent variables (first column) in all participants, diabetes, and non-diabetes. For all tests, *p*-values < 0.05 indicate a statistically significant association. β coefficients indicate the strength of the association of the eGFR and independent variables. Adjusted R² indicates the proportion of the variation in eGFR attributable to all six independent variables.

In a model including all diabetics, eGFR was inversely associated with age (β = −0.444) and E_{Cd}/C_{cr} (β = −0.244), and these two variables, plus smoking, obesity, gender, and hypertension accounted for 33% of the eGFR variability. These six independent variables

explained 29.3% of the total variation in eGFR among diabetics from a low-Cd-exposure area, and only age showed a significant association with eGFR reduction ($\beta = -0.472$). An association between the eGFR and E_{Cd}/C_{cr} was insignificant ($\beta = -0.145, p = 0.167$).

In the non-diabetic group, age, E_{Cd}/C_{cr} , smoking, obesity, gender, and hypertension together accounted for 53.4% of the total eGFR variation. Distinct from the diabetics from a low-exposure area, the eGFR among those without diabetes was inversely associated with age ($\beta = -0.574$) and E_{Cd}/C_{cr} ($\beta = -0.263$).

3.5. Inverse Association of β_2M Excretion and eGFR

To assess the association of E_{β_2M}/C_{cr} with the eGFR, we employed scatterplots and covariance analyses, where differences in age and BMI were adjusted together with interactions (Figure 2).

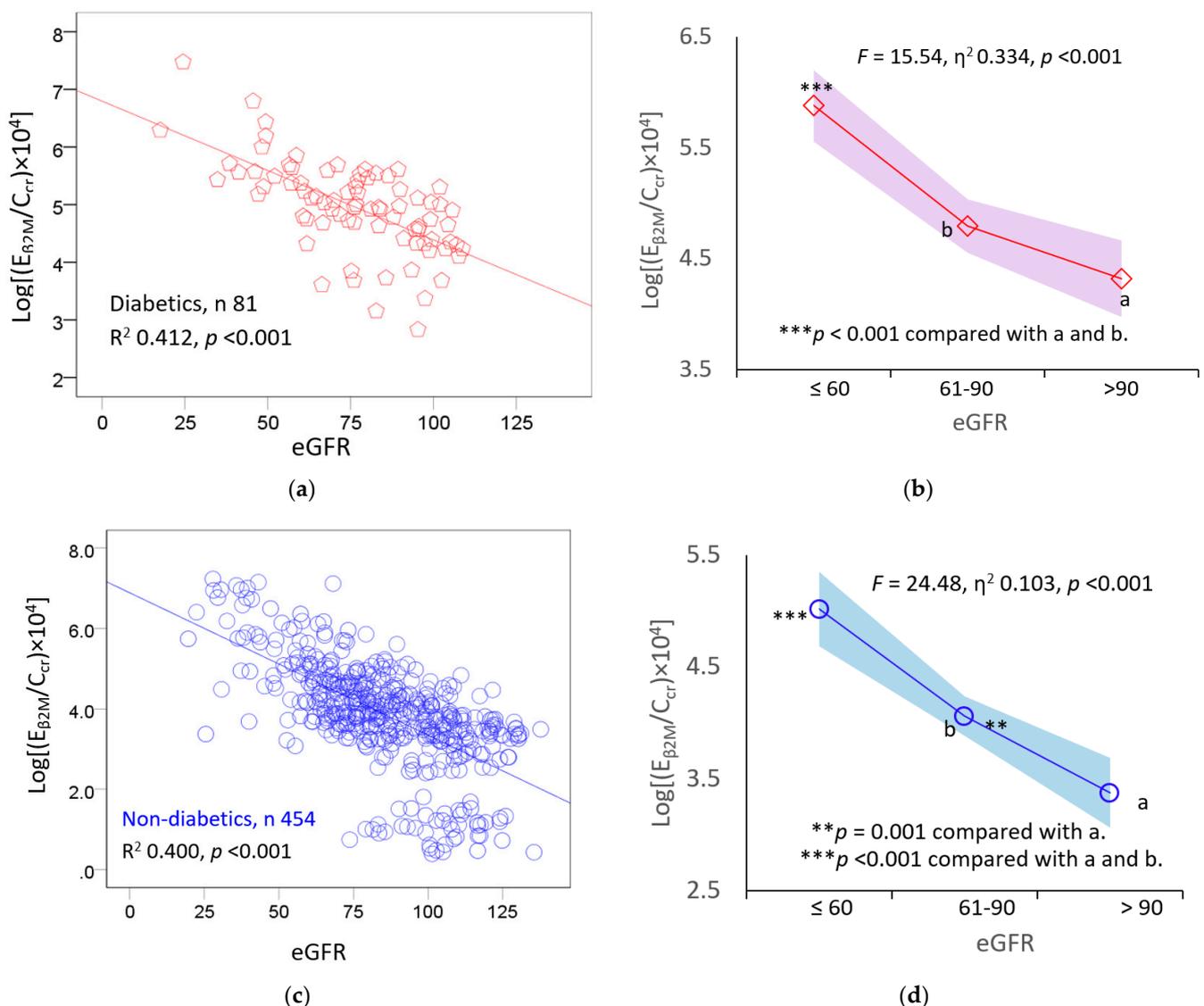


Figure 2. Excretion rates of β_2M in relation to eGFR reduction. Scatterplot (a) relates $\log[(E_{\beta_2M}/C_{cr}) \times 10^4]$ to eGFR in all diabetics. Graph (b) depicts mean $\log[(E_{\beta_2M}/C_{cr}) \times 10^4]$ and the variability of each mean in diabetics grouped by ranges of eGFR. Scatterplot (c) relates $\log[(E_{\beta_2M}/C_{cr}) \times 10^4]$ to eGFR among non-diabetics. Graph (d) depicts mean $\log[(E_{\beta_2M}/C_{cr}) \times 10^4]$ and the variability of each mean in diabetics grouped by ranges of eGFR. Coefficients of determination (R^2) and p -values are provided for all scatterplots. Mean values were adjusted for covariates and interactions. Unit of E_{β_2M}/C_{cr} is $\mu\text{g}/\text{L}$ of filtrate, and the unit of eGFR is $\text{mL}/\text{min}/1.73 \text{ m}^2$.

In the diabetic group, there was a strong inverse relationship between $E_{\beta 2M}/C_{cr}$ and the eGFR (R^2 0.412, $p < 0.001$) (Figure 2a), and the eGFR explained 33.4% of the $E_{\beta 2M}/C_{cr}$ variation across the three eGFR ranges (Figure 2b). A large proportion of the $E_{\beta 2M}/C_{cr}$ variation was explained by a single variable, eGFR.

Similarly, an inverse relationship was seen between $E_{\beta 2M}/C_{cr}$ and E_{Cd}/C_{cr} in the non-diabetic group (R^2 0.400, $p < 0.001$) (Figure 2c). The mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ was the highest, in the middle, and lowest in those with an eGFR ≤ 60 , 61–90, and >90 mL/min/1.73m², respectively (Figure 2d). The eGFR explained only 10.3% of the $E_{\beta 2M}/C_{cr}$ variation across these ranges of the eGFR ($F = 24.48$, $p < 0.001$). This variation in $E_{\beta 2M}/C_{cr}$ attributable to the eGFR was smaller compared to the diabetic group (Figure 2b).

3.6. $\beta 2M$ Excretion as a Function of GFR and Kidney Cadmium Burden

Because the GFR showed a strong influence on $\beta 2M$ excretion (Figure 2), we next compared $E_{\beta 2M}/C_{cr}$ in subsets with a normal or reduced eGFR across three E_{Cd}/C_{cr} ranges (<1 , 1–4.99 and ≥ 5 ng/L filtrate). The results of these analyses are shown in Figure 3.

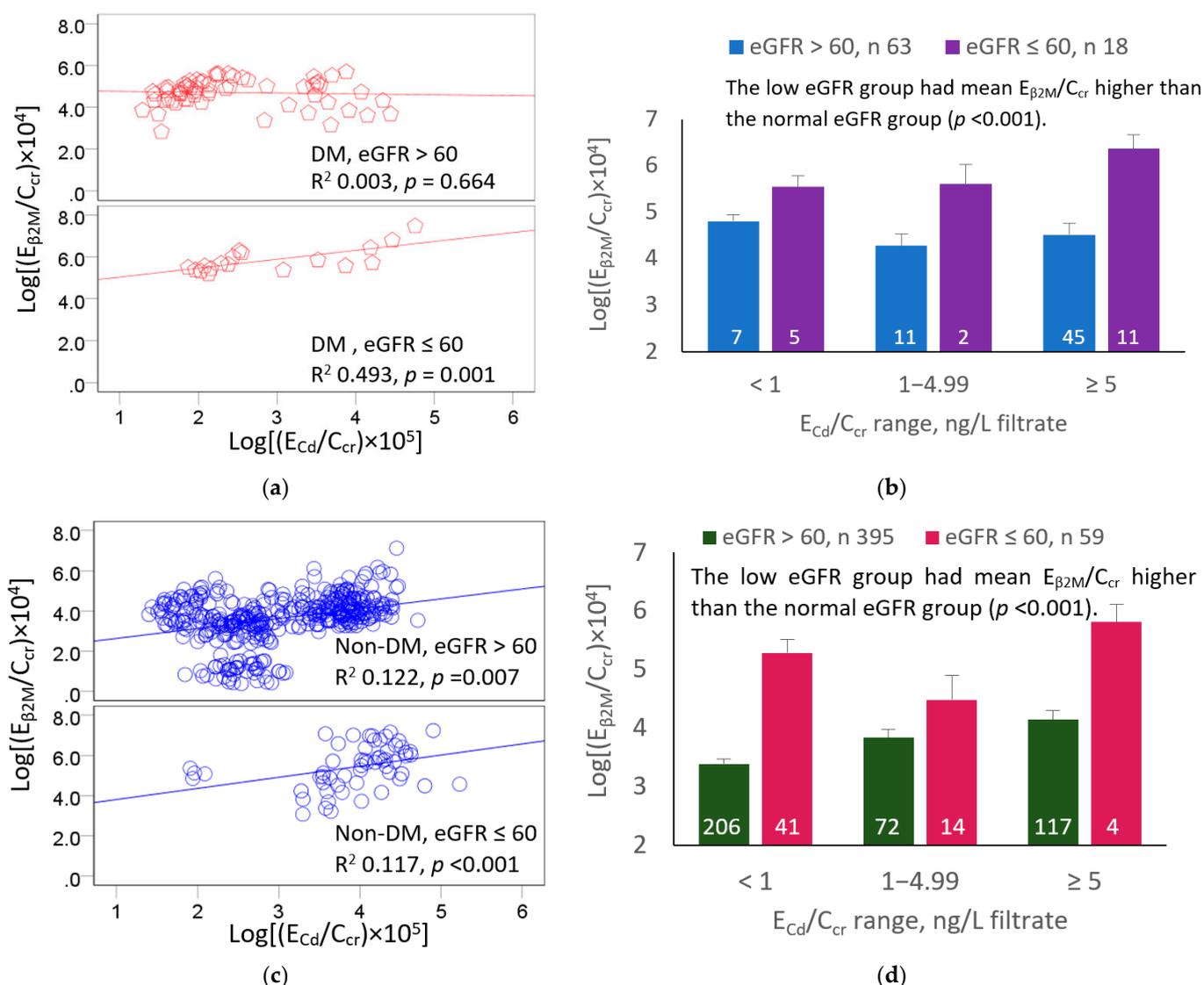


Figure 3. Increments of $\beta 2M$ excretion as a function of kidney cadmium burden and GFR. Scatterplot (a) relates $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ to $\log[(E_{Cd}/C_{cr}) \times 10^5]$ in diabetics grouped by eGFR levels > 60

and ≤ 60 mL/min/1.73 m². Bar graph (b) depicts mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ in diabetics grouped by eGFR and ranges of E_{Cd}/C_{cr} . Scatterplot (c) relates $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ to $\log[(E_{Cd}/C_{cr}) \times 10^5]$ in non-diabetics grouped by eGFR levels > 60 and ≤ 60 mL/min/1.73 m². Graph (d) depicts mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ in non-diabetics grouped by eGFR and ranges of E_{Cd}/C_{cr} . Coefficients of determination (R^2) and p -values are provided for all scatterplots. Mean values were adjusted for age and BMI differences, and interactions. Units of $E_{\beta 2M}/C_{cr}$ and E_{Cd}/C_{cr} are ng/L of filtrate, and the unit of eGFR is mL/min/1.73 m².

Among 81 diabetics, 63 had normal eGFR, while 18 (22.2%) had reduced eGFR (Figure 3a,b). A direct relationship between $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ and $\log[(E_{Cd}/C_{cr}) \times 10^5]$ was seen only in the reduced eGFR group (R^2 0.493, $p < 0.001$). In covariance analysis (Figure 2b), the mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ tended to be higher in the low eGFR subsets of all three E_{Cd}/C_{cr} ranges. The overall mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ was higher in the low eGFR than that of the high eGFR group ($p < 0.001$).

Among the 454 non-diabetics, 395 had a normal eGFR, while 59 (13%) had a reduced eGFR (Figure 3c,d). $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ rose with $\log[(E_{Cd}/C_{cr}) \times 10^5]$ in both the low and normal eGFR groups. Like the diabetics, the mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ tended to be higher in the low eGFR subsets of all three E_{Cd}/C_{cr} ranges. The overall mean $\log[(E_{\beta 2M}/C_{cr}) \times 10^4]$ was higher in the reduced eGFR than in the normal eGFR group ($p < 0.001$).

4. Discussion

In the present study, we compared the severity of Cd-induced nephropathy in diabetics and non-diabetics living in low- and high-Cd-exposure areas of Thailand. The prevalence of a reduced eGFR (below 60 mL/min/1.73 m²) in our cohort was 11.4% which is higher than that reported in studies from Spain (7%) [38] and Taiwan (6.3%) [39], but in line with the global prevalence of CKD, which varies between 8% and 16% [40]. Approximately, one in five cohort participants (27.6%) had tubular dysfunction, based on the conventional $E_{\beta 2M}/E_{cr} \geq 300$ $\mu\text{g/g}$ creatinine criterion (Table 1). We found that BMI, hypertension, smoking, and gender were independent variables that showed no significant associations with the risk for tubular proteinuria or a reduced eGFR, but age, diabetes, and measured long-term Cd exposure (body burden) did (Table 2).

For every one-year increase in age, the risks of tubular dysfunction and a reduced eGFR increased by 10% and 15%, respectively. The doubling of the Cd body burden increased the risk of tubular dysfunction by 15% while raising the risk of reduced eGFR by 50%. Thus, Cd had a particularly strong effect on the GFR in this population. This is consistent with the results of many other studies, which have linked an elevated risk of a reduced eGFR to environmental exposure to Cd, including studies from the U.S. [11,17–19], Thailand [20] Guatemala [21], Myanmar [22], Taiwan [23], and Korea [24,25].

The influence of both Cd body burden and the eGFR on $E_{\beta 2M}/C_{cr}$ were evident when participants were stratified by the eGFR and ranges of E_{Cd}/C_{cr} , a measure of Cd body burden. In the diabetic group, 33.4% of the variation in $E_{\beta 2M}/C_{cr}$ was associated with the eGFR (Figure 2b). This is a very large variation in the $E_{\beta 2M}/C_{cr}$ that was due to a single variable, eGFR. In comparison, the eGFR explained only 10.3% of the $E_{\beta 2M}/C_{cr}$ variation in the non-diabetic group (Figure 2b). After adjustment for age and BMI, the mean values of $E_{\beta 2M}/C_{cr}$ in the diabetic and non-diabetic subsets were found to be higher in those with reduced eGFR compared to the normal eGFR subsets of similar Cd body burden (Figure 3a,b).

Evidence for increased susceptibility to Cd-induced tubulopathy among the diabetics comes from covariate analysis, where a relationship between $E_{\beta 2M}/C_{cr}$ and E_{Cd}/C_{cr} was seen only in those from a low-Cd-exposure region (Figure 1, Table 3). These findings are in line with the published reports showing the high susceptibility to the nephrotoxicity of Cd among people with diabetes, as discussed below.

The Cadmibel study found that diabetics were more susceptible than non-diabetics to Cd-induced nephrotoxicity [7]. A similar observation followed in studies conducted in Sweden [8,9], the Australian Torres Strait, [10], the U.S. [11], and Korea [12]. Experimental studies have shown that nephropathy due to diabetes and Cd are magnified when both the metal and the disease are present. The injection of Cd-MT complexes into obese diabetic mice and non-obese littermates resulted in increased urinary excretion of proteins and calcium in both groups [41]. However, in the diabetic mice, the dose of Cd-MT required to induce proteinuria and calciuria was one-fourth of that required in the controls. Cd-MT induced glycosuria in both groups. Chinese hamsters with hereditary diabetes are also highly susceptible to Cd-induced nephrotoxicity [42]. In recent histopathological studies, kidney tubular degeneration and fibrosis due to Cd were more pronounced in diabetic than non-diabetic rats [43,44].

Increases in the risks of prediabetes and diabetes among U.S. adults have been associated with E_{Cd}/E_{Cr} of 1–2 $\mu\text{g}/\text{g}$ creatinine [1,2]. In a community-based study in Dallas, Texas, an elevated risk of diabetes was linked to environmental Cd exposure [3]. In a meta-analysis of pooled data from 42 studies, the risks of prediabetes and diabetes increased linearly with blood and urinary Cd; the risk of prediabetes reached a plateau at an E_{Cd}/E_{Cr} rate of 2 $\mu\text{g}/\text{g}$ creatinine, and the diabetic risk rose as blood Cd reached 1 $\mu\text{g}/\text{L}$ [5,6]. Of note, these urinary Cd and blood Cd levels have also been in the range associated with a reduced eGFR in studies conducted in many countries listed previously. In a Chinese population study, dietary Cd exposure estimates of 23.2, 29.6, and 36.9 $\mu\text{g}/\text{d}$ were associated with 1.73-, 2.93-, and 4.05-fold increments in the prevalence of CKD, compared to a 16.7 $\mu\text{g}/\text{d}$ intake level [45]. A diet high in rice, pork, and vegetables was associated with a 4.56-fold increase in the prevalence of CKD [45].

Smoking has been shown to promote both the onset and progression of CKD [46,47]. In a meta-analysis of data from 104 studies, an increase in the odds of CKD of 18% was seen among current and former smokers compared to those who never smoked [46]. A Singaporean prospective cohort ($n = 63,257$, 30.6% were smokers) has implicated smoking in the progression of CKD [47]. With adjustment for confounders, smoking increased the risk of end-stage kidney disease by 29% compared to non-smokers, while the risk of kidney failure diminished after quitting smoking for more than 10 years [47]. A strong dose-dependent association was seen between the number of years of smoking and kidney failure [47].

In the present study, the risk of CKD (reduced eGFR) was not associated with smoking; instead, it was associated with an indicator of cumulative Cd exposure from all sources. Per the doubling of Cd body burden, there was a 50% increase in the prevalence odds of CKD (Table 2). Smoking is a significant source of Cd exposure, given that cigarette smoke contains Cd in volatile metallic and oxide (CdO) forms, which have transmission rates 5 to 10 times higher than those that enter through the gut [48]. Cd exposure through smoking has been found to increase the risk of diabetic nephropathy in a Dutch cross-sectional study, including 231 patients with type 2 diabetes, where active smokers were found to have significantly higher blood Cd compared to never smokers and former smokers [49]. Data also demonstrated that smoke-derived Cd mediated this nephrotoxicity [49]. In a six-year median follow-up of these 231 diabetic patients, both Cd and active smoking were associated with progressive eGFR reduction [50]. Collectively, findings from the Dutch cohorts support the premise that exposure to even low levels of environmental Cd promotes the development and progression of diabetic kidney disease. This lends support to our observation that people with diabetes are more susceptible to Cd-induced tubulopathy than non-diabetics.

5. Conclusions

This study shows that tubular dysfunction and a reduced eGFR are more severe and more prevalent in diabetics than non-diabetics of similar age, BMI, and Cd body burden.

Public health resources that promote cessation of smoking and educate consumers about foods known to contain high levels of Cd are likely to have significant health benefits.

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Data Availability Statement: All data are contained within this article.

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References

1. Schwartz, G.G.; Il'yasova, D.; Ivanova, A. Urinary cadmium, impaired fasting glucose, and diabetes in the NHANES III. *Diabetes Care* **2003**, *26*, 468–470. [[CrossRef](#)] [[PubMed](#)]
2. Wallia, A.; Allen, N.B.; Badon, S.; El Muayed, M. Association between urinary cadmium levels and prediabetes in the NHANES 2005–2010 population. *Int. J. Hyg. Environ. Health* **2014**, *217*, 854–860. [[CrossRef](#)] [[PubMed](#)]
3. Little, B.B.; Reilly, R.; Walsh, B.; Vu, G.T. Cadmium is associated with type 2 diabetes in a Superfund Site Lead Smelter Community in Dallas, Texas. *Int. J. Environ. Res. Public Health* **2020**, *17*, 4558. [[CrossRef](#)]
4. Xiao, L.; Li, W.; Zhu, C.; Yang, S.; Zhou, M.; Wang, B.; Wang, X.; Wang, D.; Ma, J.; Zhou, Y.; et al. Cadmium exposure, fasting blood glucose changes, and type 2 diabetes mellitus: A longitudinal prospective study in China. *Environ. Res.* **2021**, *192*, 110259. [[CrossRef](#)]
5. Guo, F.F.; Hu, Z.Y.; Li, B.Y.; Qin, L.Q.; Fu, C.; Yu, H.; Zhang, Z.L. Evaluation of the association between urinary cadmium levels below threshold limits and the risk of diabetes mellitus: A dose-response meta-analysis. *Environ. Sci. Pollut. Res. Int.* **2019**, *26*, 19272–19281. [[CrossRef](#)] [[PubMed](#)]
6. Filippini, T.; Wise, L.A.; Vinceti, M. Cadmium exposure and risk of diabetes and prediabetes: A systematic review and dose-response meta-analysis. *Environ. Int.* **2022**, *158*, 106920. [[CrossRef](#)]
7. Buchet, J.P.; Lauwerys, R.; Roels, H.; Bernard, A.; Bruaux, P.; Claeys, F.; Ducoffre, G.; de Plaen, P.; Staessen, J.; Amery, A.; et al. Renal effects of cadmium body burden of the general population. *Lancet* **1990**, *336*, 699–702. [[CrossRef](#)]
8. Akesson, A.; Lundh, T.; Vahter, M.; Bjellerup, P.; Lidfeldt, J.; Nerbrand, C.; Samsioe, G.; Strömberg, U.; Skerfving, S. Tubular and glomerular kidney effects in Swedish women with low environmental cadmium exposure. *Environ. Health Perspect.* **2005**, *113*, 1627–1631. [[CrossRef](#)]
9. Barregard, L.; Bergstrom, G.; Fagerberg, B. Cadmium, type 2 diabetes, and kidney damage in a cohort of middle-aged women. *Environ. Res.* **2014**, *135*, 311–316. [[CrossRef](#)]
10. Haswell-Elkins, M.; Satarug, S.; O'Rourke, P.; Moore, M.; Ng, J.; McGrath, V.; Walmsby, M. Striking association between urinary cadmium level and albuminuria among Torres Strait Islander people with diabetes. *Environ. Res.* **2008**, *106*, 379–383. [[CrossRef](#)]
11. Madrigal, J.M.; Ricardo, A.C.; Persky, V.; Turyk, M. Associations between blood cadmium concentration and kidney function in the U.S. population: Impact of sex, diabetes and hypertension. *Environ. Res.* **2018**, *169*, 180–188. [[CrossRef](#)]
12. Hwangbo, Y.; Weaver, V.M.; Tellez-Plaza, M.; Guallar, E.; Lee, B.K.; Navas-Acien, A. Blood cadmium and estimated glomerular filtration rate in Korean adults. *Environ. Health Perspect.* **2011**, *119*, 1800–1805. [[CrossRef](#)] [[PubMed](#)]
13. Wallin, M.; Sallsten, G.; Lundh, T.; Barregard, L. Low-level cadmium exposure and effects on kidney function. *Occup. Environ. Med.* **2014**, *71*, 848–854. [[CrossRef](#)] [[PubMed](#)]
14. Satarug, S.; Vesey, D.A.; Gobe, G.C. Cadmium-induced proteinuria: Mechanistic insights from dose–effect analyses. *Int. J. Mol. Sci.* **2023**, *24*, 1893. [[CrossRef](#)] [[PubMed](#)]
15. Järup, L.; Persson, B.; Elinder, C.G. Decreased glomerular filtration rate in solderers exposed to cadmium. *Occup. Environ. Med.* **1995**, *52*, 818–822. [[CrossRef](#)]
16. Järup, L.; Persson, B.; Edling, C.; Elinder, C.G. Renal function impairment in workers previously exposed to cadmium. *Nephron* **1993**, *64*, 75–81. [[CrossRef](#)]
17. Ferraro, P.M.; Costanzi, S.; Naticchia, A.; Sturniolo, A.; Gambaro, G. Low level exposure to cadmium increases the risk of chronic kidney disease: Analysis of the NHANES 1999–2006. *BMC Public Health* **2010**, *10*, 304. [[CrossRef](#)]

18. Navas-Acien, A.; Tellez-Plaza, M.; Guallar, E.; Muntner, P.; Silbergeld, E.; Jaar, B.; Weaver, V. Blood cadmium and lead and chronic kidney disease in US adults: A joint analysis. *Am. J. Epidemiol.* **2009**, *170*, 1156–1164. [[CrossRef](#)]
19. Li, Y.S.; Ho, W.C.; Caffrey, J.L.; Sonawane, B. Low serum zinc is associated with elevated risk of cadmium nephrotoxicity. *Environ. Res.* **2014**, *134*, 33–38.
20. Satarug, S.; Gobe, G.C.; Ujjin, P.; Vesey, D.A. A comparison of the nephrotoxicity of low doses of cadmium and lead. *Toxics* **2020**, *8*, 18. [[CrossRef](#)]
21. Butler-Dawson, J.; James, K.A.; Krisher, L.; Jaramillo, D.; Dally, M.; Neumann, N.; Pilloni, D.; Cruz, A.; Asensio, C.; Johnson, R.J.; et al. Environmental metal exposures and kidney function of Guatemalan sugarcane workers. *J. Expo. Sci. Environ. Epidemiol.* **2022**, *32*, 461–471. [[CrossRef](#)] [[PubMed](#)]
22. Win-Thu, M.; Myint-Thein, O.; Win-Shwe, T.-T.; Mar, O. Environmental cadmium exposure induces kidney tubular and glomerular dysfunction in the Myanmar adults. *J. Toxicol. Sci.* **2021**, *46*, 319–328. [[CrossRef](#)]
23. Tsai, K.F.; Hsu, P.C.; Lee, C.T.; Kung, C.T.; Chang, Y.C.; Fu, L.M.; Ou, Y.C.; Lan, K.C.; Yen, T.H.; Lee, W.C. Association between enzyme-linked immunosorbent assay-measured kidney injury markers and urinary cadmium levels in chronic kidney disease. *J. Clin. Med.* **2021**, *11*, 156. [[CrossRef](#)] [[PubMed](#)]
24. Myong, J.-P.; Kim, H.-R.; Baker, D.; Choi, B. Blood cadmium and moderate-to-severe glomerular dysfunction in Korean adults: Analysis of KNHANES 2005–2008 data. *Int. Arch. Occup. Environ. Health* **2012**, *85*, 885–893. [[CrossRef](#)] [[PubMed](#)]
25. Chung, S.; Chung, J.H.; Kim, S.J.; Koh, E.S.; Yoon, H.E.; Park, C.W.; Chang, Y.S.; Shin, S.J. Blood lead and cadmium levels and renal function in Korean adults. *Clin. Exp. Nephrol.* **2014**, *18*, 726–734. [[CrossRef](#)]
26. Satarug, S.; Vesey, D.A.; Ruangyuttikarn, W.; Nishijo, M.; Gobe, G.C.; Phelps, K.R. The source and pathophysiological significance of excreted cadmium. *Toxics* **2019**, *7*, 55. [[CrossRef](#)]
27. Satarug, S.; Vesey, D.A.; Gobe, G.C.; Phelps, K.R. Estimation of health risks associated with dietary cadmium exposure. *Arch. Toxicol.* **2023**, *97*, 329–358. [[CrossRef](#)]
28. Levey, A.S.; Stevens, L.A.; Schmid, C.H.; Zhang, Y.; Castro, A.F., III; Feldman, H.I.; Kusek, J.W.; Eggers, P.; Van Lente, F.; Greene, T.; et al. A new equation to estimate glomerular filtration rate. *Ann. Intern. Med.* **2009**, *150*, 604–612. [[CrossRef](#)]
29. Phelps, K.R.; Gosmanova, E.O. A generic method for analysis of plasma concentrations. *Clin. Nephrol.* **2020**, *94*, 43–49. [[CrossRef](#)]
30. Satarug, S.; Swaddiwudhipong, W.; Ruangyuttikarn, W.; Nishijo, M.; Ruiz, P. Modeling cadmium exposures in low- and high-exposure areas in Thailand. *Environ. Health Perspect.* **2013**, *121*, 531–536. [[CrossRef](#)]
31. Yimthiang, S.; Pouyfung, P.; Khamphaya, T.; Kuraeiad, S.; Wongrith, P.; Vesey, D.A.; Gobe, G.C.; Satarug, S. Effects of environmental exposure to cadmium and lead on the risks of diabetes and kidney dysfunction. *Int. J. Environ. Res. Public Health* **2022**, *19*, 2259. [[CrossRef](#)]
32. Zhang, Z.W.; Shimbo, S.; Watanabe, T.; Srianujata, S.; Banjong, O.; Chitchumroonchokchai, C.; Nakatsuka, H.; Matsuda-Inoguchi, N.; Higashikawa, K.; Ikeda, M. Non-occupational lead and cadmium exposure of adult women in Bangkok, Thailand. *Sci. Total Environ.* **1999**, *226*, 65–74. [[CrossRef](#)] [[PubMed](#)]
33. Zarcinas, B.A.; Pongsakul, P.; McLaughlin, M.J.; Cozens, G. Heavy metals in soils and crops in Southeast Asia. 2. Thailand. *Environ. Geochem. Health* **2004**, *26*, 359–371. [[CrossRef](#)] [[PubMed](#)]
34. Suwatvitayakorn, P.; Ko, M.S.; Kim, K.W.; Chanpiwat, P. Human health risk assessment of cadmium exposure through rice consumption in cadmium-contaminated areas of the Mae Tao sub-district, Tak, Thailand. *Environ. Geochem. Health* **2020**, *42*, 2331–2344. [[CrossRef](#)] [[PubMed](#)]
35. Swaddiwudhipong, W.; Nguntra, P.; Kaewnate, Y.; Mahasakpan, P.; Limpatanachote, P.; Aunjai, T.; Jeekeeree, W.; Punta, B.; Funkhiew, T.; Phopueng, I. Human health effects from cadmium exposure: Comparison between persons living in cadmium-contaminated and non-contaminated areas in northwestern Thailand. *Southeast Asian J. Trop. Med. Publ. Health* **2015**, *46*, 133–142.
36. Hornung, R.W.; Reed, L.D. Estimation of average concentration in the presence of nondetectable values. *Appl. Occup. Environ. Hyg.* **1990**, *5*, 46–51. [[CrossRef](#)]
37. White, C.A.; Allen, C.M.; Akbari, A.; Collier, C.P.; Holland, D.C.; Day, A.G.; Knoll, G.A. Comparison of the new and traditional CKD-EPI GFR estimation equations with urinary inulin clearance: A study of equation performance. *Clin. Chim. Acta* **2019**, *488*, 189–195. [[CrossRef](#)]
38. Grau-Perez, M.; Pichler, G.; Galan-Chilet, I.; Briongos-Figuero, L.S.; Rentero-Garrido, P.; Lopez-Izquierdo, R.; Navas-Acien, A.; Weaver, V.; Garcia-Barrera, T.; Gomez-Ariza, J.L.; et al. Urine cadmium levels and albuminuria in a general population from Spain: A gene-environment interaction analysis. *Environ. Int.* **2017**, *106*, 27–36. [[CrossRef](#)]
39. Tsai, H.J.; Hung, C.H.; Wang, C.W.; Tu, H.P.; Li, C.H.; Tsai, C.C.; Lin, W.Y.; Chen, S.C.; Kuo, C.H. Associations among heavy metals and proteinuria and chronic kidney disease. *Diagnostics* **2021**, *11*, 282. [[CrossRef](#)]
40. Kalantar-Zadeh, K.; Jafar, T.H.; Nitsch, D.; Neuen, B.L.; Perkovic, V. Chronic kidney disease. *Lancet* **2021**, *398*, 786–802. [[CrossRef](#)]
41. Jin, T.; Frankel, B.J. Cadmium-metallothionein nephrotoxicity is increased in genetically diabetic as compared with normal Chinese hamsters. *Pharmacol. Toxicol.* **1996**, *79*, 105–108. [[CrossRef](#)]
42. Jin, T.; Nordberg, G.F.; Sehlin, J.; Leffler, P.; Wu, J. The susceptibility of spontaneously diabetic mice to cadmium-metallothionein nephrotoxicity. *Toxicol.* **1994**, *89*, 81–90. [[CrossRef](#)]
43. Riaz, M.A.; Nisa, Z.U.; Mehmood, A.; Anjum, M.S.; Shahzad, K. Metal-induced nephrotoxicity to diabetic and non-diabetic Wistar rats. *Environ. Sci. Pollut. Res. Int.* **2019**, *26*, 31111–31118. [[CrossRef](#)]

44. Riaz, M.A.; Nisa, Z.U.; Anjum, M.S.; Butt, H.; Mehmood, A.; Riaz, A.; Akhtar, A.B.T. Assessment of metals induced histopathological and gene expression changes in different organs of non-diabetic and diabetic rats. *Sci. Rep.* **2020**, *10*, 5897. [[CrossRef](#)]
45. Shi, Z.; Taylor, A.W.; Riley, M.; Byles, J.; Liu, J.; Noakes, M. Association between dietary patterns, cadmium intake and chronic kidney disease among adults. *Clin. Nutr.* **2018**, *37*, 276–284. [[CrossRef](#)] [[PubMed](#)]
46. Kelly, J.T.; Su, G.; Zhang, L.; Qin, X.; Marshall, S.; González-Ortiz, A.; Clase, C.M.; Campbell, K.L.; Xu, H.; Carrero, J.J. Modifiable lifestyle factors for primary prevention of CKD: A systematic review and meta-analysis. *J. Am. Soc. Nephrol.* **2021**, *32*, 239–253. [[CrossRef](#)] [[PubMed](#)]
47. Jin, A.; Koh, W.P.; Chow, K.Y.; Yuan, J.M.; Jafar, T.H. Smoking and risk of kidney failure in the Singapore Chinese health study. *PLoS ONE* **2013**, *8*, e62962. [[CrossRef](#)]
48. Pappas, R.S.; Fresquez, M.R.; Watson, C.H. Cigarette smoke cadmium breakthrough from traditional filters: Implications for exposure. *J. Anal. Toxicol.* **2015**, *39*, 45–51. [[CrossRef](#)] [[PubMed](#)]
49. Hagedoorn, I.J.M.; Gant, C.M.; Huizen, S.V.; Maatman, R.G.H.J.; Navis, G.; Bakker, S.J.L.; Laverman, G.D. Lifestyle-related exposure to cadmium and lead is associated with diabetic kidney disease. *J. Clin. Med.* **2020**, *9*, 2432. [[CrossRef](#)]
50. Oosterwijk, M.M.; Hagedoorn, I.J.M.; Maatman, R.G.H.J.; Bakker, S.J.L.; Navis, G.; Laverman, G.D. Cadmium, active smoking and renal function deterioration in patients with type 2 diabetes. *Nephrol. Dial. Transplant* **2023**, *38*, 876–883. [[CrossRef](#)]

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