



Article National Culture and Corporate Rating Migrations

Huong Dieu Dang

Department of Economics and Finance, University of Canterbury, Christchurch 8140, New Zealand; Huong.dang@canterbury.ac.nz

Received: 14 October 2018; Accepted: 7 November 2018; Published: 14 November 2018



Abstract: The informal constraints that arise from the national culture in which a firm resides have a pervasive impact on managerial decision making and corporate credit risk, which in turn impacts on corporate ratings and rating changes. In some cultures, firms are naturally predisposed to rating changes in a particular direction (downgrade *or* upgrade) while, in other cultures, firms are more likely to migrate from the current rating in either direction. This study employs a survival analysis framework to examine the effect of national culture on the probability of rating transitions of 5360 firms across 50 countries over the period 1985–2010. Firms located in *long-term oriented* cultures are less likely to be downgraded and, in some cases, more likely to be upgraded. Downgrades occur more often in strong *uncertainty-avoiding* countries and less often in large *power distance* (*hierarchy*) and *embeddedness* countries. There is some evidence that *masculinity* predisposes firms to more rating transitions. Studying culture helps enrich our understanding of corporate rating migrations, and helps develop predictive models of corporate rating changes across countries.

Keywords: national culture; survival analysis; hazard model; rating migrations

1. Introduction

Hofstede, Hofstede and Minkov (Hofstede et al. 2010) define culture as "the collective programming of the mind that distinguishes the members of one category of people from those of another." In a global survey conducted in May 2008 by PricewaterhouseCoopers, 78% of the participants considered culture and excessive risk taking the contributing factors that led to the credit crisis (PricewaterhouseCoopers 2008). Contributing to this view is the evidence that culture affects corporate risk taking (Li et al. 2013), bank risk taking (Ashraf et al. 2016), and bank performance (Boubakri et al. 2017).

Pan et al. (2017) observe similarity in risk attitudes inside U.S. firms, which are rooted in the founders' risk attitudes and retained through the appointments of leaders with similar mindsets. Unlike Pan et al., this study does not focus on founders' risk attitudes and corporate risk culture.¹ As Ramirez and Tadesse (2009) suggest, the culture in which an individual resides influences his/her perception of uncertainty and his/her "mechanism" to cope with uncertainty and ambiguity. Corporate managers, regardless of their original cultural backgrounds, have to tailor their business priorities and corporate policies to the cultural contexts in which their firms operate. This study therefore focuses on the national culture in which a firm resides.

In practice, credit rating has been widely used as a measure of credit risk. A strong understanding of credit rating decisions has important implications to market participants. As the President of the European Commission José Manuel Barroso remarked at the European Parliament on 5 May 2010, ratings are "too cyclical, too reliant on the general market mood rather than on fundamentals—regardless of whether market mood is too optimistic or too pessimistic".

¹ Pan et al. acknowledge that they do not account for social influences and shared experience inside a firm, thus not capturing a firm's risk culture entirely.

Investors from different cultures react differently to an information shock and are affected by

different cognitive biases (Afego 2018). Behavioral pitfalls, such as overreaction and herding, are related to culture values. For example, people in strong *uncertainty avoiding* countries, such as Greece and Italy, show a high level of anxiety and a natural sense of urgency (Hofstede et al. 2010). People in *collectivism* (*embeddedness*) cultures tend to exhibit herding behavior (Chang and Lin 2015). These culture values affect investor sentiment and market mood, and market mood impacts on credit ratings, as stated by the President of the European Commission. Thus, it is expected that culture indirectly influences rating decisions.

Empirical studies suggest that culture feeds into the rating review process through several channels (Appendix A). Culture directly influences managerial decision making thereby indirectly influencing corporate risk taking (Li et al. 2013), capital structure (Chui et al. 2002), debt maturity (Zheng et al. 2012), cash management (Ramirez and Tadesse 2009; Chang and Noorbakhsh 2009), and corporate investment (Shao et al. 2013). Corporate risk taking, corporate investment, capital structure, debt maturity and cash management are important criteria used in the rating review process (S&P RatingsDirect 2013).

The informal constraints that stem from culture have an extensive influence on daily behaviors, and this influence extends far beyond formal laws (North 1990). Culture directly affects managers' perceptions of questionable business practices (Vitell et al. 1993; Cohen et al. 1996) and indirectly influences managers' choices relating to financial disclosures, financial report quality (Gray 1988; Hope 2003) and the degree of earnings management (Han et al. 2010; Desender et al. 2011). Jorion et al. (2009) consider increased earnings management the contributing factor that has led to a deterioration in the quality of corporate investment ratings.

The conflict of interest between shareholders and creditors, and the extent of protection toward shareholders and creditors vary markedly across countries (Li et al. 2013). Paredes and Wheatley (2017) find that measures of investor protection are "subsumed" by culture. Culture affects managers' decisions to hire a Big Four auditor (Hope et al. 2008). High quality audits reduce information asymmetries and agency conflicts between corporate managers (shareholders) and creditors (Jensen and Meckling 1976). Culture also affects dividend policy (Shao et al. 2010; Fidrmuc and Jacob 2010; Bae et al. 2012), and dividend policy has been widely used as a monitoring mechanism to minimize agency problems. Culture may also complicate the negotiations to avoid a default between corporate managers (shareholders) and creditors.²

There is increasing recognition of the role culture plays in corporate risk taking and corporate credit risk. An interesting question which has not been addressed in the literature is if and how culture affects *changes* in credit risk, specifically, *changes* in ratings. Though culture is very stable, I argue that some cultures may predispose to more rating changes while other cultures may lead to directional predisposition for rating changes over time. As discussed in Section 2.3, a strong uncertainty avoidance trait may lead to more downgrades while a long-term orientation trait may result in fewer downgrades as time passes. Analysts may be more prone to change corporate ratings in countries with a high score for masculinity, individualism, and power distance.

Culture is expected to affect rating migrations through two channels (Appendix A). First, culture influences macro-economic activities³ through its roles as an informal constraint and through its effects on managerial decision making (Zheng et al. 2012). Ratings move pro-cyclically (Lobo et al. 2017) and rating regrades are primarily affected by macro-economic conditions (Blume et al. 1991). Accelerated downgrades and defaults occur more often during economic contractions while upgrades tend to

² Contributing to this view is the evidence that the cultural differences between Greece and Germany made Greece's negotiations to avoid a default much more difficult (Guiso et al. 2016).

³ For example, long-term oriented countries have a higher national saving rate and a higher growth rate. Individualistic countries achieve a higher GNI per capita (Hofstede et al. 2010, pp. 38, 263–65).

outweigh downgrades during economic expansions (Bangia et al. 2002; Lobo et al. 2017). Rating volatility is intensified during business cycle troughs and subdued during business cycle peaks (Nickell et al. 2000).

Second, culture feeds into the rating review process (as discussed above) and thus affects a firm's rating and its rating history. Empirical studies suggest that the current rating and various aspects of rating history impact on subsequent rating changes. For example, issuers in the boundary between investment and speculative rating grades (BBB–/BB+) exhibit different migration dynamics compared with other issuers (Carty and Fons 1994; Johnson 2004). Issuers downgraded to a given rating are more likely to be downgraded than those upgraded to that rating grade.⁴

The above discussion motivates this study to examine the impact of national culture on corporate rating migrations. The effects of culture are examined after accounting for variables that have been found to be significant in previous studies on rating migrations. The results of this study provide robust evidence that national culture significantly impacts the probability of rating changes of 17,109 ratings (5360 firms) across 50 countries over the period 1985–2010. The evidence in favor of including culture variables is stronger for rating downgrades and when numeric scores are used to represent culture values.

This study finds that *long-term orientation (LTO)* value is associated with a lower downgrade hazard, and in some cases, a higher upgrade probability. The effect of *LTO* on downgrades is robust to alternative samples (with and without U.S. firms), alternative measures of culture (Hofstede's and Schwartz's culture scores) and alternative study periods (crisis and non-crisis times). *Uncertainty avoidance* trait makes a downgrade more likely whereas *power distance (hierarchy)* and *embeddedness* trait reduces the risk of a downgrade. There is some evidence that *masculinity* dimension predisposes firms to more rating transitions.

This study extends the literature in two ways. First, this is the first study to explore the effects of culture values established by Hofstede et al. (2010) and Schwartz (1994) on the rating migration hazards of 5360 firms in 50 countries over a 26-year period. The study suggests a cross-disciplinary explanation and establishes a link between culture literature and rating migration literature. Second, the study applies a survival analysis framework (Allison 1995) and develops Cox's dynamic hazard model (Cox 1972) which offers three attractive features: It accounts for the sequence of rating migrations of the same firm, includes both time-fixed and time-varying variables in the estimation process, and allows for non-Markovian behaviors, such as within-rating heterogeneity and time-heterogeneity, in corporate rating migration dynamics.

The remainder of this paper is organized as follows. Section 2 describes the data and discusses culture values established by Hofstede et al. (2010) and Schwartz (1994). Section 3 presents the method, variables and samples. Section 4 discusses the results, and Section 5 summarizes the key findings.

2. National Culture and Corporate Rating Migrations

2.1. Rating Data

This study employs Standard and Poor's foreign currency issuer ratings retrieved from Ratings Xpress database on 28 September 2010. Firm-specific data other than credit ratings are not available. Standard and Poor's (S&P) ratings are used as S&P's decisions are generally timelier than Moody's (Güttler 2011). Foreign currency ratings are used as the persistence of national culture dimensions is relevant for firms which issue debts in international markets.

Financial institutions and utilities are excluded from this study for two reasons. First, previous studies suggest that financial institutions exhibit different rating migration dynamics compared with industrial firms (Nickell et al. 2000; Lando and Skodeberg 2002). Second, financial institutions and

⁴ For the evidence of downward momentum in rating migration dynamics, see Altman and Kao (1992); Carty and Fons (1994); Altman (1998); Bangia et al. (2002); Lando and Skodeberg (2002); Güttler and Wahrenburg (2007); Figlewski et al. (2012); Dang and Partington (2014)

utilities operate in highly regulated environments and are closely monitored by their respective regulatory agencies. This suggests less latitude for corporate managers in these sectors to exhibit flexibility in decision making. Thus, financial institutions and utilities deserve a separate study and this article focuses entirely on other sectors. The study spans the period from 1 January 1985⁵ to 28 September 2010.

2.2. Culture Data

Hofstede's culture values have been widely employed in various business disciplines including finance. This study employs Hofstede et al. (2010)'s five culture values, namely *power distance index*, *uncertainty avoidance index, individualism versus collectivism, masculinity versus femininity,* and *long-term versus short-term orientation*. A country's score for a culture value does not represent an absolute position but reflects its ranking relative to other countries. The validity of culture values persists for a long period of time. Hofstede (2001, p. 255) stresses that in the long term "cultures shift, but they shift in formation so that the differences between them remain intact."

This study employs the numerical scores of Hofstede's five culture values and five dummy variables, each based on a country's score relative to the mean score. Scores greater than or equal to the mean take a value of one or zero otherwise. As the differences between cultures remain stable over time, the use of dummy variables helps maintain the relative rankings between cultures over the study period.

Most culture values introduced in recent studies are conceptually related to and empirically correlated with Hofstede's culture values (Leung et al. 2005).⁶ Two widely known frameworks that categorize culture along similar values are the GLOBE study conducted by Robert House and his team in 1991 and Schwartz (1994). GLOBE values are not used in this study as, despite using the same terms, their meanings are different compared with Hofstede's values. The approach the GLOBE team used to formulate survey questions suffers from major disadvantages (Hofstede et al. 2010, pp. 42–43).

There are significant correlations between Schwartz's scores and Hofstede's scores (Hofstede et al. 2010, p. 41). Given their comparability, this study employs the numeric scores of two of Schwartz's culture measures, *embeddedness*⁷ and *hierarchy*, as alternative measures of Hofstede's *collectivism* (the opposite pole of *individualism*) and *power distance index*, respectively. As in Zheng et al. (2012), this study does not employ Schwartz's *mastery* trait because it captures such values as success, independence, ambition, and capability, which overlap with those of Hofstede's *individualism* and *masculinity* traits.

2.3. Culture Dimensions

The following discussion elaborates on the possible links between culture values and corporate rating migrations, as outlined in Appendix A. In some cultures, firms are naturally predisposed to rating changes in a particular direction (downgrade *or* upgrade) while in other cultures firms tend to migrate from the current rating in either direction.

2.3.1. Power Distance Index (PDI) or Hierarchy

Hofstede's *power distance index* (PDI) and Schwartz's *hierarchy* reflect the extent to which people accept an unequal, hierarchical distribution of power, authority and wealth (Hope 2003; Licht et al. 2005). Small power distance (PD) cultures strive to equalize the distribution of power, and justification

⁵ S&P rating scales were changed in 1983. To calculate the annual changes of employed macro-economic variables in 1985, the values in 1984 and 1985 were needed. Thus, 1985 was chosen as the starting year of the study.

⁶ Tang and Koveos (2008) suggest that institutional factors, such as language, religion, climate and legal origin, are subsumed by Hofstede's *uncertainty avoidance* and *masculinity* traits. The correlations between Hofstede's culture scores and other measures "do not tend to become weaker over time" (Hofstede et al. 2010, p. 39).

⁷ Embeddedness is referred to as conservatism in some studies, such as Johnson and Lenartowicz (1998); Chui et al. (2002); Shao et al. (2010).

is required for any inequalities. Large PD cultures accept a hierarchical order in which everybody has a place and no further justification is required (Hofstede et al. 2010).

Agency conflicts are less severe in large PD countries as there is greater acceptance of wealth and power inequalities (Fidrmuc and Jacob 2010). As a result, firms adopt a lower dividend payout policy and are less likely to hire a Big Four auditor (Hope et al. 2008). It is difficult for firms to get access to long-term credit markets given the greater risk of violence in domestic politics. Thus, firms employ more short-term debt than long-term debt and exhibit a lower degree of financial risk (Zheng et al. 2012). Contributing to this view is the evidence that banks in large PD countries tend to take fewer risks (Ashraf et al. 2016).

A high PD score is related to a high degree of earnings management (Paredes and Wheatley 2017). Corporate managers exhibit strong influences on financial reporting choices and prefer not to disclose information to preserve power inequalities (Zarzeski 1996). People often view a questionable business practice as ethical (Cohen et al. 1996, p. 58) and are more tolerant of corruption and tax evasion (Husted 1999; Tsakumis et al. 2007). Policy debates and political discussions are not often seen in larger PD societies, which indicates a lower degree of transparency and accountability.

The above discussion suggests that a high PD score may predispose firms to rating changes in either direction (downgrade *and* upgrade).

2.3.2. Individualism (IDV) versus Collectivism (Embeddedness)

People in individualistic countries (high IDV score) show autonomy and are encouraged to pursue personal goals and stand up for their rights (Licht et al. 2005). Hofstede's *individualism* trait may predispose firms to rating changes in either direction (downgrade *and* upgrade), as discussed below.

Individualistic countries have strong economies (Hofstede 2001, p. 519) and tend to adopt market-based financial systems, which encourage corporate risk taking (Li et al. 2013). Their preferred Anglo-Saxon corporate governance systems focus on shareholders' interests (Griffin et al. 2015). Firms employ a high degree of leverage (Chui et al. 2002), invest more in long-term assets, R&D projects (instead of physical assets), and employ excess cash to increase R&D (instead of increasing dividends) (Shao et al. 2013). Aggressive risk-taking results in more volatile operating income (Li et al. 2013). Consequently, managers are more likely to manage earnings (Han et al. 2010) and engage in income smoothing (Fonseca and Gonzalez 2008). Agency conflicts are inherently more severe (Chui et al. 2002) as corporate insiders exhibit a strong tendency to pursue their own personal interests rather than adhere to different stakeholders' preferences.

On the positive side, individualistic countries have effective regulatory systems. Laws and rights are equal for all groups, and formal institutions are established to protect the rights of competing parties, such as shareholders and creditors (Licht et al. 2005). Firms tend to adopt a high dividend payout policy to minimize agency problem concerns (Shao et al. 2010; Fidrmuc and Jacob 2010). The business environment in individualistic cultures is more competitive and less secretive than in collectivistic cultures (Gray 1988). Firms are open to extensive accounting disclosures (Gray and Vint 1995) and are more likely to hire a Big Four auditor (Hope et al. 2008). People are less tolerant of corruption and tax evasion (Tsakumis et al. 2007).

In collectivistic (embeddedness) cultures (low IDV score), people have less need for autonomy and an active determination of their own lives. Schwartz's *embeddedness* trait is expected to lower downgrade hazard and raise upgrade probability, as discussed below.

Collectivistic (embeddedness) cultures focus on maintaining a harmonious relationship within a society and preserving public images. People value conformity and adherence to societal norms and regulations (Hofstede 1980). Firms tend to adopt autocratic and paternalistic management systems. Corporate managers are very concerned about the liquidation costs to its stakeholders (Titman 1984). Capital structure is decided by corporate leaders who have strong views that there should be no detrimental effect on employees. As a high degree of financial leverage restricts firms' flexibility and raises the probability of bankruptcy, firms in embeddedness cultures tend to employ less debt

(Chui et al. 2002) and use more short-term debt (Zheng et al. 2012). Avoiding bankruptcy is also consistent with preserving a firm's public images.

Collectivistic (embeddedness) cultures are more likely to have bank-based financial systems and relation-based corporate governance systems, which place a greater emphasis on the interests of a firm's different stakeholders (Griffin et al. 2015). Agency conflicts are less severe and excessive risk taking is not encouraged as managers tend to act in line with the interests of all stakeholders (Jensen and Meckling 1976; Fidrmuc and Jacob 2010; Shao et al. 2013). Corporate managers share a strong sense of responsibility and are less likely to differentiate between their own and others' welfares.

Overall, *individualism* may predispose firms to more rating regrades whereas *embeddedness* is expected to raise upgrade hazard and reduce downgrade risk.

2.3.3. Masculinity (MAS) versus Femininity

Masculine cultures (high MAS score) value assertiveness, ambition, competition, challenge, recognition, material accomplishment, and success. *Feminine* cultures (low MAS score) value cooperation, modesty, tenderness, security, caring for the weak, and quality of life (Hofstede et al. 2010).

Masculine cultures encourage taking risky decisions and achieving performance goals. Men often exhibit overconfidence or self-attribution biases (Barber and Odean 2001). Compared with women, men are less likely to assess the accurate levels of risk associated with an assigned task (Byrnes et al. 1999). Men are also more likely to experience failure or unfavorable outcomes, and can be more aggressive than women about default. Contributing to this view is the evidence that banks in masculine countries are more likely to incur a large loss during a crisis (Kanagaretnam et al. 2011).

Masculine cultures are open to head-on confrontation (Licht et al. 2005) and encourage aggressive behaviors (Kanagaretnam et al. 2011). Conflicts are often resolved by fighting in masculine countries and by negotiation in feminine countries.⁸ The political environment in masculine countries tends to be adversarial, as opposed to cooperative coalitions in feminine countries (Hofstede et al. 2010, pp. 173–80).

Countries with high MAS scores are less likely to perceive ethical issues in business practices (Vitell et al. 1993, p. 758). However, masculine countries incorporate economic interests in legal form (Licht et al. 2005), and are less permissive in dealing with lawbreakers. Masculine countries focus more on punishment while feminine countries are more lenient and place greater emphasis on correction and rehabilitation (Hofstede 2001, p. 319).

Firms in masculine cultures are open to extensive information sharing (Hope 2003; Doupnik and Tsakumis 2004) as societies are more business oriented and value visible achievement. Corporate managers are concerned that debt covenants may restrict their aggressive business plans and fears of financial distress may interfere with their bold investment strategies. Firms thus prefer less debt financing (Chui et al. 2002) and use more short-term debt (Zheng et al. 2012). By contrast, feminine cultures may accept higher indebtedness to support their welfare systems "even at the expense of accomplishment and performance" (Weaver 2001, p. 9).

The above discussion suggests that a high MAS score may predispose firms to rating migrations in either direction (downgrade *and* upgrade).

⁸ An example is the handling of the Åland Islands crisis and the Falkland Islands crisis. The Åland Islands crisis was resolved by negotiations in 1921 between feminine countries Finland and Sweden. The Åland Islands remained Finnish but the pro-Swedish islands gained substantial regional autonomy. The Falklands Islands crisis in 1982 involved Argentinean military and British expeditionary forces. The crisis between two masculine countries cost "725 Argentinean and 225 British lives and enormous financial expense." The Falklands Islands have remained a disputed territory and required "constant British subsidies and military presence" (Hofstede et al. 2010, p. 173).

2.3.4. Uncertainty Avoidance Index (UAI)

Uncertainty avoidance (UA) is the extent to which members of a society feel uncomfortable with uncertainty and find ambiguity stressful (Chui et al. 2002). People from strong UA countries are subject to a natural sense of urgency and are more likely to suffer from anxiety (Hofstede et al. 2010). A strong UA trait tends to exacerbate investor panic and create more uncertainty during market turmoil. Deteriorating market mood and heightened concerns about uncertainty often lead to drastic and accelerated downgrades.⁹

Uncertainty avoidance differs from risk avoidance (Tsakumis et al. 2007). High UA cultures perceive ambiguity as a continuous threat. People tend to "engage in risky behavior in order to reduce ambiguities, such as starting a fight with a potential opponent rather than sitting back and waiting". They are open to familiar risks such as driving faster and having more fatal accidents but they show fears of ambiguity and unfamiliar risks (Hofstede et al. 2010, pp. 197–98).

Strong UA countries tend to adopt bank-based financial systems (Kwok and Tadesse 2006) and maintain a low degree of economic freedom. Firms in these countries have a negative view of competition. Debt financing is minimized as the use of debt imposes constraints and may lead to financial instability (Chui et al. 2002). Corporate managers are concerned about potential financial distress and the cash shortage that may require capital raising under unfavorable market conditions or fire sales of a firm's assets (Chang and Noorbakhsh 2009). Access to long-term credit markets is difficult given the greater risk of political instability.¹⁰ So, firms prefer to use more short-term debt (Zheng et al. 2012), adopt a lower dividend payout policy (Fidrmuc and Jacob 2010) and hold more cash (Ramirez and Tadesse 2009; Chang and Noorbakhsh 2009). Firms are not open to extensive accounting disclosure (Gray and Vint 1995) and are less likely to hire a Big Four auditor (Hope et al. 2008). Information asymmetries and agency conflicts between managers (shareholders) and creditors tend to be greater.

In strong UA countries, "ineffective rules can satisfy people's emotional need for formal structure" (Hofstede et al. 2010, p. 209). Corporate managers are less likely to perceive ethical problems and are not opposed to contravening an unjust law. In situations where outcomes cannot be determined with certainty, corruption is viewed as an option to secure a predictable result. Tax evasion is considered a means of reducing ambiguity (Vitell et al. 1993, p. 757; Husted 1999; Tsakumis et al. 2007; Hofstede et al. 2010, p. 223). By contrast, in weak UA countries rules are often more likely to be followed (Hofstede et al. 2010), investors' legal rights are stronger (Licht et al. 2005), and people are less likely to view tax evasion as a viable option.

In strong UA societies, people are also slower in paying their bills (De Mooij 2004, p. 154). These countries have more precise formal laws, informal rules, rigid safety and security measures in place. However, it takes more time for their citizens to comply with two simple civil procedures: collecting a bounced check which had been refused by a bank and evicting a tenant for non-payment of rent (Hofstede et al. 2010, pp. 216–17). Debt burdened firms in strong UA countries may not waste time negotiating with creditor(s) to reduce their delinquent debts. Driven by a high level of anxiety and a natural sense of urgency, these firms may simply walk away to avoid any uncertainty associated with the debt restructure process.

In the light of this discussion, a high UA score is expected to make downgrades more likely and upgrades less likely.

⁹ Contributing to this view is the remark of the President of the European Commission José Manuel Barroso at the European Parliament in May 2010 that ratings are "too cyclical, too reliant on the general market mood rather than on fundamentals...".

¹⁰ Strong UA countries are intolerant of political ideologies, are "more likely to harbor extremist minorities within their political landscape" and have more "native terrorists" (Hofstede et al. 2010, p. 221).

Long-term oriented societies (high LTO score) value pragmatic virtues related to the future, such as savings and adapting to changing circumstances (Hofstede et al. 2010, p. 239). Firms invest in building up strong market positions even at the expense of poor short-term performance. Managers are allowed time and resources to make sustained efforts with an aim to serve stakeholders and future generations (Hofstede et al. 2010, p. 244). Young people are taught the importance of savings and learn to put money aside for future uses.

Short-term oriented societies (low LTO score)—for example, Arab countries—value virtues related to the past and present such as national pride, respect for tradition, preservation of public images, and fulfilling social obligations (Hofstede et al. 2010, p. 239). Young people, affected by a culture of consumption, often live beyond their means. Corporate managers focus on the immediate results and are judged by their short-term performance. The cost of short-term decisions in terms of "pecuniary considerations, myopic decisions, work process control, hasty adoption, and quick abandonment of novel ideas" is evident (Hofstede et al. 2010, pp. 244–45).

The economic success of Taiwan, South Korean, Singapore, Hong Kong and Japan in the early 1990s highlights the value of their LTO cultures. These countries encourage thrift, perseverance (Anderson et al. 2011), and support entrepreneurial activities. People are persistent in the pursuit of their goals and are less tolerant of questionable business activities (Cohen et al. 1996). Their thrift translates into higher savings and growth rates (Hofstede et al. 2010, pp. 38, 263–65) and the availability of capital for reinvestment. Firms are less likely to rely heavily on debt financing. The sense of shame prevalent in LTO societies encourages interrelatedness through social contacts and stresses the importance of keeping commitments (Hofstede et al. 2010, pp. 243–44). Firms have strong ethical motivations to pay debts on time, and the severity of agency conflicts is inherently lower given their focus on stakeholders and future generations.

In the light of this discussion, a high LTO score is expected to make upgrades more likely and downgrades less likely.

3. Models and Variables

3.1. Estimation Model

This study applies a survival analysis framework (Allison 1995) and develops Cox's dynamic hazard model (Cox 1972) to examine the effect of culture on the probability of corporate rating migrations. Previous studies suggest that upgrade and downgrade follow different dynamics (Figlewski et al. 2012; Dang and Partington 2014). Thus, upgrade and downgrade are treated as competing risks, and separate models are estimated for these two migration outcomes.

Rating observations are arranged in event time (gap time) risk sets; each risk set includes all the firms that are at risk of a migration of interest at event time *t*. The clock is reset when a firm is assigned a rating grade. An upgrade (downgrade) is treated as a migration of interest (a censored observation) when estimating the upgrade model, and vice versa. Ratings which started before the beginning of the study or ended after the end of the study are also treated as censored. The survival time of a rating is the time a firm maintains a rating grade measured from the time it enters the rating grade subsequent to the start of the study until the time it either migrates to another rating grade or becomes censored.

Cox's hazard model is the premier technique in survival analysis. The main attraction of the estimated model is that it is convenient to handle repeated migrations. A typical firm experienced several migrations during the study period, which may lead to dependence among ratings. Accounting for repeated migrations is important as 42.6% (18%) of the firms in the main sample experienced between 2 and 15 downgrades (upgrades); the most volatile firms experienced 26 downgrades (25 upgrades). This highlights the need to consider unobserved heterogeneity in the underlying hazard of a rating change and to ensure that a firm is not considered at risk of a rating change before all previous rating changes have already occurred (Hosmer et al. 2008).

The problem arising from repeated migrations is minimized in two ways. First, the Wei et al. (1989)'s method for multiple failure time data is applied to account for the dependence among ratings of the same firm. The advantage of this method is that it does not require any assumptions about the nature or the structure of the dependence (Allison 1995, p. 242). Second, the estimation of the stratified hazard model, as in as Dang and Partington (2014), takes into account the sequence of rating changes. Each stratum includes ratings that have the same number of prior changes. The underlying (baseline) hazard of a rating change differs according to the number of prior changes whereas the effect of a variable is assumed to be the same across strata.

Another attractive feature of the estimated hazard model is that it accommodates time-varying outlook and rating age. A firm may be assigned a negative, positive, stable or developing outlook while its rating remains unchanged. A negative (positive) outlook signals the deterioration (improvement) in a firm's credit quality, and indicates S&P's opinion regarding the potential direction of a long-term credit rating over the intermediate term (six months to two years) (S&P RatingsDirect 2009). Estimating stratified dynamic hazard models with time-varying variables on a sample of 17,109 ratings (5360 firms) across 50 countries over a 26-year period creates substantial computational challenges in this study.

The hazard of a rating change *m* in stratum *s* is given by the product of the underlying (baseline) hazard $h_{(0,s)}(t)$ and the effect of the risk factors (covariates), and can be expressed as follows:

$$h_{m,s}(t, Z, Z(t)) = h_{(0,s)}(t) \exp[Z_j^m \beta_j + Z_p^m(t)\beta_p],$$
(1)

where: $h_{m,s}(t, Z, Z(t))$ is the migration hazard of rating *m* in stratum *s* at time *t* given its time-fixed covariate vector Z_j^m and its time-varying covariate vector $Z_p^m(t)$; $h_{(0,s)}(t)$ is the baseline hazard of a migration in stratum *s* at time *t*; β_p is the vector of estimated coefficients for time-varying covariates $Z_p^m(t)$; and β_j is the vector of estimated coefficients for time-fixed covariates Z_j^m .

The full rating migration model is estimated by multiplying together the individual likelihood functions for all the strata in the sample.

3.2. Variables

The definitions of variables employed, the data sources, and the references to examples of the relevant literature are given in Table 1.

The first group of variables includes Hofstede's five culture values and two of Schwartz's culture values, as discussed in the previous section.

The second group of variables accounts for the current rating (*current rating grade*), its proximity to investment/speculative rating threshold (*dummy investment boundary*, *dummy junk boundary*), rating history (*logarithm of age since first rated*, *dummy lag one downgrade*, *lag one duration*, *dummy prior fallen angel*, *dummy large downgrade*, *dummy large upgrade*, *rating volatility*) and time-varying rating outlook (*dummy negative outlook*, *dummy positive outlook*). Previous studies on corporate rating migrations have widely documented that the direction of lagged rating changes.¹¹ The use of lagged dependent variable (*lag one duration*) and important control variables (*dummy lag one downgrade*, *dummy negative outlook*, *dummy positive outlook*, *dummy lag one downgrade*, *dummy negative outlook*, *dummy positive outlook*. If the use of lagged dependent variable (*lag one duration*) and important control variables (*dummy lag one downgrade*, *dummy negative outlook*, *dummy positive outlook*). Let the use of lagged dependent variable (*lag one duration*) and important control variables (*dummy lag one downgrade*, *dummy negative outlook*, *dummy positive outlook*) minimizes endogeneity concerns, Li (2016).

¹¹ See, for example, Altman and Kao (1992); Carty and Fons (1994); Altman (1998); Bangia et al. (2002); Lando and Skodeberg (2002); Vazza et al. (2005b); Güttler and Wahrenburg (2007); Figlewski et al. (2012); Dang and Partington (2014).

Table 1. Definitions of variables used in the analysis.

Variable	Definition	References
Hofstede's culture dimensions Hofstede (1980) conducted surveys with Hofstede et al. (2010), the scores on the for Value Survey conducted in 23 countries is	IBM employees in over 50 countries and used the survey responses to identify our dimensions were listed for 76 countries. The fifth dimension, <i>long-term</i> versu n 1987, and extended to 93 countries in 2010 by Michael Minkov (Hofstede et a	four national culture dimensions that were virtually uncorrelated. In <i>s short-term orientation,</i> was introduced by Michael Bond in his Chinese 1. 2010)
Power distance index	Power distance index expresses the degree to which members of a society accept and expect that power and authority is distributed unequally.	Licht et al. (2005); Hope et al. (2008); Hofstede et al. (2010); Fidrmuc and Jacob (2010); Kanagaretnam et al. (2011); Zheng et al. (2012); Paredes and Wheatley (2017)
Individualism vs. collectivism	Individualism encourages the pursuit of personal interests, autonomy and an active determination of one's destiny. Collectivism stresses conformity and adherence to societal norms and regulations	Chui et al. (2002); Licht et al. (2005); Tsakumis et al. (2007); Fidrmuc and Jacob (2010); Han et al. (2010); Hofstede et al. (2010); Shao et al. (2010); Kanagaretnam et al. (2011); Zheng et al. (2012); Li et al. (2013); Shao et al. (2013)
Masculinity vs. femininity	Masculine countries strive for a performance society and value assertiveness, material accomplishment, ambition, competition and success. Feminine countries strive for a welfare society and value cooperation, modesty, caring for the weak and quality of life	Vitell et al. (1993); Chui et al. (2002); Licht et al. (2005); Hofstede et al. (2010); Anderson et al. (2011); Kanagaretnam et al. (2011); Zheng et al. (2012)
Uncertainty avoidance index	The uncertainty avoidance index expresses the degree to which the members of a society feel uncomfortable with uncertainty and ambiguity.	Licht et al. (2005); Kwok and Tadesse (2006); Ramirez and Tadesse (2009); Tsakumis et al. (2007); Hope et al. (2008); Fidrmuc and Jacob (2010); Han et al. (2010); Hofstede et al. (2010); Anderson et al. (2011); Kanagaretnam et al. (2011); Zheng et al. (2012); Li et al. (2013)
Long-term vs. short-term orientation	Long-term oriented cultures are oriented toward the future and value perseverance and thrift. Short-term oriented societies foster virtues related to the past and present.	Cohen et al. (1996); Hofstede et al. (2010); Anderson et al. (2011)
Schwartz's culture dimensions Schwartz (1994) collected survey data fro and <i>hierarchy</i> are employed in this study	m school teachers and university students in more than 60 countries. He classifi	ed national cultures into six dimensions. Two dimensions embeddedness
Embeddedness (conservatism)	Embedded cultures value social relationships, emphasize maintaining the status quo and restraining actions that may disrupt in-group solidarity and traditional order	Johnson and Lenartowicz (1998); Chui et al. (2002); Licht et al. (2005); Shao et al. (2010); Zheng et al. (2012)
Hierarchy	Hierarchical cultures view the unequal distribution of power, roles, and wealth as legitimate and even desirable.	Chui et al. (2002); Licht et al. (2005); Zheng et al. (2012)

Table 1. Cont.

11 of 27

Variable	Definition	References
S&P's rating data: Source: Standard & Poo	r's Ratings Xpress	
Current rating grade	The current rating (start rating) for the rating transition being observed.	Carty and Fons (1994); Figlewski et al. (2012); Dang and Partington (2014)
Investment rating boundary	The dummy takes the value of one if the current rating is in the investment grade boundary (BBB–, BBB, BBB+) or zero otherwise	Carty and Fons (1994); Johnson (2004); Dang and Partington (2014)
Junk rating boundary	The dummy takes the value of one if the current rating is in the speculative (junk) grade boundary (BB–, BB, BB+) or zero otherwise	Carty and Fons (1994); Johnson (2004); Dang and Partington (2014)
Logarithm of age since first rated	Age since first rated is a time-varying variable measuring the duration since a firm was first rated. This variable is updated whenever a migration of interest occurs in the sample	Altman (1998); Figlewski et al. (2012); Dang and Partington (2014)
Dummy lag one downgrade	The dummy takes the value of one if the lag one rating ends with a downgrade, and zero otherwise	Carty and Fons (1994); Lando and Skodeberg (2002); Bangia et al. (2002); Figlewski et al. (2012); Dang and Partington (2014)
Lag one duration (years)	The duration of the rating immediately preceding the current rating	Carty and Fons (1994); Lando and Skodeberg (2002); Dang and Partington (2014)
Dummy prior fallen angel	This variable takes the value of one if a firm had experienced a downgrade from an investment-grade rating to a speculative-grade rating as of the start of the current rating, and zero otherwise	Mann et al. (2003); Vazza et al. (2005a); Güttler and Wahrenburg (2007); Dang and Partington (2014)
Dummy large downgrade	This variable takes the value of one if a firm had experienced a big downgrade of at least three rating notches as of the start of the current rating, and zero otherwise	Carty and Fons (1994); Dang and Partington (2014)
Dummy large upgrade	This variable takes the value of one if a firm had experienced a big upgrade of at least three rating notches as of the beginning of the current rating, and zero otherwise	Dang and Partington (2014)
Rating volatility	This is the average number of migrations per year over a firm's rating history. It is calculated as the number of migrations a firm had experienced as of the beginning of the current rating divided by <i>age since first rated</i> .	Dang and Partington (2014)
S&P's outlook Source: Standard & Poor's I S&P issues an outlook to indicate its opinion can be positive (rating may be raised), nega	Ratings Xpress n regarding the potential direction of a long-term credit rating over the interme ative (rating may be lowered), stable (rating unlikely to change), or developin	ediate term (six months-two years) (S&P RatingsDirect 2009). Outlooks og (rating may be raised/ lowered)
Dummy negative outlook	This time-varying variable takes the value of one if a firm was assigned a negative outlook by S&P, and zero otherwise.	Vazza et al. (2005a); Hill et al. (2010)
Dummy positive outlook	This time-varying variable takes the value of one if a firm was assigned a positive outlook by S&P, and zero otherwise.	Vazza et al. (2005a); Hill et al. (2010)

Variable	Definition	References
Macro-economic and financial conditions	Source: World Bank databases unless otherwise stated	
Dummy prior default	This dummy takes the value of one if a country where a firm resides had a foreign currency-denominated debt default prior to the start of the rating under study, and zero otherwise. Source: S&P Global Ratings' S&P Global Ratings' Credit Research (2013)	Mora (2006); Hill et al. (2010)
Dummy debt crisis	This variable takes a value of one if a rating commences during a period of sovereign debt/ banking crisis as listed in Manasse et al. (2003), Laeven and Valencia (2008), or De Paoli et al. (2009), and zero otherwise.	Ferri et al. (1999); Mora (2006)
Dummy OECD member	This variable takes a value of one if the country where a firm resides is a member of the OECD at the start of the current rating, and zero otherwise.	Ferri et al. (1999); Mora (2006)
Logarithm of GDP per capita	The logarithm of real GDP per capita	Ramirez and Tadesse (2009); Zheng et al. (2012); Figlewski et al. (2012); Shao et al. (2013); Li et al. (2013); Dang and Partington (2014)
Change in real GDP growth rate	The change in the real GDP growth rate over the year prior to the start of the rating.	Ferri et al. (1999); Mora (2006); Hill et al. (2010); Shao et al. (2013)
Change in inflation	The change in the inflation rate over the year prior to the start of the rating.	Ramirez and Tadesse (2009); Zheng et al. (2012)
Change in current account surplus/GDP	The change in the current account surplus or deficit divided by GDP	Ferri et al. (1999); Mora (2006); Hill et al. (2010)
Change in term trade	The change in terms of trade. The terms of trade effect equals capacity to imports less exports of goods and services in constant prices. Data are in constant local currency.	
Logarithm of ratio stock market capitalization/GDP	The logarithm of the ratio of stock market capitalization to GDP	Zheng et al. (2012); Li et al. (2013); Shao et al. (2013)
Return of world stock market index	The average return of the World-Datastream stock market index, which is calculated using daily data over a 63-trading day rolling window prior to the start of the rating under study. Source: Datastream	Hill et al. (2010)
Political rights and civil liberties Source: I The political risk rating comprises the score conflict, investment profile, law and order,	nternational Country Risk Guide database. s of 12 metrics including government stability, bureaucracy quality, corruption military in politics, religion in politics, and socioeconomics conditions.	n, democratic accountability, external conflict, ethnic tensions, internal
Dummy high political risk	This dummy takes a value of one if a country's political rating score is less than or equal to 40, and zero otherwise 40	
Dummy low political risk	This dummy takes a value of one if a country's political rating score is higher than or equal to 80, and zero otherwise.	
Table 1 defines the variables used in Cox	s's hazard models. The table provides the data sources, variable names, the defi	initions of the variables, and references to examples of the literature

that have used the variables.

The third group of variables takes into account country characteristics. Potential variables are identified from previous studies on rating migrations. The choices of variables eliminate multi-collinearity concerns and aim at creating a sample from a large number of cultures. Selected variables are proxies for changes in macro-economic conditions (*change in real GDP growth rate, change in inflation, change in current account surplus/GDP, change in term trade*), the level of economic development (*logarithm of GDP per capita, dummy OECD member*), the degree of stock market capitalization (*logarithm of ratio stock market capitalization/GDP*), the existence of a sovereign debt/banking crisis (*dummy debt crisis*), and a history of sovereign foreign currency debt defaults (*dummy prior default*). Selected variables also account for the performance of global stock market (*return of world stock market index*), and the political risk in each country (*dummy high political risk, dummy low political risk*). By definition, *dummy high political risk* capture institutional factors such as religion and law (Table 1). The use of these variables further minimizes endogeneity concerns.

The variables are of two types: static (time-fixed) and dynamic (time-varying). The time-varying variables (*dummy negative outlook, dummy positive outlook, logarithm of age since first rated*) are updated over the duration of a rating as its outlook changes or as a migration of interest occurs. The time-fixed variables take the values observed at the start or closest to the start of a rating observation and are not changed over the duration of a rating.¹² The values of static variables only change when the rating changes. The durations of ratings, particularly for downgrades, were small, as depicted in Figure 1. Thus, substantial changes in the values of static variables during the duration of a rating were unlikely.

3.3. Samples

Since the data on outlook and the last rating change are required, only firms having at least one rating outlook and experiencing at least one prior migration during the study period are included in the final dataset. The main sample, sample A(H), includes 17,109 ratings of 5360 firms from 50 countries. This sample consists of ratings which have no missing data on Hofstede (H)'s five culture values and control variables.

Several robustness tests are conducted using alternative samples, alternative study periods, and alternative culture measures. In the first test, Schwartz's numeric scores for *embeddedness* and *hierarchy* are used as alternative measures of Hofstede's numeric scores for *collectivism* (the opposite pole of *individualism*) and *power distance index*, respectively. The requirement of data availability for Schwartz (S)'s culture scores causes a small reduction in the size of sample A(H), resulting in sample A(H-S) of 16,966 ratings. Each rating in sample A(H-S) has no missing data for three of Hofstede (H)'s culture values (*masculinity, uncertainty avoidance index*, and *long-term orientation*), two of Schwartz (S)'s culture values (*embeddedness* and *hierarchy*), and all control variables.

In the second robustness test, U.S. firms are excluded from samples A(H) and A(H-S), which results in samples B(H) and B(H-S), respectively. Sample B(H) (with Hofstede's culture values) includes 4745 ratings of 1717 firms from 49 countries. Sample B(H-S) (with Hofstede and Schwartz's culture values) includes 4602 ratings.

In the third test, ratings in sample A(H) are pooled across countries at the time they were (were not) experiencing a debt/a banking crisis, as defined in Manasse et al. (2003), Laeven and Valencia (2008), and De Paoli et al. (2009). Restricting the study period to crisis and non-crisis times results in samples C(H-1) and C(H-2), respectively. Sample C(H-1) (crisis times) includes 3927 ratings of 2088 firms from 32 countries. Sample C(H-2) (non-crisis times) includes 13,182 ratings of 4614 firms from 50 countries. Excluding U.S firms from sample C(H-2) results in sample C(H-3) of 3714 ratings (1535 firms).

¹² Most static variables are updated annually. *Return of world stock market index* is calculated using daily data over a 63-trading day rolling window prior to the beginning of each rating. *Dummy OECD member, dummy debt crisis and dummy prior default* are updated at the beginning of each rating.

3.4. Statistics

A numerical rating scale is employed to represent S&P's alphabetical ratings, ranging from AAA as 21 to C as 1. The descriptive statistics of the current rating grades (*start rating*) for observations in the three samples with Hofstede's culture scores are presented in Panel A of Table 2. On average, firms in all three samples had a speculative grade median rating (BB for the whole sample and BB+ for the sample excluding U.S firms). During crisis times the median rating dropped by two notches from BB to B+, suggesting a substantial decline in credit quality.

The descriptive statistics of the survival times for downgrades and upgrades in the three samples with Hofstede's culture scores are respectively given in Panel B and Panel C of Table 2. The median survival time was 1 year (1.74 years) for downgrades (upgrades) in the entire sample A(H), and about 0.38 year (1.01 year) for downgrades (upgrades) in the crisis sample C(H-1).

Downgrades (upgrades) account for 60.85% (23.48%), 56.08% (23.1%) and 46.47% (13.3%) of the entire sample A(H), non-U.S. sample B(H), and crisis sample C(H-1), respectively. The surprisingly low frequency of downgrades in the crisis sample is due to the lower than average frequency of downgrades of U.S. firms during crisis times. This is not too surprising as the study does not consider financial sector and thus does not examine accelerated downgrades which occurred to a large number of U.S. financial institutions during the financial crisis 2007–2009. Additional analysis shows that for non-U.S. firms, the frequency of downgrades is comparable (56%) across business cycles. For U.S. firms, the frequency of downgrades is markedly higher in non-crisis times (68.63%) than in crisis times (43.23%). For both U.S. and non-U.S. firms, upgrades occur much less often in crisis times.

	Panel A: Statistics of S&P's	Numerical Rating Grades											
	Sample A(H): All Firms	Sample B(H): Non-US Firms	Sample C(H-1): Crisis Sample										
Sample size	17,109	4745	3927										
Mean	10.29	11.08	8.79										
Median	10 (BB)	11 (BB+)	8 (B+)										
Std dev	4.18	4.22	3.9										
Min	1 (C)	1 (C)	2 (CC)										
Max	21 (AAA)	21 (AAA)	20 (AA+)										
	Panel B: Statistics of Survival Time for Downgrades												
	Sample A(H): All Firms	Sample B(H): Non-US Firms	Sample C(H-1): Crisis Sample										
Number of downgrades	10,411	2661	1825										
Frequency of downgrades	60.85%	56.08%	46.5%										
Mean (years)	1.82	1.55	0.66										
Median (years)	1	0.92	0.38										
Std dev	2.25	1.78	0.76										
Min (years)	~0	0.01	0.01										
Max (years)	23.43	14.28	7.41										
	Panel C: Statistics of Sur	vival Time for Upgrades											
	Sample A(H): All Firms	Sample B(H): Non-US	Sample C(H-1): Crisis										
	•	Firms	Sample										
Number of upgrades	4018	1096	521										
Frequency of upgrades	23.48%	23.1%	13.3%										
Mean (years)	2.25	1.99	1.34										
Median (years)	1.74	1.55	1.01										
Std dev	1.96	1.63	1.32										
Min (years)	0.02	0.02	0.04										
Max (years)	19.72	10.3	10.3										

Table 2. Statistics of	f rating grac	des and	survival	time.
------------------------	---------------	---------	----------	-------

Panel A of Table 2 shows the statistics of rating grades, Panel B and Panel C present the statistics of survival time for downgrades and upgrades, respectively, in the whole sample A(H), non-U.S. sample B(H), and crisis sample C(H-1). For brevity reasons, this table presents the statistics for ratings in the three samples with Hofstede's culture values.

The histograms of survival times for downgrades and upgrades in sample A(H) (the entire sample) and sample C(H-1) (crisis times) are depicted in Figure 1. Overall, downgrades had a shorter survival time than upgrades and were heavily massed in durations shorter than a year. About 50% of downgrades during the study period and about 80% of downgrades during crisis times retained their current rating for less than a year.





Panel B: Distribution of ratings by duration in crisis time



Figure 1. Distribution of rating observations by duration (survival time). Panel A and Panel B of Figure 1 shows the distribution of migrated ratings by duration (survival time) for the entire sample A(H) and crisis sample C(H-1), respectively. Duration is the length of time a firm stays in a rating grade measured from the time it enters the rating grade subsequent to the beginning of the study until the time it either migrates to another rating grade or becomes censored. A typical firm contributes multiple ratings to the dataset. Sample A(H) and sample C(H-1) includes 17,109 ratings and 3927 ratings, respectively.

Panels A and B of Table 3 give the descriptive statistics of the numeric scores on Hofstede's five culture values and Schwartz's two culture values for firms in samples A(H) and B(H), respectively. There is a wide range of variation in the scores for each of Hofstede's culture values. Compared with firms in sample A(H), non-U.S. firms in sample B(H) show characteristics of larger power distance (*PDI*), more feminine (*MAS*), more collectivistic (*IDV*), stronger risk-avoiding (*UAI*), and more long-term oriented (*LTO*) cultures.

Panel A: Descriptive Statistics of	Culture Val	Panel A: Descriptive Statistics of Culture Values for Samples A(H) and A(H-S) of All Firms											
	Mean	Median	Std Dev	Min	Max								
Hofstede (H) culture values (N = 17,109)													
Power distance index (PDI)	42.7	40	10.43	11	104								
Individualism vs. collectivism (IDV)	83.26	91	17.64	12	91								
Masculinity vs. femininity (MAS)	60.15	62	10.4	5	110								
Uncertainty avoidance index (UAI)	50.28	46	13.46	8	112								
Long-term vs. short-term orientation (LTO)	32.91	26	15.99	13	100								
Schwartz (S) culture values (N = 16,966)													
Embeddedness	3.62	3.67	0.16	3.03	4.35								
Hierarchy	2.34	2.37	0.17	1.49	3.23								
Panel B: Descriptive Statistics of C	ulture Value	s for Sample B(H) and B(H-S) o	of Non-U.S. Fi	rms								
	Mean	Median	Std Dev	Min	Max								
Hofstede (H) culture values (N = 4745)													
Power distance index (PDI)	49.75	39	17.98	11	104								
Individualism vs. collectivism (IDV)	63.08	71	23.63	12	90								
Masculinity vs. femininity (MAS)	55.32	56	18.91	5	110								
Uncertainty avoidance index (UAI)	61.44	53	21.94	8	112								
Long-term vs. short-term orientation (LTO)	50.93	51	21.74	13	100								
Schwartz (S) culture values (N = 4602)													
Embeddedness	3.49	3.46	0.27	3.03	4.35								
Hierarchy	2.25	2.22	0.32	1.49	3.23								

Table 3. Descriptive statistics of national culture scores.

Table 3 shows the descriptive statistics for scores of Hofstede (H)'s and Schwartz (S)'s culture values for the whole samples and non-U.S. samples. The definitions of the culture variables are as in Table 1.

4. Results

Separate models are estimated for downgrades and upgrades. Overall, culture values have significant effects on downgrades and upgrades after accounting for time-varying rating outlook, various aspects of rating history, macro-economic conditions and political risks.

4.1. Models for the Whole Sample (Samples A(H) and A(H-S))

The results of the models for the whole sample are given in Table 4. For each migration outcome, three models are estimated and three sets of results are given: first, using dummy variables where each dummy is set according to a country's score on Hofstede's culture value relative to the mean (model 1); second, using numeric scores for Hofstede's culture values (model 2); and, third using numeric scores for Hofstede's culture values (model 3).

As presented in Table 4, most of the retained variables are significant at the 10% level or better based on a Wald chi-square test. Culture values have stronger effects on downgrades than upgrades. The downgrade model with Hofstede and Schwartz's scores (model 3) features the significance of all five culture measures. The sign of a significant culture variable in model (3) is consistent with its sign (if significant) in model (1) and model (2).

The three models for downgrades feature one common culture value, which is *long-term orientation* (*LTO*)/*dummy LTO*, and its sign is as expected. Being in a *LTO* country reduces the downgrade probability by 25.2% (model 1) whereas a one-unit increase in *LTO* score makes a downgrade 0.9% (model 2) or 1.1% (model 3) less likely.¹³

Power distance index (PDI) is significant in model (2) for downgrades while *hierarchy* is present in model (3) for both downgrades and upgrades. Schwartz's score of *hierarchy* has a stronger effect than Hofstede's score of *PDI*. For example, a one-unit increase in *hierarchy* makes a downgrade 29.8% less

¹³ Subtracting one from the hazard ratio (HR) gives the change in risk for a one-unit change in the independent variable. *Dummy LTO*'s HR of 0.748 represents a 25.2% reduction in downgrade risk for firms in a *LTO* country (model 1). *Dummy LTO* (model 1) has a stronger impact than *LTO* (models 2 and 3). A larger effect of *dummy LTO* is not unusual in hazard modelling, often because a switch from short-term to long-term orientation represents a substantial change.

likely and an upgrade 73.7% more likely (model 3). A one-unit increase in *PDI* reduces the downgrade hazard by only 1% (model 2). This is not surprising as the score of Hofstede's *PDI* value for a country is much higher than the respective score of Schwartz's *hierarchy* value (Table 3). The risk being changed is very small for a one-unit increase in *PDI*.

Apart from *LTO* and *PDI/hierarchy*, models (2) and (3) for downgrades share another common culture value which is *uncertainty avoidance index (UAI)*. As expected, firms residing in strong UA countries have unfavorable experiences toward downgrades.

Individualism (IDV)/ Dummy IDV is not significant in models (1) and (2). However, *embeddedness* is significant in model (3) for downgrades. *Embeddedness* is comparable to *collectivism* which is the opposite pole of *individualism*. As expected, firms in *embeddedness* cultures are less likely to experience a downgrade (model 3).

For upgrades, *hierarchy* is the only significant variable in model (3) (as discussed above) while *masculinity* (*MAS*)/*dummy MAS* is the only significant variable in models (1) and (2). *MAS*'s positive sign in models (1) and (2) (for upgrades) and model 3 (for downgrades) suggests that firms in *MAS* cultures are more likely to experience rating changes in either direction.

With regard to control variables, the following discussion focuses on those significant in all three models for downgrades (upgrades). Firms rated around the speculative rating boundary (BB+/BB/BB) are more likely to become rising stars. Firms rated close to the investment rating threshold (BBB+/BBB/BBB-) or those with a positive outlook have lower downgrade risk and higher upgrade probability. Consistent with prior literature, issuers with a lagged downgrade or a negative outlook exhibit a strong tendency to travel downward on the rating spectrum. Older firms, firms with a high current rating, a volatile rating history or those which have experienced a fallen angel event tend to retain the current rating grade. Firms with a longer lagged rating or those with a prior substantial downgrade are more likely to experience a subsequent migration.

With regard to macro-economic environment, a strong global equity market is associated with a lower rating volatility, raising the probability that a firm will retain its current rating grade. Firms in OECD countries or countries with strong current account balances tend to go up the rating scales. Firms in countries high on inflation or countries with equity market-based economies experience fewer upgrades. A high GDP growth rate, a strong current account balance, and an emergence from a sovereign default lower the probability of corporate downgrades.

In terms of political risk, the effect of *dummy low political risk* is in contrast to initial expectations. Firms in countries with low political risk are more likely to be downgraded and less likely to be upgraded. Perhaps this surprising effect is due to the sample being dominated by U.S. firms.

A natural question is whether the impact of culture remains robust when the sample is restricted to non-U.S. firms and the study period is restricted to crisis/ non-crisis times.

4.2. Robustness Tests

4.2.1. Models for Non-U.S. Firms (Samples B(H) and B(H-S))

The results of the models for non-U.S. firms are given in Table 5. The evidence in favor of including culture variables is weaker in the upgrade models and stronger in the downgrade models (compared with the respective models in Table 4).

For downgrades, *dummy LTO/LTO* remains significant in three models. Consistent with the results presented in Table 4, firms in *LTO* cultures are less likely to be downgraded. Model (3) for downgrades features only four significant culture values compared with five significant culture values in the respective model for the whole sample (Table 4). *UAI* is no longer present in model (3) whereas *dummy UAI* and *dummy large PDI* become significant in model (1). The negative sign of *dummy large PDI* (model 1) is consistent with the negative sign of *PDI* (model 2), *hierarchy* (model 3) and the results presented in Table 4. Similarly, the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *dummy UAI* (model 1) is consistent with the positive sign of *UAI* (model 2).

For upgrades, model (1) does not feature any significant culture dummies. In model (3), *hierarchy* is no longer significant and is replaced by *MAS*. The positive effect of *MAS* on upgrades (models 2 and 3) and on downgrades (model 3) is as expected. Firms in *MAS* cultures are more volatile and exhibit a tendency to migrate from the current rating in either direction.

4.2.2. Models for Crisis and Non-Crisis Periods (Samples C(H-1), C(H-2), C(H-3))

Table 6 presents the results of the models for crisis sample C(H-1), non-crisis sample C(H-2), and non-crisis non-U.S. sample C(H-3). For reasons of brevity, only the results of the models with Hofstede's numeric scores are reported in Table 6.

Overall, the statistical significance of culture variables for crisis and non-crisis samples is weaker. Consistent with the results presented in Tables 4 and 5, *LTO* is the only culture value which is significant in the downgrade models for three samples. Firms in *LTO* cultures are less likely to be downgraded and during crisis times, are more likely to be upgraded.

UAI is only present in the downgrade model for the crisis sample. A strong UA culture (*UAI*) predisposes to more downgrades during crisis times. The effect is consistent with its effect on the whole sample (Table 4) and non-U.S. sample (Table 5). *PDI* is significant in the downgrade model for the crisis sample and the upgrade model for the non-crisis non-U.S. sample. A high PDI score reduces the downgrade hazard during crisis times and raises the upgrade probability of non-U.S. firms during non-crisis periods. The impact of *PDI* on downgrades during crisis times is consistent with its impact on the whole sample (Table 4) and non-U.S. sample (Table 5). *MAS* is not present in any downgrade models while *IDV* is not significant in any models.

The upgrade models for three samples do not share any significant common culture variable. However, *MAS* is significant in the upgrade models for the two non-crisis samples. Its positive sign suggests that firms in *MAS* cultures are more likely to be upgraded during non-crisis periods.

4.2.3. Other Robustness Tests

For each migration outcome, model (1) is re-estimated for three samples using dummy variables where each dummy is set according to a country's score on Hofstede's culture value relative to the *median* score. Three samples are used in this test: the whole sample (sample A(H)), non-U.S. sample (sample B(H)) and crisis sample (sample C(H-1)). Untabulated results across the three samples consistently show that firms in *LTO* countries (*dummy LTO*) are less likely to go down the rating scales. *Dummy LTO* is significant in both downgrade and upgrade models for the whole sample, and its opposite effects are as expected.

Dummy masculinity (MAS) is also significant in the models for the two large samples A(H) and B(H). The positive effect of *dummy MAS* on upgrades in sample A(H) is consistent with the result of the respective model when *dummy MAS* is set based on a country score relative to the mean score (Table 4, model 1 for upgrades).

While *dummy individualism/ IDV* is not significant in any models presented in Tables 4–6, it is significant when each dummy is set based on a country score relative to the median score. Firms in individualistic countries are more likely to be downgraded (for the whole sample and non-U.S. sample) and less likely to be upgraded (for the crisis sample).

	Culture D	Numeric C	ulture Sco	re (Hofstede, N	= 17,109)	Numeric Score (Hofstede & Schwartz, N = 16,966)						
	Downgrad	le (1)	Upgrade	e (1)	Downgra	ade (2)	Upgrad	le (2)	Downgra	ade (3)	Upgrad	de (3)
Variables	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR
Hofstede's national culture dimensions												
Power distance index (PDI)	NA		NA		-0.01 ***	0.99			NA		NA	
Individualism vs. collectivism (IDV)	NA		NA						NA		NA	
Masculinity vs. femininity (MAS)	NA		NA				0.004 **	1.004	0.006 ***	1.006		
Uncertainty avoidance index (UAI)	NA		NA		0.009 ***	1.009			0.004 ***	1.004		
Long-term vs. short-term orientation (LTO)	NA		NA		-0.009 ***	0.991			-0.011 ***	0.989		
Dummy large power distance index					NA		NA		NA		NA	
Dummy individualism					NA		NA		NA		NA	
Dummy masculine			0.17907 ***	1.196	NA		NA		NA		NA	
Dummy strong uncertainty avoidance					NA		NA		NA		NA	
Dummy long-term orientation	-0.29086 ***	0.748			NA		NA		NA		NA	
Schwartz's national culture dimensions												
Embeddedness	NA		NA		NA		NA		-0.274 **	0.76		
Hierarchy	NA		NA		NA		NA		-0.353 ***	0.702	0.552 ***	1.737
S&P's rating data												
Current rating grade	-0.01021 ***	0.99	-0.10447 ***	0.901	-0.009 **	0.991	-0.106 ***	0.9	-0.01 **	0.99	-0.108 ***	0.898
Investment rating boundary (BBB-/BBB/BBB+)	-0.23186 ***	0.793	0.16734 ***	1.182	-0.241 ***	0.786	0.167 ***	1.182	-0.24 ***	0.786	0.164 ***	1.179
(BBB / BBB / BBB / BB / BB / BB / BB /			0 14226 ***	1 153			0 143 ***	1 154			0 141 ***	1 152
Dummy negative outlook (time-varying)	0 29138 ***	1.338	-1 7834 ***	0.168	0 285 ***	1.33	-1 783 ***	0.168	0 282 ***	1.326	-1 776 ***	0.169
Dummy positive outlook (time-varying)	-1.73029 ***	0.177	1.206 ***	3.34	-1.721 ***	0.179	1.209 ***	3.35	-1.713 ***	0.18	1.207 ***	3.345
Logarithm of age since first rated												
(time-varving)	-1.57095 ***	0.208	-1.45531 ***	0.233	-1.556 ***	0.211	-1.439 ***	0.237	-1.552 ***	0.212	-1.434 ***	0.238
Dummy lag one downgrade	0.66822 ***	1.951			0.659 ***	1.933			0.665 ***	1.945		
Lag one rating duration	0.06974 ***	1.072	0.03321 ***	1.034	0.07 ***	1.072	0.033 ***	1.034	0.07 ***	1.072	0.033 ***	1.033
Dummy prior fallen angel event(s)	-0.13936 ***	0.87			-0.137 ***	0.872			-0.119 ***	0.888		
Dummy large downgrade	0.15244 ***	1.165	0.28276 ***	1.327	0.171 ***	1.187	0.285 ***	1.33	0.163 ***	1.177	0.281 ***	1.325
Dummy large upgrade									0.088 *	1.092		
Rating volatility	-0.05591	0.946	-0.13391 ***	0.875	-0.053	0.948	-0.134 ***	0.874	-0.051	0.951	-0.135 ***	0.874
Macro-economic and financial conditions												
Dummy prior default	-0.47471 ***	0.622			-0.536 ***	0.585			-0.442 ***	0.643	-0.24 *	0.787
Dummy debt crisis			-0.1048 *	0.901			-0.112 **	0.894				
Dummy OECD member			0.23915 ***	1.27	-0.249 ***	0.78	0.306 ***	1.358	-0.269 ***	0.764	0.212 *	1.236
Logarithm of GDP per capita											0.164 **	1.178
Change in real GDP growth rate	-0.02713 ***	0.973			-0.028 ***	0.972			-0.028 ***	0.973		
Change in inflation			-0.02144 ***	0.979			-0.022 ***	0.979			-0.022 ***	0.978
Change in current account surplus/GDP	-0.04317 ***	0.958	0.07859 ***	1.082	-0.04 ***	0.961	0.079 ***	1.082	-0.044 ***	0.957	0.069 ***	1.071
Change in term trade												
Logarithm of ratio stock market cap/GDP			-0.09241 ***	0.912	0.06 ***	1.062	-0.078 ***	0.925	0.08 ***	1.083	-0.161 ***	0.852
Return of world stock market index	-0.58725 ***	0.556	-0.24199 **	0.785	-0.525 ***	0.592	-0.228 *	0.796	-0.532 ***	0.588	-0.251 **	0.778

Table 4. Cox's regression models for rating changes of all firms.

	Culture D	Culture Dummy Mean (Hofstede, N = 17,109)					ore (Hofstede, N	= 17,109)	Numeric Sco	Numeric Score (Hofstede & Schwartz, N = 16,966)			
	Downgrad	Downgrade (1)		Upgrade (1)		Downgrade (2)		Upgrade (2)		Downgrade (3)		Upgrade (3)	
Variables	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	
Political risks Dummy low political risk Dummy high Political risk	0.09715 ***	1.102	-0.08271 **	0.921	0.067 **	1.069	-0.094**	0.911	0.072 ***	1.074	-0.066 *	0.936	
Events/ sample size Likelihood ratio χ^2	60.85% 8319.85	60.85% 8319.85 ***		23.48% 5022.41 ***		60.85% 8379.83 ***		23.48% 5017.3 ***		60.97% 8350.8 ***		23.44% 4977.2 ***	

Table 4. Cont.

Table 4 presents the results of Cox's hazard models for upgrades and downgrades in the entire samples. Downgrades (upgrades) were treated as events (censored) in the downgrade model and vice versa. The scores for Hofstede's culture values are coded as dummy variables (model 1) and numeric variables (model 2 and model 3). The scores for Schwartz's culture values are coded as numeric variables (model 3). In model (1), scores greater than or equal to the mean take a value of one or zero otherwise. Significant variables were retained in the model using the stepwise selection procedure. A variable has to be significant at the 0.25 level before it can be entered into the model, and a variable has to be significant at the 0.1 level for it to remain in the model. Parameter estimates are given first followed by the corresponding *p*-values based on Wald chi-square test, with ***, **, * representing significance at the 0.01, 0.05 and 0.10 levels, respectively, using the Wald test. Subtracting 1 from the hazard ratio (HR) gives the percentage change in the hazard for a one-unit change in the independent variable. The likelihood ratio is based on a comparison of the model with variables added versus a model without variables, and tests the null hypothesis that all variables included in the model have coefficients of 0.

	Culture D	ummy Mea	n (Hofstede, N = 4	4745)	Numeric C	Culture Sco	ore (Hofstede, N	= 4745)	Numeric Score (Hofstede & Schwartz, N = 4602)			
	Downgrade (1)		Upgrade	e (1)	Downgra	Downgrade (2)		e (2)	Downgrade (3)		Upgrad	e (3)
Variables	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR
Hofstede's national culture dimensions												
Power distance index (PDI)	NA		NA		-0.008 ***	0.993			NA		NA	
Individualism vs. collectivism (IDV)	NA		NA						NA		NA	
Masculinity vs. femininity (MAS)	NA		NA				0.003 *	1.003	0.004 ***	1.004	0.003 *	1.003
Uncertainty avoidance index (UAI)	NA		NA		0.004 ***	1.004						
Long-term vs. short-term orientation (LTO)	NA		NA		-0.007 ***	0.993			-0.008 ***	0.992		
Dummy large power distance index	-0.37814 ***	0.685			NA		NA		NA		NA	
Dummy individualism					NA		NA		NA		NA	
Dummy masculine					NA		NA		NA		NA	
Dummy strong uncertainty avoidance	0.17973 ***	1.197			NA		NA		NA		NA	
Dummy long-term orientation	-0.1631 ***	0.85			NA		NA		NA		NA	
Schwartz's national culture dimensions												
Embeddedness	NA		NA		NA		NA		-0.448 ***	0.639		
Hierarchy	NA		NA		NA		NA		-0.337 ***	0.714		
S&P's rating data												
Current rating grade	-0.10145 ***	0.904	-0.10145 ***	0.904			-0.103 ***	0.902			-0.102 ***	0.904
Investment rating boundary (BBB-/BBB/BBB+)	0.25624 ***	1.292	0.25624 ***	1.292	-0.292 ***	0.747	0.25 ***	1.284	-0.263 ***	0.769	0.219 ***	1.244
Junk rating boundary (BB-/BB/BB+)	0.17256 **	1.188	0.17256 **	1.188			0.17 **	1.185			0.16 **	1.173

Table 5. Cox's regression	n models for rating	changes of	non-U.S. firms
fuble of con b regression	i modelo for futility	crianges or	11011 0.0. 111110

	Culture D	oummy Mea	an (Hofstede, N =	4745)	Numeric C	Culture Sco	ore (Hofstede, N	= 4745)	Numeric Score (Hofstede & Schwartz, N = 4602)					
	Downgrade (1)		Upgrade	Upgrade (1)		Downgrade (2)		le (2)	Downgrade (3)		Upgrade (3)			
Variables	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR		
Dummy negative outlook (time-varying)	-1.61499 ***	0.199	-1.61499 ***	0.199	0.354 ***	1.425	-1.618 ***	0.198	0.341 ***	1.406	-1.576 ***	0.207		
Dummy positive outlook (time-varying)	1.27989 ***	3.596	1.27989 ***	3.596	-1.787 ***	0.167	1.282 ***	3.605	-1.767 ***	0.171	1.306 ***	3.691		
Logarithm of age since first rated (time-varying)	-2.77147 ***	0.063	-2.77147 ***	0.063	-2.299 ***	0.1	-2.765 ***	0.063	-2.335 ***	0.097	-2.742 ***	0.064		
Dummy lag one downgrade	-0.13508 **	0.874	-0.13508 **	0.874	0.598 ***	1.818	-0.133 **	0.876	0.575 ***	1.777	-0.13 **	0.878		
Lag one rating duration	0.15485 ***	1.167			0.167 ***	1.181	0.156 ***	1.169	0.184 ***	1.202	0.16 ***	1.173		
Dummy prior fallen angel event(s)	0.22181 **	1.248					0.227 **	1.255	0.174 **	1.19	0.225 **	1.253		
Dummy large downgrade														
Dummy large upgrade	-0.24784 **	0.78			-0.232 *	0.793			-0.189	0.828				
Rating volatility	-0.07356	0.929	-0.21696 **	0.805	-0.075	0.928	-0.215 **	0.807	-0.067	0.936	-0.204 **	0.816		
Macro-economic and financial conditions														
Dummy prior default	-0.45461 ***	0.635			-0.4 ***	0.671			-0.391 ***	0.676				
Dummy debt crisis	0.16082 **	1.174			0.172 ***	1.188			0.171 **	1.186				
Dummy OECD member									-0.191 **	0.826				
Logarithm of GDP per capita														
Change in real GDP growth rate	-0.02736 ***	0.973			-0.03 ***	0.971			-0.031 ***	0.97				
Change in inflation	-0.016 ***	0.984	-0.01629 ***	0.984	-0.017 ***	0.983	-0.016 ***	0.984	-0.017 ***	0.984	-0.017 ***	0.983		
Change in current account surplus/GDP			0.03483 ***	1.035			0.035 ***	1.036			0.037 ***	1.038		
Change in term trade														
Logarithm of ratio stock market cap/GDP														
Return of world stock market index	-0.60958 ***	0.544			-0.609 ***	0.544			-0.683 ***	0.505				
Political risks														
Dummy low political risk														
Dummy high Political risk														
Events / sample size	56.08%	6	23.1%	, n	56.08	%	23.1	%	56 37%		22.92%			
Likelihood ratio v^2	2858.04	- ***	1681.21	***	2874	***	1684 7	***	2860	***	2860 *** 1630 5 ***			

Table 5. Cont.

Table 5 presents the results of Cox's hazard models for rating upgrades and downgrades in the samples of non-U.S. firms. Downgrades (upgrades) were treated as events (censored) in the downgrade model and vice versa. The scores for Hofstede's culture values are coded as dummy variables (model 1) and numeric variables (model 2 and model 3). The scores for Schwartz's culture values are coded as numeric variables (model 3). In model (1), scores greater than or equal to the mean take a value of one and zero otherwise. Significant variables were retained in the model using the stepwise selection procedure. A variable has to be significant at the 0.25 level before it can be entered into the model, and a variable in the model has to be significant at the 0.1 level for it to remain in the model. Parameter estimates are given first followed by the corresponding *p*-values based on Wald chi-square test, with ***, ** representing significance at the 0.01, 0.05 and 0.10 levels, respectively, using the Wald test. Subtracting 1 from the hazard ratio (HR) gives the percentage change in the hazard for a one-unit change in the independent variable. The likelihood ratio is based on a comparison of the model with variables added versus a model without variables, and tests the null hypothesis that all variables included in the model have coefficients of 0.

	Crisi	s Sample (Hofstede, N = 392	27)	Non-Cris	sis Sample ((Hofstede, N = 13,1	182)	Non-Crisis Non-U.S. Sample (Hofstede, N = 3714)				
	Downg	rade	Upgra	de	Downgr	ade	Upgrad	le	Downgr	ade	Upgra	de	
Variables	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	Coefficient	HR	
Hofstede's national culture dimensions													
Power distance index (PDI)	-0.019 ***	0.981									0.00405 **	1.004	
Individualism vs. collectivism (IDV)													
Masculinity vs. femininity (MAS)							0.00442 ***	1.004			0.00326 *	1.003	
Uncertainty avoidance index (UAI)	0.009 ***	1.009											
Long-term vs. short-term orientation (LTO)	-0.005 ***	0.995	0.015 ***	1.015	-0.0068 ***	0.993			-0.00473 ***	0.995			
S&P's rating data													
Current rating grade	-0.055 ***	0.947	-0.281 ***	0.755			-0.09783 ***	0.907			-0.105 ***	0.9	
Investment rating boundary (BBB-/BBB/BBB+)	-0.417 ***	0.659	0.445 *	1.56	-0.21511 ***	0.806	0.18773 ***	1.207	-0.22824 ***	0.796	0.29756 ***	1.347	
Junk rating boundary (BB-/BB/BB+)			0.472 ***	1.603			0.15556 ***	1.168			0.17947 **	1.197	
Dummy negative outlook (time-varying)	0.406 ***	1.501	-1.512 ***	0.22	0.2374 ***	1.268	-1.93868 ***	0.144	0.26543 ***	1.304	-1.97581 ***	0.139	
Dummy positive outlook (time-varying)	-1.668 ***	0.189	1.327 ***	3.769	-1.72525 ***	0.178	1.17788 ***	3.247	-1.91491 ***	0.147	1.20023 ***	3.321	
Logarithm of age since first rated (time-varying)	-1.048 ***	0.351	-0.69 ***	0.502	-1.73288 ***	0.177	-1.47874 ***	0.228	-3.19282 ***	0.041	-2.99776 ***	0.05	
Dummy lag one downgrade	0.877 ***	2.403			0.61724 ***	1.854			0.58564 ***	1.796			
Lag one rating duration	0.066 ***	1.068			0.07298 ***	1.076	0.03799 ***	1.039	0.18513 ***	1.203	0.17943 ***	1.197	
Dummy prior fallen angel event(s)					-0.19549 ***	0.822					0.30926 ***	1.362	
Dummy large downgrade	0.194 ***	1.214			0.17731 ***	1.194	0.32772 ***	1.388			0.17525	1.192	
Dummy large upgrade					0.14199 ***	1.153					-0.26509 **	0.767	
Rating volatility					-0.107	0.899	-0.10982 **	0.896	-0.46349 ***	0.629	-0.19194 **	0.825	
Macro-economic and financial conditions													
Dummy prior default	-0.654 ***	0.52											
Dummy OECD member	-1.008 ***	0.365			0.28907 ***	1.335	0.2497 ***	1.284	0.42189 ***	1.525			
Logarithm of GDP per capita			0.564 ***	1.758									
Change in real GDP growth rate	-0.055 ***	0.946											
Change in inflation	-0.019 ***	0.981	-0.027 ***	0.973									
Change in current account surplus/GDP	-0.063 ***	0.939	0.142 ***	1.152			0.04871 ***	1.05					
Change in term trade			0.000002 ***	1			0.000037 ***	1			0.00003 ***	1	
Logarithm of ratio stock market cap/GDP					0.11571 ***	1.123	-0.10045 ***	0.904					
Return of world stock market index	-0.367 ***	0.693	-0.61 ***	0.544	-0.45234 ***	0.636			-0.47447 ***	0.622			
Political risks													
Dummy low political risk	-0.198 ***	0.82			0.11886 ***	1.126	-0.10071 **	0.904					
Dummy high Political risk													
Events/ sample size	46.47	%	13.3%	6	65.13%	6	26 53%		56.22%	6	26.17	/0	
Likelihood ratio χ^2	1825.3	***	697 **	F#	6808.65	***	4414.7 *	**	2367.02	***	1568 68 ***		
	1020.0		557		0000.00		1111.,		2007.02		1000.00		

Table 6. Cox's regression model for corporate rating changes: crisis vs. non-crisis periods.

Table 6 presents the results of Cox's hazard models for corporate ratings pooled across countries during the time they were (were not) experiencing a debt crisis or a banking crisis as defined in Manasse et al. (2003), Laeven and Valencia (2008), and De Paoli et al. (2009). For brevity reasons, only the results of the models with Hofstede's numeric scores are reported. Downgrades (upgrades) were treated as events (censored) in the downgrade model and vice versa. Significant variables were retained in the model using the stepwise selection procedure. A variable has to be significant at the 0.25 level before it can be entered into the model, and a variable has to be significant at the 0.1 level for it to remain in the model. Parameter estimates are given first followed by the corresponding *p*-values based on Wald chi-square test, with ***, **, * representing significance at the 0.01, 0.05 and 0.10 levels, respectively, using the Wald test. Subtracting 1 from the hazard ratio (HR) gives the percentage change in the hazard for a one-unit change in the independent variable. The likelihood ratio is based on a comparison of the model with variables added versus a model without variables, and tests the null hypothesis that all variables included in the model have coefficients of 0.

5. Conclusions

The informal constraints (social influences) that arise from the national culture in which a firm resides have a pervasive impact on managerial decision making and corporate credit risk, which in turn impacts on corporate ratings and rating changes. Some cultures may lead to more rating changes whereas other cultures may result in a directional predisposition for rating changes (downgrade *or* upgrade). This study is the first attempt to explore the effects of national culture values established by Hofstede et al. (2010) and Schwartz (1994) on the rating migration dynamics of 5360 firms across 50 countries over the period 1985–2010. The study enriches the literature by presenting empirical evidence that national culture provides a better explanation of corporate rating migrations. The effects of culture are significant after accounting for variables which have been found significant in previous studies on rating migrations. The study overcomes computational challenges in estimating stratified dynamic hazard models (with time-varying variables) for a large sample of 17,109 ratings.

Overall, the evidence in favor of including culture variables is generally stronger in the downgrade models or when numeric scores are employed to represent culture. The evidence of statistical significance is strongest for the effect of *long-term orientation*. Firms located in *long-term oriented* cultures (*LTO*) are less likely to be downgraded and in some cases, more likely to climb up the rating scales. The effect of *LTO* on downgrades is robust to alternative samples, alternative measures of culture and alternative study periods. There is some evidence that downgrades occur more often in strong *uncertainty-avoiding* countries and less often in large *power distance* (*hierarchy*) and *embeddedness* countries. *Masculinity* predisposes to a higher rating volatility, raising the probability of downgrades and upgrades.

This study is somewhat limited as apart from credit rating, other firm-specific data is not available. In the absence of variables such as firm size, it is *not* possible to examine the effects of culture on rating *levels*. This study therefore focuses on the effects of national culture on the probability of rating *migrations*. The estimated models of rating changes include the current rating, which captures firm-specific characteristics such as firm size and leverage, and time-varying rating outlooks, which signal the potential direction of a long-term credit rating over the intermediate term. Thus, the lack of firm-specific data such as firm size is not a substantial concern for this study.

This study emphasizes the need to understand the role of culture in managerial decision making and corporate policies, which feed into the rating *review* process. Studying culture helps enrich our understanding of corporate rating migration dynamics. This knowledge in turn can be helpful in developing predictive models of corporate rating changes across countries. The results of this study have practical implications to investors who use credit ratings to make investment decisions.

Acknowledgments: I would like to thank two anonymous reviewers, the editors, and Jean Helwege, Gady Jacoby, Andreas Knetsch, Bing Li, Graham Partington, as well as participants at the 2017 Cross Country Perspectives in Finance Symposium for helpful suggestions. Thanks are also due to George Griffiths for proofreading this article, and to the Knowledge Unlatched and the MDPI for supporting this manuscript. The rating data used in this study was purchased thanks to the generous funding from the University of Munich (LMU). I am grateful for Andreas Richter, Susanne Weber, and the Center for Advanced Studies at LMU for the funding support.

Conflicts of Interest: The author declares no conflicts of interest.

Appendix A

Appendix A outlines different channels over which national culture impacts on corporate rating migration dynamics.



Figure A1. Possible links between national culture and corporate rating migrations.

References

- Afego, Pyemo N. 2018. Index shocks, investor action and long-run stock performance in Japan: A case of cultural behaviouralism? *Journal of Behavioral and Experimental Finance* 18: 54–66. [CrossRef]
- Allison, Paul D. 1995. Survival Analysis Using SAS: A Practical Guide. Cary: SAS Institute Inc. ISBN 978-1-59994-884-3.
- Altman, Edward. 1998. The importance and subtlety of credit rating migration. *Journal of Banking and Finance* 22: 1231–47. [CrossRef]
- Altman, Edward I., and Duen L. Kao. 1992. The implications of corporate bond rating drifts. *Financial Analysts Journal* 48: 64–75. [CrossRef]
- Anderson, Christopher W., Mark Fedenia, Mark Hirschey, and Hilla Skiba. 2011. Cultural influences on home bias and international diversification by institutional investors. *Journal of Banking and Finance* 35: 916–34. [CrossRef]
- Ashraf, Badar N., Changjun Zheng, and Sidra Arshad. 2016. Effects of national culture on bank risk-taking behavior. *Research in International Business and Finance* 27: 309–26. [CrossRef]
- Bae, Sung C., Kiyoung Chang, and Eun Kang. 2012. Culture, corporate governance and dividend policy: International evidence. *The Journal of Financial Research* 35: 289–316. [CrossRef]
- Bangia, Anil, Francis X. Diebold, André Kronimus, Christian Schagen, and Til Schuermann. 2002. Rating Migrations and the Business Cycle, with Applications to Credit Portfolio Stress Testing. *Journal of Banking and Finance* 26: 445–74. [CrossRef]

- Barber, Brad M., and Terrance Odean. 2001. Boys will be boys: Gender, overconfidence and common stock investment. *Quarterly Journal of Economics* 116: 261–92. [CrossRef]
- Blume, Marshall E., Donald B. Keim, and Sandeep A. Patel. 1991. Returns and volatility of low-grade bonds, 1977–1989. *Journal of Finance* 46: 49–74. [CrossRef]
- Boubakri, Narjess, Ali Mirzaei, and Anis Samet. 2017. National culture and bank performance: Evidence from the recent financial crisis. *Journal of Financial Stability* 29: 36–56. [CrossRef]
- Byrnes, James P., David C. Miller, and William Schafer. 1999. Gender Differences in Risk Taking: A Meta-Analysis. *Psychology Bulletin* 75: 367–83. [CrossRef]
- Carty, Lee V., and Jerome S. Fons. 1994. Measuring Changes in Corporate Credit Quality. *Journal of Fixed Income* 4: 27–41. [CrossRef]
- Chang, Chih-Hsiang, and Shih-Jia Lin. 2015. The effects of national culture and behavioral pitfalls on investors' decision-making: Herding behavior in international stock markets. *International Review of Economics and Finance* 37: 380–92. [CrossRef]
- Chang, Kiyoung, and Abbas Noorbakhsh. 2009. Does national culture affect international corporate cash holdings? Journal of Multinational Financial Management 19: 323–42. [CrossRef]
- Chui, Andy, Alison E. Lloyd, and Chuck C. Y. Kwok. 2002. The determination of capital structure: Is national culture a missing piece to the puzzle? *Journal of International Business Studies* 33: 99–127. [CrossRef]
- Cohen, Jeffrey R., Laurie W. Pant, and David J. Sharp. 1996. A methodological note on cross-cultural accounting ethics research. *International Journal of Accounting* 31: 55–66. [CrossRef]
- Cox, David R. 1972. Regression models and life tables. *Journal of Royal Statistical Society Series B (Methodological)* 34: 187–220. [CrossRef]
- Dang, Huong, and Graham Partington. 2014. Rating migrations: The effects of rating history and time. *ABACUS* 50: 174–202. [CrossRef]
- De Mooij, Marieke. 2004. *Consumer Behaviour and Culture: Consequence for Global Marketing and Advertising*. Thousand Oaks: Sage. ISBN 9781412979900.
- De Paoli, Bianca, Glenn Hoggarth, and Victoria Saporta. 2009. Output Costs of Sovereign Crises: Some Empirical Estimates. Bank of England Working Paper No. 362. Available online: https://www.bankofengland.co.uk/-/media/boe/files/working-paper/2009/output-costs-ofsovereign-crises-some-emperical-estimates.pdf (accessed on 2 November 2018).
- Desender, Kurt A., Christian E. Castro, and Sergio A. Escamilla De Leon. 2011. Earnings management and cultural values. *American Journal of Economics and Sociology* 70: 639–70. [CrossRef]
- Doupnik, Timothy S., and George T. Tsakumis. 2004. A critical review of tests of Gray's theory of cultural relevance and suggestions for future research. *Journal of Accounting Literature* 23: 1–48.
- Ferri, Giovanni, Li-gang Liu, and Joseph E. Stiglitz. 1999. The procyclical role of rating agencies: Evidence from the East Asian crisis. *Economic Notes* 28: 335–55. [CrossRef]
- Fidrmuc, Jana P., and Marcus Jacob. 2010. Culture, agency costs and dividends. *Journal of Comparative Economics* 38: 321–39. [CrossRef]
- Figlewski, Stephen, Halina Frydman, and Weijian Liang. 2012. Modeling the Effects of Macroeconomic Factors on Corporate Default and Credit Rating Transitions. *International Review of Economics and Finance* 21: 87–105. [CrossRef]
- Fonseca, Ana R., and Francisco Gonzalez. 2008. Cross country determinants of bank income smoothing by managing loan-loss provisions. *Journal of Banking and Finance* 32: 217–28. [CrossRef]
- Gray, Sidney. 1988. Towards a theory of cultural influence on the development of accounting systems internationally. *ABACUS* 24: 1–15. [CrossRef]
- Gray, Sidney J., and Hazel M. Vint. 1995. The impact of culture on accounting disclosures: Some international evidence. *Asia Pacific Journal of Accounting* 21: 33–43. [CrossRef]
- Griffin, Dale, Omrane Guedhami, Chuck C. Y. Kwok, Kai Li, and Liang Shao. 2015. National Culture, Corporate Governance Practices and Firm Performance. Working Paper. Available online: https://memento.epfl.ch/ public/upload/files/Paper_Li.pdf (accessed on 2 November 2018).
- Guiso, Luigi, Helios Herrera, and Massimo Morelli. 2016. Cultural differences and institutional integration. *Journal of International Economics* 99: 97–113. [CrossRef]
- Güttler, André. 2011. Lead-lag Relationships and Rating Convergence among Credit Rating Agencies. *Journal of Credit Risk* 7: 95–119. [CrossRef]

- Güttler, André, and Mark Wahrenburg. 2007. The adjustment of credit ratings in advance of defaults. *Journal of Banking and Finance* 31: 751–67. [CrossRef]
- Han, Sam, Tony Kang, Stephen Salter, and Yong Keun Yoo. 2010. A cross-country study on the effects of national culture on earnings management. *Journal of International Business Studies* 41: 123–41. [CrossRef]
- Hill, Paula, Robert Brooks, and Robert Faff. 2010. Variations in sovereign credit quality assessments across rating agencies. *Journal of Banking and Finance* 34: 1324–43. [CrossRef]
- Hofstede, Geert. 1980. Culture's Consequences: International Differences in Work-Related Values. Beverly Hills: Sage. ISBN 9780803913066.
- Hofstede, Geert. 2001. Culture's Consequences: Comparing Values, Behaviors, Institutions and Organizations Across Nations. Thousand Oaks: Sage. ISBN 9780803973244.
- Hofstede, Geert, Geert Hofstede, and Michael Minkov. 2010. *Cultures and Organizations: Software of the Mind*, 3rd ed. New York: McGraw-Hill. ISBN 978-0071664189.
- Hope, Ole-Kristian. 2003. Firm-level disclosures and the relative roles of culture and legal origin. *Journal of International Financial Management and Accounting* 14: 218–48. [CrossRef]
- Hope, Ole-Kristian, Tony Kang, Wayne Thomas, and Yong Keun Yoo. 2008. Culture and auditor choice: A test of the secrecy hypothesis. *Journal of Accounting Public Policy* 27: 357–73. [CrossRef]
- Hosmer, David W., Stanley Lemeshow, and Susanne May. 2008. *Applied Survival Analysis: Regression Modeling of Time-to-Event Data*, 2nd ed. New York: Wiley. ISBN 978-0-471-75499-2.
- Husted, Bryan. 1999. Wealth culture and corruption. Journal of International Business Studies 3: 339–59. [CrossRef]
- Jensen, Michael, and William Meckling. 1976. Theory of the firm: Managerial behavior, agency costs and ownership structure. *Journal of Financial Economics* 3: 305–60. [CrossRef]
- Johnson, Richard. 2004. Rating agency actions around the investment grade boundary. *Journal of Fixed Income* 13: 25–37. [CrossRef]
- Johnson, James, and Tomasz Lenartowicz. 1998. Culture freedom and economic growth: Do cultural values explain economic growth? *Journal of World Business* 33: 332–56. [CrossRef]
- Jorion, Philippe, Charles Shi, and Sanjian Zhang. 2009. Tightening Credit Standards: The Role of Accounting Quality. *Review of Accounting Study* 14: 123–60. [CrossRef]
- Kanagaretnam, Kiridaran, Chee Yeow Lim, and Gerald Lobo. 2011. Effects of National Culture on Earnings Quality of Banks. *Journal of International Business Studies* 42: 853–74. [CrossRef]
- Kwok, Chuck C. Y., and Solomon Tadesse. 2006. National culture and financial systems. *Journal of International Business Studies* 37: 227–47. [CrossRef]
- Laeven, Luc, and Fabian Valencia. 2008. Systemic Banking Crises: A New Database. IMF Working Paper WP/08/224. Available online: https://www.imf.org/external/pubs/ft/wp/2008/wp08224.pdf (accessed on 2 November 2018).
- Lando, David, and Torben M. Skodeberg. 2002. Analyzing ratings transitions and rating drift with continuous observations. *Journal of Banking and Finance* 26: 423–44. [CrossRef]
- Leung, Kwok, Rabi S. Bhagat, Nancy R. Buchan, Miriam Erez, and Cristina B. Gibson. 2005. Culture and international business: Recent advances and their implications for future research. *Journal of International Business Studies* 36: 357–78. [CrossRef]
- Li, Zhichuan. 2016. Endogeneity in CEO power: A survey and experiment. *Investment Analysts Journal* 45: 149–62. [CrossRef]
- Li, Kai, Dale Griffin, Heng Yue, and Longkai Zhao. 2013. How does culture influence corporate risk-taking? Journal of Corporate Finance 23: 1–22. [CrossRef]
- Licht, Amir, Chanan Goldschmidt, and Shalom Schwartz. 2005. Culture, law and corporate governance. International Review of Law and Economics 25: 229–55. [CrossRef]
- Lobo, Gerald, Luc Paugam, Herve Stolowy, and Pierre Astolfi. 2017. The Effect of Business and Financial Market Cycles on Credit Ratings: Evidence from the Last Two Decades. *ABACUS* 53: 59–93. [CrossRef]
- Manasse, Paolo, Nouriel Roubini, and Axel Schimmelpfennig. 2003. Predicting Sovereign Debt Crisis. IMF Working Paper 03/221. Available online: https://www.imf.org/en/Publications/WP/Issues/2016/12/30/ Predicting-Sovereign-Debt-Crises-16951 (accessed on 2 November 2018).
- Mann, Christopher, David Hamilton, Praveen Varma, and Richard Cantor. 2003. *What Happens to Fallen Angels? A Statistical Review* 1982–2003. Moody's Investor Service Special Comment. Available online: https://www.moodys.com/sites/products/DefaultResearch/200200000425343.pdf (accessed on 2 November 2018).

- Mora, Nada. 2006. Sovereign credit ratings: Guilty beyond reasonable doubt? *Journal of Banking and Finance* 30: 2041–62. [CrossRef]
- Nickell, Pamela, William Perraudin, and Simone Varotto. 2000. Stability of ratings transitions. *Journal of Banking* and Finance 24: 203–22. [CrossRef]
- North, Douglass. 1990. Institutions, Institutional Change and Economic Performance. Cambridge: Cambridge University Press. ISBN 9780521397346.
- Pan, Yihui, Stephen Siegel, and Tracy Yue Wang. 2017. Corporate Risk Culture. *Journal of Financial and Quantitative* Analysis 52: 2327–67. [CrossRef]
- Paredes, Angel A. P., and Clark Wheatley. 2017. The influence of culture on real earnings management. *International Journal of Emerging Markets* 12: 38–57. [CrossRef]
- PricewaterhouseCoopers. 2008. Reward: A New Paradigm? Available online: https://www.pwc.com/gx/en/ banking-capital-markets/pdf/reward.pdf (accessed on 2 November 2018).
- Ramirez, Andrés, and Solomon Tadesse. 2009. Corporate Cash Holdings, Uncertainty Avoidance and the Multinationality of Firms. *International Business Review* 18: 387–403. [CrossRef]
- S&P Global Ratings' Credit Research. 2013. *Default Study: Sovereign Defaults and Rating Transition Data,* 2012 Update. New York: Alacra Store.
- S&P RatingsDirect. 2009. Use of CreditWatch and Outlooks. September 14. Available online: http://www.maalot. co.il/publications/MT20131212112758c.pdf (accessed on 2 November 2018).
- S&P RatingsDirect. 2013. Corporate Methodology. November 19. Available online: https://www.spratings.com/ scenario-builder-portlet/pdfs/CorporateMethodology.pdf (accessed on 2 November 2018).
- Schwartz, Shalom H. 1994. Beyond Individualism-Collectivism: New Dimensions of Values. In Individualism and Collectivism: Theory, Method and Application. Edited by Uichol Kim, Harry C. Triandis, Çiğdem Kâğitçibaşi, Sang-Chin Choi and Gene Yoon. Thousand Oaks: Sage. ISBN 978-0803957633.
- Shao, Liang, Chuck C. Y. Kwok, and Omrane Guedhami. 2010. National culture and dividend policy. *Journal of International Business Studies* 41: 1391–414. [CrossRef]
- Shao, Liang, Chuck C. Y. Kwo, and Ran Zhang. 2013. National culture and corporate investment. *Journal of International Business Studies* 44: 745–63. [CrossRef]
- Tang, Linghui, and Peter Koveos. 2008. A framework to update Hofstede's cultural value indices: Economic dynamics and institutional stability. *Journal of International Business Studies* 39: 1045–63. [CrossRef]
- Titman, Sheridan. 1984. The effect of capital structure on a firm's liquidation decision. *Journal of Financial Economics* 13: 137–52. [CrossRef]
- Tsakumis, George, Anthony Curatola, and Thomas Porcano. 2007. The relation between national cultural dimensions and tax evasion. *Journal of International Accounting, Auditing and Taxation* 16: 131–47. [CrossRef]
- Vazza, Diane, Devi Aurora, and Ryan Schneck. 2005a. *Crossover Credit: A 24-Year Study of Fallen Angel Rating Behavior*. New York: Standard & Poor's Global Fixed Income Research.
- Vazza, Diane, Edward Leung, Marya Alsati, and Mike Katz. 2005b. *Credit Watch and Ratings Outlooks: Valuable Predictors of Rating Behaviour*. New York: Standard & Poor's Global Fixed Income Research.
- Vitell, Scott, Saviour Nwachukwu, and James Barnes. 1993. The effects of culture on ethical decisions making: An application of Hofstede's typology. *Journal of Business Ethics* 12: 753–60. [CrossRef]
- Weaver, Gary R. 2001. Ethics Programs in Global Businesses: Culture's Role in Managing Ethics. *Journal of Business Ethics* 30: 3–15. [CrossRef]
- Wei, Lee-Jen, Danyu Lin, and Lisa Weissfeld. 1989. Regression analysis of multivariate incomplete failure time data by modelling marginal distributions. *Journal of the American Statistical Association* 84: 1065–73. [CrossRef]
- Zarzeski, Marilyn T. 1996. Spontaneous harmonization effects of culture and market forces on accounting disclosure practices. *Accounting Horizons* 10: 18–37.
- Zheng, Xiaolan, Sadok Ghoul, Omrane Guedhami, and Chuck C. Y. Kwok. 2012. National culture and corporate debt maturity. *Journal of Banking and Finance* 36: 468–88. [CrossRef]



© 2018 by the author. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (http://creativecommons.org/licenses/by/4.0/).