



Article Tree Inference: Response Time and Other Measures in a Binary Multinomial Processing Tree, Representation and Uniqueness of Parameters

Richard Schweickert * D and Xiaofang Zheng

Department of Psychological Sciences, Purdue University, West Lafayette, IN 47907, USA; zheng331@purdue.edu * Correspondence: schweick@purdue.edu

Abstract: A Multinomial Processing Tree (MPT) is a directed tree with a probability associated with each arc and partitioned terminal vertices. We consider an additional parameter for each arc, a measure such as time. Each vertex represents a process. An arc descending from a vertex represents selection of a process outcome. A source vertex represents processing beginning with stimulus presentation and a terminal vertex represents a response. An experimental factor selectively influences a vertex if changing the factor level changes parameter values on arcs descending from that vertex and no others. Earlier work shows that if each of two factors selectively influences a different vertex in an arbitrary MPT it is equivalent to one of two simple MPTs. Which applies depends on whether the two selectively influenced vertices are ordered by the factors or not. A special case, the Standard Binary Tree for Ordered Processes, arises if the vertices are ordered and the factor selectively influencing the first vertex changes parameter values on only two arcs. We derive necessary and sufficient conditions, testable by bootstrapping, for this case. Parameter values are not unique. We give admissible transformations for them. We calculate degrees of freedom needed for goodness of fit tests.

Keywords: Multinomial Processing Tree; response time; selective influence

1. Introduction

A Multinomial Processing Tree (MPT) consists of a finite set of vertices and a set of arcs, each arc being an ordered pair of vertices, such that there is no more than one directed path from one vertex to another, see Figure 1. Associated with each arc is a probability parameter and possibly other parameters. The sum of the probabilities on arcs descending from a vertex is 1. A source vertex is a vertex with no incoming arc, that is, a vertex that is not the second vertex of any arc. A terminal vertex is a vertex with no outgoing arc, that is, a vertex that is not the first vertex of any arc. The terminal vertices are partitioned into classes.

Multinomial Processing Trees are widely used to model cognitive processing in perception, memory, decision making, and movement tasks. For review and discussion, see [1–4]. In an MPT model of a task, each vertex represents a process, for example, an attempt to retrieve an item from memory. The process has possible outcomes, such as successful or unsuccessful retrieval. Each possible outcome is represented by a vertex with an arc directed to it from the vertex representing the process. The probability associated with the arc is the probability the corresponding outcome occurs after the process starts. When the task starts, with, say, presentation of a stimulus, processing begins at a source vertex. One of the outcomes of the source vertex occurs, with the probability associated with the corresponding arc. The second vertex of the arc is reached and the process represented by that vertex starts. Steps continue in this fashion until a terminal vertex is reached, at which a response is made. Responses are partitioned into classes, for example correct and



Citation: Schweickert, R.; Zheng, X. Tree Inference: Response Time and Other Measures in a Binary Multinomial Processing Tree, Representation and Uniqueness of Parameters. *Mathematics* 2022, *10*, 267. https://doi.org/10.3390/ math10020267

Academic Editor: Ronaldo Vigo

Received: 1 December 2021 Accepted: 13 January 2022 Published: 16 January 2022

Publisher's Note: MDPI stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.



Copyright: © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/).



incorrect. One class of responses is made at each terminal vertex, although several terminal vertices may be associated with the same response class.

Figure 1. A Multinomial Processing Tree. It is the Standard Binary Tree for Ordered Processes. A probability denoted by *P* labels each arc, also a measure denoted by *t*. See text for further notation.

Here we focus on the special case of binary trees, those in which exactly two arcs descend from every nonterminal vertex, as in Figure 1. Empirically, many MPTs are binary trees. For example, two most often used MPTs, the Process Dissociation Model [5] and the Quad Model [6] are binary trees. Binary trees are of theoretical interest also. Any MPT representing response probability can be reparameterized as a binary MPT [7,8]. Further, a way of representing binary MPTs as strings in a context-free language has been developed [9], facilitating analysis of model structure.

Multinomial Processing Tree models are often tested by goodness of fit. Although goodness of fit is a useful consideration, several authors have argued that it is not sufficient for evaluation of a model (e.g., [10,11]. Here we derive conditions that can serve as the basis for alternative tests of the MPT model in Figure 1. A distribution free test of each condition can be made with bootstrapping, in the manner of [12]. Goodness of fit is an evaluation of a model as a whole. Each condition here focuses on a particular parameter. If a test of a condition fails, one learns that the corresponding parameter does not behave well. A result here is useful for testing with goodness of fit. Evaluation of a goodness of fit statistic, such as G^2 , requires degrees of freedom and they are provided here.

An MPT may have several sources. Usually, these correspond to different conditions. For example, one source may be for stimuli that are verbs, another for nouns, and in a given condition only one source is relevant and the others can be ignored. We assume the MPTs under consideration have only one source.

The probability of a directed path from one vertex to another is the product of the probabilities associated with the arcs on the path. On any particular trial of the task, a directed path is followed from the source to a terminal vertex. The probability that a response is in a particular class is the sum of probabilities of paths from the source to the terminal vertices of that class.

When a response is made, an observation might be made of the time required to make the response, or of some other quantity. To account for these, response times and other measures are sometimes incorporated in MPTs (e.g., [12–18]. Here we assume that each arc has associated with it, in addition to a probability, another parameter. We often consider the additional parameter to be the time required for the outcome corresponding to the arc to occur, but the parameter could have another interpretation, such as a cost, so we call it a measure. We assume the measure for a directed path from one vertex to another is the sum of the measures associated with the arcs on the path.

1.1. Selective Influence

An experimental factor, such as the brightness of a stimulus, selectively influences a vertex if changing a level of the factor changes parameter values on some arcs descending from the vertex and on no other arcs. For example, in an MPT model of a list learning, it was found that changing word frequency changed the probability of target recollection, but changed no other parameter [19]. We sometimes say the factor selectively influences the process that the vertex represents. For a review of selective influence in MPTs and more examples see [20].

Suppose an experiment is carried out with two factors, Φ and Ψ . A level of Factor Φ is denoted *i*, with *i* = 1, ..., *I* and a level of Factor Ψ is denoted *j*, with *j* = 1, ..., *J*. We are interested in the situation in which each factor selectively influences a different vertex in an MPT. For example, in Figure 1, the source vertex is selectively influenced by Factor Φ and a vertex following the source vertex is selectively influenced by Factor Ψ . Parameters on arcs descending from the source vertex are indexed by the level of Factor Φ . When Factor Φ is at level *i*, the probability the vertex on the left is the outcome of the processing at the source vertex is $p_A(i)$ and $t_A(i)$ is the measure associated with this outcome (e.g., the time required for this outcome to occur). Other notation is similar.

For the MPT in Figure 1, when Factor Φ is at level *i* and Factor Ψ is at level *j* the probability p(i, j) of a correct response is

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j).$$

The measure t(i, j) associated with a correct response, given a correct response is made, is a mixture of measures of each of the two paths to a terminal vertex for a correct response. That is,

$$t(i, j) = [p_A(i)p_D / p(i, j)][t_A(i) + t_D] + [p_B(i)p_F(j) / p(i, j)][t_B(i) + t_F(j)].$$

More conveniently,

$$p(i, j)t(i, j) = p_A(i)p_D[t_A(i) + t_D] + p_B(i)p_F(j)[t_B(i) + t_F(j)]$$

In equations above and following, probabilities such as $p_A(i)$ and measures such as $t_A(i)$ are fixed real numbers. They can be considered as population means. In previous papers, e.g., [12], we consider the measure associated with each arc in an MPT to be a random variable. We derive distribution free bootstrapping statistical tests that two factors selectively influence different vertices in an arbitrary MPT. Tests are based on observed means, density functions and cumulative distribution functions. Use of the tests is demonstrated with simulations. Similar tests specifically for the MPT in Figure 1 can be used, but are beyond the scope of this paper.

1.2. Vertex Arrangements

Consider an MPT in which each of two factors selectively influences a different vertex. There are only two ways the two vertices can be arranged in the MPT. Suppose there is a directed path from the source vertex to a terminal vertex, and on this path there is an arc whose parameter values depend on the level of one of the factors, and also an arc whose parameter values depend on the level of the other factor. Then we say the vertices are *ordered by the factors*, or for short, *ordered*. If there is no such path, we say the vertices are *unordered by the factors*, or for short, *unordered*. (Note that a path directed from one selectively influenced vertex to the other will not suffice for the vertices to be ordered by the factors, if no arc on the path has a parameter whose value depends on a level of one

of the factors.) In the MPT in Figure 1 the selectively influenced vertices are ordered by the factors.

Two MPTs are *equivalent* for Factors Φ and Ψ , with respective levels i = 1, ..., I and j = 1, ..., J, if the MPTs lead to the same values of p(i, j) and t(i, j) for every i and j. Earlier work shows that if each of two factors selectively influences a different vertex in an arbitrary MPT with two response classes, that MPT is equivalent to one of two relatively simple MPTs [21,22]. If in the arbitrary MPT the selectively influenced vertices are ordered by the factors, the MPT is equivalent to the Standard Tree for Ordered Processes in Figure 2. Otherwise, the arbitrary MPT is equivalent to the Standard Tree for Unordered Processes in Figure 3. The MPTs are similar in that in each has two special vertices, one selectively influenced by each factor. They differ in whether or not the two special vertices are connected by a directed path.



Figure 3. The Standard Tree for Unordered Processes.

In what follows, we assume there are exactly two response classes. For binary trees there is nothing new to consider for selectively influenced vertices that are unordered by the factors, because the Standard Tree for Unordered Processes is itself a binary tree. For vertices ordered by the factors, consider the binary tree in Figure 1, with exactly two arcs descending from each nonterminal vertex. It turns out that if in an arbitrary MPT with two response classes, two factors selectively influence vertices v_1 and v_2 ordered by the factors, with v_1 preceding v_2 , the arbitrary MPT is equivalent for response probabilities to the MPT in Figure 1, provided a simple condition is met. The condition is that the factor selectively influencing vertex v_1 changes parameters on only two descending arcs [23].

An example of a binary Multinomial Processing Tree is in [19]. It is an extension of the MPT usually used in the Process Dissociation model for list learning [5]. In that model, a test item can be recognized as correct via recollection, i.e., a conscious vivid reinstatement of features present at study, or if that fails, via the familiarity of the item. The Dual Recollection Principle [19] is that there are two components of recollection, recollection of the item (Target Recollection) and recollection of the context (Context Recollection). With the Process Dissociation Procedure, items are tested in several situations, and a separate binary MPT is used for each (a forest of trees) [19]. The paper presents experiments on all, which overall support the Dual Recollection Principle.

There are several MPTs, each with its own source, but in any particular situation only one MPT applies. For this example, we confine our attention to one MPT, for the situation in which a test item could be from List 1 or List 2 or neither, and the test item is actually from List 1. The participant's task is to respond accept if the item is from either List 1 or List 2 and reject if not.

In the MPT for this situation, a response is made as follows. The MPT is binary because each process has two possible outcomes. With probability R_{C1} , Context Recollection occurs and the item is accepted. If that fails, with probability R_{T1} , Target Recollection occurs and the item is accepted. If that fails, with probability F_1 , the item provokes a high level of familiarity and the item is accepted. If that fails, with response bias probability b_{12} , the item is accepted and with probability $1 - b_{12}$ the item is rejected.

In Experiment 1, lists consist of words. From various considerations, the authors propose that the word frequency of a test item from List 1 will selectively influence R_{T1} and the concreteness of the item will selectively influence F_1 . Let *i* denote a level of word frequency and *j* denote a level of concreteness. Let p(i, j) denote the probability of acceptance. Then the model predicts

$$p(i, j) = R_{C1} + (1 - R_{C1}) \{ R_{T1}(i) + (1 - R_{T1}(i)) [F_1(j) + (1 - F_1(j))b_{12}] \}$$

Analysis of Experiment 1 indicated that changing word frequency indeed changed values of R_{T1} and of no other parameter. However, concreteness did not selectively influence F_1 . Although changing concreteness changed values of F_1 , as proposed, it also changed values of b_{12} .

Follow up questions come to mind, for which analysis here is helpful. To account for response probabilities, does any binary MPT exist in which word frequency and concreteness selectively influence two different ordered vertices? In a later paper [13], response time was modeled with an MPT using a method of [14]. For a response at a particular terminal vertex, the response time was considered to be the duration of the entire path from the source to that vertex. Is there a way to consider durations of individual arcs on such paths? If so, and a experiment is to be done with word frequency and concreteness as factors, how many levels of each factor would be needed to test the MPT?

The first question can be readily answered. If there is any MPT that can account for response probabilities and in the MPT word frequency and concreteness selectively influence two ordered vertices, the factors would have to do so in the Standard Binary Tree for Ordered Processes of Figure 1 [23]. Any other suitable MPT must "collapse" into this one. For a level *i* of word frequency and a level *j* of concreteness, it predicts

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j).$$

The equation above for the MPT of Brainerd, et al. (2015) can easily be written in this form, with

$$p_A(i) = R_{C1} + (1 - R_{C1})R_{T1}(i)$$

$$p_D = 1$$

$$p_B(i) = (1 - R_{C1})(1 - R_{T1}(i))$$

$$p_F(j) = F_1(j) + (1 - F_1(j) b_{12}).$$

(It is by coincidence that in $p_F(j)$ and in $F_1(j)$ symbol F occurs). Parameter b_{12} was found to depend on the level j of concreteness [19], so it can be rewritten as $b_{12}(j)$. Nonetheless, parameter $p_F(j)$ depends on the level j of concreteness and not on the level i of word frequency. The two factors selectively influence two ordered vertices in the binary MPT in Figure 1. We emphasize that if this did not happen, no binary MPT exists in which the two factors selectively influence two ordered vertices.

2. Results

The MPT in Figure 1 is the Standard Binary Tree for Ordered Processes. This paper presents three results about it. First, the theorem to follow derives conditions this MPT predicts for response probability and an additional measure such as response time. The theorem shows further that if the three conditions hold, then the Standard Binary Tree for Ordered Processes can account for the response probabilities and the additional measure, no matter how they may have been generated. Second, suppose parameter values are found that enable the Standard Binary Tree for Ordered Processes to account for response probabilities and the additional measure. It might seem at first that these are the only parameter values that can do so, and one might propose theoretical interpretations of the values. We show that the parameter values are not unique, and show how different sets of values must be related to each other. Finally, if one wants to test the model with a goodness of fit statistic, the degrees of freedom are needed and we derive them.

By saying every level of a factor selectively influencing a vertex is *effective* we mean the factor has at least two levels and there are no two levels such that when the levels of all the other factors are fixed, in every response class the probability of a response is the same for both levels and the additional measure is the same for both levels. (It is possible that in some response class the probability of a response is the same for both levels, but the measure changes, or vice versa.) We assume there are two response classes; we label them as correct and incorrect, but the classes may be of any kind. Suppose Factor Φ is at level *i* and Factor Ψ is at level *j*. The probability of a correct response is denoted p(i, j). The probability of an incorrect response is 1 - p(i, j). The measure associated with a correct response is denoted t(i, j) and that associated with an incorrect response is denoted $t_w(i, j)$. For convenience, we denote these quantities with matrices. For example, probability matrix **P** has p(i, j) in the cell in row *i* and column *j*.

Definition 1. Probability matrix $\mathbf{P} = (p(i, j))$, correct-response-measure matrix $\mathbf{T} = (t(i, j))$, and incorrect-response-measure matrix $\mathbf{T}_{\mathbf{w}} = (t_w(i, j))$ are produced by two factors selectively influencing two vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , if the following are true. Both factors are effective and for all i and j there are probability parameters $p_A(i)$, $p_B(i)$, p_C , p_D , $p_E(j)$, $p_F(j)$ such that

$$0 \le p_A(i), p_B(i), p_C, p_D, p_E(j), p_F(j) \le 1$$

 $p_A(i) + p_B(i) = 1$
 $p_C + p_D = 1$
 $p_E(j) + p_F(j) = 1$

and

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j)$$

Further, there are measure parameters $t_A(i)$, $t_B(i)$, t_D , $t_F(j)$ such that

$$p(i, j)t(i, j) = p_A(i)p_D[t_A(i) + t_D] + p_B(i)p_F(j) [t_B(i) + t_F(j)].$$

And there are further measure parameters t_C and $t_E(j)$ such that

$$[1 - p(i, j)]t_w(i, j) = p_A(i)[1 - p_D][t_A(i) + t_C] + p_B(i)p_E(j)[t_B(i) + t_E(j)].$$

2.1. Representation

The following theorem gives necessary and sufficient conditions for response probabilities and response measures to be produced by factors selectively influencing vertices in the Standard Binary Tree for Ordered Processes. We note that response probabilities p(i, j) are required to be strictly between 0 and 1 for all levels *i* and *j* of the factors, to avoid dividing by 0 at certain places in the proof. The proof uses a result from [22], which is provided in the Appendix A here.

Theorem 1. Consider probability matrix $\mathbf{P} = (p(i, j))$, with for all *i* and *j*, 0 < p(i, j) < 1; correct-response measure matrix $\mathbf{T} = (t(i, j))$; and incorrect-response measure matrix $\mathbf{T}_{\mathbf{w}} = (t_w(i, j))$. The following two statements are equivalent.

A. Matrices **P**, **T** and **T**_w are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ .

B. There is a level n of Factor Ψ and for every level i of Factor Φ there are numbers $r_i \ge 0$ and s_i , such that the following three conditions are true, with h a value of i such that max $\{r_i\} = r_h$:

1. There is a constant k, $0 \le k \le 1$, *such that for every* $i \ne h$ *and* $j \ne n$

$$p(i, j)p(h, n) - p(i, n)p(h, j) = k[p(i, j) - p(h, j) - p(i, n) + p(h, n)].$$

 $p(i, j) - k = r_i[p(h, j) - k].$

2. For every j,

3. For every j,

$$r_{h} r_{i} s_{i} [p(h, j) - p(h, n)]$$

$$= r_{h} [p(i, j)t(i, j) - p(i, n)t(i, n)] - ri [p(h, j)t(h, j) - p(h, n)t(h, n)]$$

$$= -r_{h}\{[1 - p(i, j)]t_{w}(i, j) - [1 - p(i, n)]t_{w}(i, n)]\}$$

$$+ r_{i}\{[1 - p(h, j)]t_{w}(h, j) - [1 - p(h, n)]t_{w}(h, n)]\}.$$
(1)

Proof. I. Suppose probability matrix $\mathbf{P} = (p(i, j)), 0 < \mathbf{P} < 1$, and measure matrix $\mathbf{T} = (t(i, j))$ are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Ψ .

Then for any *i* and *j*,

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j)$$
 (2)

and

$$p(i, j)t(i, j) = p_A(i)p_D[t_A(i) + t_D] + p_B(i)p_F(j)[t_B(i) + t_F(j)]$$

with

$$p_A(i) = 1 - p_B(i).$$

Let *h* be a value of *i* such that $\max\{p_B(i)\} = p_B(h)$. For every *i*, let $r_i = \frac{p_B(i)}{p_B(h)}$. Note that $r_h = \max\{r_i\}$. Clearly, $0 \le r_i$.

Let *n* be a value of *j* such that $min\{p(h, j)t(h, j)\} = p(h, n)t(h, n)$.

With a little algebra it follows from Equation (2) that for every $i \neq h$ and $j \neq n$

 $p(i, j)p(h, n) - p(i, n)p(h, j) = p_D[p(i, j) - p(h, j) - p(i, n) + p(h, j)].$

Hence, Condition 1 is true, with $k = p_D$.

Note that if $p_B(h) = 0$, then for every i, $p_B(i) = 0$ so Factor Φ is ineffective. Hence $p_B(h)$ does not equal 0.

From Equation (2), for every *i* and for every *j*,

$$p(i, j) - k = k + p_B(i)[p_F(j) - k] - k$$

= [p_B(i) / p_B(h)]p_B(h)[p_F(j) - k]
= r_i[p(h, j) - k].

Hence, Condition 2 is true.

We turn to Condition 3.

If $r_i = 0$, let $s_i = 0$.

For *i* such that $r_i \neq 0$, let $s_i = t_B(i) - t_B(h)$. Clearly, s_i does not depend on *j*. We now consider the right of Equation (1) and show that it equals $r_h r_i s_i [p(h, j) - p(h, n)]$. Suppose $r_i = 0$. Then $s_i = 0$, so $r_h r_i s_i [p(h, j) - p(h, n)] = 0$. Because $r_i = p_B(i)/p_B(h) = 0$,

it follows that $p_B(i) = 0$ and $p_A(i) = 1$. The right hand side of Equation (1) is also 0, that is,

$$r_h [p(i, j)t(i, j) - p(i, n)t(i, n)] = r_h [p_A(i)p_D - p_A(i)p_D] = 0.$$

Then Condition 3 is true if $r_i = 0$.

Suppose $r_i \neq 0$.

Select *j* such $p(h, j) \neq p(h, n)$. Because neither r_h , r_i nor p(h, j) - p(h, n) is 0, we can divide expressions by $r_h r_i [p(h, j) - p(h, n)]$.

On division by $r_h r_i [p(h, j) - p(h, n)]$, the left hand side of Equation (1) is s_i . On division by $r_h r_i [p(h, j) - p(h, n)]$, the right hand side of Equation (1) is

$$\frac{[p(i, j)t(i, j) - p(i, n)t(i, n)]/r_i - [p(h, j)t(h, j) - p(h, n)t(h, n)]/r_h}{p(h, j) - p(h, n)}$$
(3)

The first term in the numerator of Equation (3) is

$$[p(i, j)t(i, j) - p(i, n)t(i, n)]/r_i = \{p_B(i)p_F(j)[t_B(i)+t_F(j)] - p_B(i)p_F(n)[t_B(i)+t_F(n)]\}/r_i = p_B(h)\{p_F(j)[t_B(i)+t_F(j)] - p_F(n)[t_B(i)+t_F(n)]\}$$

Similarly, the second term in the numerator of Equation (3) is

$$[p(h, j)t(h, j) - p(h, n)t(h, n)]/r_h = p_B(h)\{p_F(j)[t_B(h) + t_F(j)] - p_F(n)[t_B(h) + t_F(n)]\}$$

The numerator of Equation (3) becomes

$$p_B(h)\{p_F(j)[t_B(i) - t_B(h)] - p_F(n)[t_B(i) - t_B(h)]\} = p_B(h)[p_F(j) - p_F(n)][t_B(i) - t_B(h)]$$

The denominator of Equation (3) is $p(h, j) - p(h, n) = p_B(h)[p_F(j) - p_F(n)]$. Then the ratio in Equation (1) becomes

$$t_B(i) - t_B(h) = s_i.$$

Hence, Condition 3 is true.

II. Suppose Conditions 1, 2 and 3 are true. We begin by showing that matrices **P** and **T** are produced by the Standard Tree for Ordered Processes, and then we show they are also produced by the Standard Binary Tree for Ordered Processes.

In [23] it is shown that from Conditions 1 and 2, parameters $p_A(i)$, $p_B(i)$, p_C , p_D , $p_E(j)$, and $p_F(j)$ exist, such that

$$0 \le p_A(i), p_B(i), p_C, p_D, p_E(j), p_F(j) \le 1$$
$$p_A(i) + p_B(i) = 1$$
$$p_C + p_D = 1$$
$$p_E(j) + p_F(j) = 1$$

and

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j),$$
 (4)

with $p_D = k$.

We now show that **P** and **T** are produced by Factor Φ and Factor Ψ selectively influencing two vertices in the Standard Tree for Ordered Processes. Necessary and sufficient conditions are in [22] (see Appendix A).

Renumber the levels of Factor Ψ so $p_F(1) \leq \ldots \leq p_F(j) \leq \ldots \leq p_F(J)$. Then by Equation (4), for every i, $p(i, 1) \leq \ldots \leq p(i, j) \leq \ldots \leq p(i, J)$. Hence, Condition 1 of Theorem 5 of [22] is satisfied.

By Condition 2 of the theorem to be proved, for every level *i* of Factor Φ there exists a number $r_i \ge 0$ such that for every level *j*,

$$p(i, j) - k = r_i[p(h, j) - k],$$

where *h* is a level such that $\max\{r_i\} = r_h$. Let *n* be a level of Factor Ψ such that $\min\{p(h, j)t(h, j)\} = p(h, n)t(h, n)$. Then for every *i* and *j*,

$$p(i, j) - p(i, n) = p(i, j) - k - [p(i, n) - k]$$

= $r_i[p(h, j) - k] - r_i[p(h, n) - k]$
= $r_i[p(h, j) - p(h, n)].$

Hence Condition 2 of Theorem 5 is satisfied (with $i^* = h$ and $j^* = n$).

Condition 3 of Theorem 5 is satisfied because it is the same as Condition 3 of the theorem to be proved, and we have assumed Condition 3 to be true.

Because Conditions 1, 2 and 3 of Theorem 5 are satisfied, there exist probability parameters (denoted φ) and measure parameters (denoted μ) of the Standard Tree for Ordered Processes such that the following equations are true for every level *i* of Factor Φ and every level *j* of Factor Ψ , see Figure 2. The arc denoted *D* in Figure 2 (for the Standard Tree for Ordered Processes) is denoted Δ here to distinguish it from arc *D* in the Standard Binary Tree for Ordered Processes (Figure 1).

$$p(i,j) = \varphi_{\Delta}(i) + \varphi_B(i)\varphi_F(j).$$
(5)

$$p(i,j)t(i,j) = \varphi_{\Delta}(i)\mu_{\Delta}(i) + \varphi_{B}(i)\varphi_{F}(j)[\mu_{B}(i) + \mu_{F}(j)].$$
(6)

From the proof of Theorem 5, parameters can be assigned so $\mu_F(n) = 0$.

Equation (4) has the form of an equation accounting for p(i, j) with the Standard Tree for Ordered Processes. Then Equations (4) and (5) provide two sets of probability parameters that account for p(i, j) with the Standard Tree for Ordered Processes. Then by Theorem 6 of Schweickert and Zheng (2019b), with scaling parameters *c* and *d* we can transform probability parameters in Equation (5) to those in Equation (4), as follows.

For every level *i*, $p_B(i) = \varphi^*_B(i) = c\varphi_B(i)$.

For every level *i*, $p_A(i)p_D = \varphi^*_{\Delta}(i) = \varphi_{\Delta}(i) - cd\varphi_B(i)$.

For every level j, $p_F(j) = \varphi^*_F(j) = \varphi_F(j)/c + d$.

Now by Theorem 6 of Schweickert and Zheng (2019b), the following transformations of the measure parameters in Equation (6) are admissible transformations, with scaling parameters c and d, and with j' = n and e = 0.

For every level *i*, let $\mu^*_B(i) = \mu_B(i)$. Recall that $\mu_F(n) = 0$. Let $\mu^*_F(n) = \mu_F(n) = 0$. For every level *i*, let

$$\mu_{\Delta}^{*}(i) = \frac{-\varphi_{\Delta}(i)\mu_{\Delta}(i) + cd\varphi_{B}(i)\mu_{B}(i)}{cd\varphi_{B}(i) - \varphi_{\Delta}(i)}$$

For every level *j*, let

$$\mu_F^*(j) = \frac{\varphi_F(j)\mu_F(j)}{cd + \varphi_F(i)}$$

Then by Theorem 6, with the transformed parameters, for every level *i* and every level *j*,

$$p(i, j) = \varphi^*_{\Delta}(i) + \varphi^*_B(i) \varphi^*_F(j)$$

= $p_A(i)p_D + p_B(i)p_F(j)$

and

$$p(i, j)t(i, j) = \varphi^*_{\Delta}(i)\mu^*_{\Delta}(i) + \varphi^*_{B}(i)\varphi^*_{F}(j)[\mu^*_{B}(i) + \mu^*_{F}(j)]$$

= $p_{A}(i)p_{D}[\mu^*_{\Delta}(i) + 0] + p_{B}(i)p_{F}(j)[\mu^*_{B}(i) + \mu^*_{F}(j)].$

Then **P** and **T** are produced by the Standard Binary Tree for Ordered Processes, with $t_A(i) = \mu^*_{\Delta}(i)$, $t_D = 0$, $t_B(i) = \mu^*_B(i)$ and $t_F(j) = \mu^*_F(j)$.

The subtree of the Standard Binary Tree for Ordered Processes that produces correct responses has the same form as the subtree that produced incorrect responses. So, reasoning for **P** and **T**_w is similar to that above. \Box

2.2. Uniqueness of Parameters

Suppose the Standard Binary Tree for Ordered Processes accounts for observed response probabilities and response measures with a particular set of parameter values. Those parameter values are not necessarily the only ones that can account for the data.

2.2.1. Numerical Example

Table 1 gives two different sets of values, old and new, for parameters of the Standard Binary Tree for Ordered Processes that make the same predictions. The parameter values are for a particular level *i* of Factor Φ and a particular level *j* of Factor Ψ . When Factors Φ and Ψ have levels *i* and *j*, respectively, the old parameter values predict for the probability of a correct response

$$p(i, j) = p_A(i)p_D + p_B(i)p_F(j) = 0.50 \times 0.40 + 0.50 \times 0.16 = 0.28.$$

The new parameter values predict the same

$$p(i, j) = p_A^*(i)p_D^* + p_B^*(i)p_F^*(j) = 0.20 \times 0.40 + 0.80 \times 0.25 = 0.28.$$

Likewise, the old parameter values predict for the product of correct response probability and response measure

$$p(i, j)t(i, j) = p_A(i)p_D[t_A(i) + t_D] + p_B(i)p_F(j)[t_B(i) + t_F(j)] = 2.26.$$

And the new parameter values predict the same

$$p(i, j)t(i, j) = p^*{}_A(i)p^*{}_D[t^*{}_A(i) + t^*{}_D] + p^*{}_B(i)p^*{}_F(j)[t^*{}_B(i) + t^*{}_F(j)] = 2.26.$$

Old and new parameter values are related through the admissible transformations in Table 2. There are four scaling parameters, *c*, *e*, *f*, and $t^*_F(j')$, where j' is an arbitrary level of Factor Ψ , chosen so $p^*_F(j') \neq 0$. To insure that new probability parameters are between 0 and 1, the scaling parameters must satisfy the bounds in Table 3, which were derived by Schweickert and Chen (2008).

 Table 1.
 Numerical Example of Transformed Parameters in The Standard Binary Tree for

 Ordered Processes.
 Processes.

	Old	New ¹
$p_A(i)$	0.50	0.20
$p_B(i)$	0.50	0.80
p_D	0.40	0.40
$p_F(j)$	0.16	0.25
$t_A(i)$	4.50	7.50
t_D	4.00	7.00
$\overline{t_B(i)}$	2.00	3.00
$t_F(j)$	5.00	2.50

¹ New parameter values were obtained from old ones, including $p_F(j') = 0.20$ and $t_F(j') = 8$, through the admissible transformations in Table 2, using scaling parameter values c = 1, e = 3, f = 1 and $t^*_F(j') = 4$.

Table 2. Admissible Transformations of Parameters.

Standard Binary Tree for Ordered Processes ¹		
$p^*_B(i) = cp_B(i)$		
$p^*_{D=} p_D$		
$p_{F}^{*}(j) = p_{F}(j)/c + (c-1) p_{D}/c$		
If $p^*_B(i) \neq 0$, $t^*_B(i) = t_B(i) + f$		
If $p^*_F(j) \neq 0$,		
$t_F^*(j) = \frac{p_F(j)t_F(j) + f[p_F(j') - p_F(j)] + t_F^*(j')[p_F(j') + (c-1)p_D] - p_F(j')t_F(j')}{p_F(j) - (c-1)p_D}$		
If $p_{T}^* \neq 0$, $t_{T}^* = t_D + e$		
If $p^*_A(i) \neq 0$,		
$t_{A}^{*}(i) = \frac{p_{A}(i)t_{A}(i) + p_{B}(i)\{[t_{B}(i) - t_{D}](1 - c) + [f + t_{F}^{*}(j')][1 - c - \frac{p_{F}(j')}{p_{D}}] + \frac{p_{F}(j')t_{F}(j')}{p_{D}} + ce\} - e}{1 - cn_{B}(i)}$		
If $p^*_{C} \neq 0$, $t^*_{C} = t_{C} + k$		
If $p^*_F(I) \neq 0$.		
$t_{E}^{*}(j) = \frac{p_{E}(j)t_{E}(j) + f[p_{E}(j') - p_{E}(j)] + [c - p_{F}(j') - (c - 1)p_{D}]t_{E}^{*}(j') - p_{E}(j')t_{E}(j')}{c - p_{F}(j) - (c - 1)p_{D}}$		

¹ Level j' of Factor Ψ is chosen so $0 \neq p^*_F(j') \neq 1$.

Table 3. Bounds on Scaling Parameters in Admissible Transformations For the Standard Binary Tree for Ordered Processes For probability parameters to be between 0 and 1.

$0 < c \le 1/\max\{p_B(i)\}$	
$p_D - \min\{p_F(j)\} \le p_D^* c$	
$\max\{p_F(j)\} - p_D \le (1 - p_D)^* c$	

2.2.2. Admissible Transformations

Suppose response probability and response measure are accounted for by two factors selectively influencing two vertices in the Standard Binary Tree for Ordered Processes. If two sets of parameter values are possible, the following theorem gives the relations between them.

Theorem 2. Suppose probability matrix **P** and measure matrix **T** are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , with probability parameters $p_A(i)$, $p_B(i)$, p_D , and $p_F(j)$, and measure parameters $t_A(i)$, $t_B(i)$, t_D , and $t_F(j)$.

Then **P** and **T** are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , with probability parameters $p^*_A(i)$, p^*_D , $p^*_B(i)$ and $p^*_F(j)$, and measure parameters $t^*_A(i)$, t^*_D , $t^*_B(i)$ and $t^*_F(j)$ if and only if there are constants *c*, *e* and *f* and a level *j'* of Factor Ψ such that the admissible transformations in Table 2 apply and the bounds in Table 3 are satisfied.

Proof. Suppose probability matrix $\mathbf{P} = (p(i, j))$ and measure matrix $\mathbf{T} = (t(i, j))$ are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , with probability parameters $p_A(i)$, p_D , $p_B(i)$ and $p_F(j)$, and measure parameters $t_A(i)$, t_D , $t_B(i)$ and $t_F(j)$.

I. Suppose **P** and **T** = (t(i, j)) are also produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , with probability parameters $p^*_A(i)$, p^*_D , $p^*_B(i)$ and $p^*_F(j)$, and measure parameters $t^*_A(i)$, t^*_D , $t^*_B(i)$ and $t^*_F(j)$.

The admissible transformations in Table 2, with bounds in Table 3, for the probability parameters $p_A^*(i)$, p_D^* , $p_B^*(i)$ and $p_F^*(j)$ were shown to apply in [23].

We turn to admissible transformations for the measure parameters.

Suppose $p_{A}^{*}(i)$, p_{D}^{*} , $p_{B}^{*}(i)$, $p_{F}^{*}(j)$, $t_{A}^{*}(i)$, t_{D}^{*} , $t_{B}^{*}(i)$, and $t_{F}^{*}(j)$ exist with

$$p(i, j)t(i, j) = p_A^*(i)p_D^*[t_A^*(i) + t_D^*] + p_B^*(i)p_F^*(j)[t_B^*(i) + t_F^*(j)]$$

Also, for any *i* and *j*,

$$p(i, j)t(i, j) = p_A(i)p_D [t_A(i) + t_D] + p_B(i)p_F(j) [t_B(i) + t_F(j)].$$

Let *j* and *j'* be two different values of *j*. Then

$$p(i, j)t(i, j) - p(i, j')t(i, j') = p_A(i)p_D [t_A(i) + t_D] + p_B(i)p_F(j) [t_B(i) + t_F(j)] - p_A(i)p_D [t_A(i) + t_D] - p_B(i)p_F(j') [t_B(i) + t_F(j')] = p_B(i)p_F(j) [t_B(i) + t_F(j)] - p_B(i)p_F(j') [t_B(i) + t_F(j')].$$
(7)

Additionally, for the same *i* and *j*, due to the assumption of

$$p(i, j)t(i, j) = p_{A}^{*}(i)p_{D}^{*}[t_{A}^{*}(i) + t_{D}^{*}] + p_{B}^{*}(i)p_{F}^{*}(j)[t_{B}^{*}(i) + t_{F}^{*}(j)],$$

we have for any *i*, *j* and j'

$$p(i, j)t(i, j) - p(i, j')t(i, j') = p^*{}_B(i)p^*{}_F(j) [t^*{}_B(i) + t^*{}_F(j)] - p^*{}_B(i)p^*{}_F(j') [t^*{}_B(i) + t^*{}_F(j')].$$
(8)

Because Equations (7) and (8) have the same left side, the right sides are the same, that is,

 $p_B(i)p_F(j) [t_B(i) + t_F(j)] - p_B(i)p_F(j') [t_B(i) + t_F(j')]$ = $p^*_B(i)p^*_F(j) [t^*_B(i) + t^*_F(j)] - p^*_B(i)p^*_F(j') [t^*_B(i) + t^*_F(j')].$

Substitute

$$p^{*}{}_{B}(i) = c p_{B} (i)$$

$$p^{*}{}_{F}(j) = p_{F}(j)/c + (c - 1) p_{D}/c$$

We have

$$p_B(i)p_F(j) [t_B(i) + t_F(j)] - p_B(i)p_F(j') [t_B(i) + t_F(j')] = cp_B(i)[p_F(j)/c + (c - 1) p_D/c][t_B^*(i) + t_F^*(j)] - cp_B(i)[p_F(j')/c + (c - 1) p_D/c][t_B^*(i) + t_F^*(j')]$$

Equivalently,

$$\begin{aligned} p_F(j)t_B(i) + p_F(j)t_F(j) &- p_F(j')t_B(i) - p_F(j')t_F(j') \\ &= p_F(j)t^*_B(i) + p_F(j)t^*_F(j) + (c-1)p_D[t^*_F(j) - t^*_F(j')] \\ &- p_F(j')t^*_B(i) - p_F(j')t^*_F(j') \end{aligned}$$

Then we get

$$\begin{split} & [t^*{}_B(i) - t_B(i)][p_F(j') - p_F(j)] \\ &= p_F(j)t^*{}_F(j) + (c-1)p_D[t^*{}_F(j) - t^*{}_F(j')] - p_F(j')t^*{}_F(j') \\ &\quad - p_F(j)t_F(j) + p_F(j')t_F(j') \end{split}$$

As a result, for any *j* such that $p_F(j) \neq p_F(j')$

$$t_B^*(i) - t_B(i) = \frac{p_F(j)t_F^*(j) + (c-1)p_D[t_F^*(j) - t_F^*(j')] - p_F(j')t_F^*(j') - p_F(j)t_F(j) + p_F(j')t_F(j')}{p_F(j') - p_F(j)}$$

The left hand side of the above equation cannot change when *j* changes. So the left hand side must be a constant, denote it as *f*.

Hence, $t^*_B(i) = t_B(i) + f$.

Further, the right hand side equals *f*.

$$\frac{p_F(j)t_F^*(j) + (c-1)p_D[t_F^*(j) - t_F^*(j')] - p_F(j')t_F^*(j') - p_F(j)t_F(j) + p_F(j')t_F(j')}{p_F(j') - p_F(j)} = f$$

Hence, there is

$$t_F^*(j) = \frac{p_F(j)t_F(j) + f[p_F(j') - p_F(j)] + t_F^*(j')[p_F(j') + (c-1)p_D] - p_F(j')t_F(j')}{p_F(j) + (c-1)p_D}$$

Similarly, according to the assumptions,

$$p(i, j)t(i, j) = p_A(i)p_D [t_A(i) + t_D] + p_B(i)p_F(j) [t_B(i) + t_F(j)]$$

$$p(i, j)t(i, j) = p_A^*(i)p_D^* [t_A^*(i) + t_D^*] + p_B^*(i)p_F^*(j) [t_B^*(i) + t_F^*(j)].$$

Also, for any *i* such that $p_B(i) \neq 0$ and $p^*_B(i) \neq 0$, we have

$$\frac{p(i, j)t(i, j)}{p_B(i)} = \frac{p_A(i)p_D[t_A(i) + t_D]}{p_B(i)} + p_F(j)[t_B(i) + t_F(j)]$$
(9)

$$\frac{p(i, j)t(i, j)}{p_B^*(i)} = \frac{p_A^*(i)p_D^*\left[t_A^*(i) + t_D^*\right]}{p_B^*(i)} + p_F^*(j)[t_B^*(i) + t_F^*(j)]$$
(10)

Because $p^*_B(i) = cp_B(i)$, we have

$$\frac{p(i, j)t(i, j)}{p_B^*(i)} = \frac{p(i, j)t(i, j)}{cp_B(i)}$$

Also, there is $p^*_D = p_D$ from Table 2. So Equation (10) can be transformed into

$$\frac{p(i, j)t(i, j)}{p_B(i)} = c * \frac{p_A^*(i)p_D^*[t_A^*(i) + t_D^*]}{p_B^*(i)} + cp_F^*(j)[t_B^*(i) + t_F^*(j)]$$
(11)

Then according to Equation (9), for any *i* and *i'* which can be any different value of *i* with $p_B(i') \neq 0$,

$$\frac{p(i,j)t(i,j)}{p_{B}(i)} - \frac{p(i',j)t(i',j)}{p_{B}(i')} = \frac{p_{A}(i)p_{D}[t_{A}(i)+t_{D}]}{p_{B}(i)} + p_{F}(j)[t_{B}(i) + t_{F}(j)] - \frac{p_{A}(i')p_{D}[t_{A}(i')+t_{D}]}{p_{B}(i')} - p_{F}(j)[t_{B}(i') + t_{F}(j)] = \frac{p_{A}(i)p_{D}t_{A}(i)}{p_{B}(i)} + \frac{p_{A}(i)p_{D}t_{D}}{p_{B}(i)} + p_{F}(j)[t_{B}(i) - t_{B}(i')] - \frac{p_{A}(i')p_{D}t_{A}(i')}{p_{B}(i')} - \frac{p_{A}(i')p_{D}t_{D}}{p_{B}(i')} = \frac{p_{A}(i)p_{D}t_{A}(i)}{p_{B}(i)} - \frac{p_{A}(i')p_{D}t_{A}(i')}{p_{B}(i')} + p_{F}(j)[t_{B}(i) - t_{B}(i')] + \frac{p_{A}(i)p_{D}t_{D}}{p_{B}(i')} - \frac{p_{A}(i')p_{D}t_{D}}{p_{B}(i')} = \frac{p_{A}(i)p_{D}t_{A}(i)}{p_{B}(i)} - \frac{p_{A}(i')p_{D}t_{A}(i')}{p_{B}(i')} + p_{F}(j)[t_{B}(i) - t_{B}(i')] + \left[\frac{1-p_{B}(i)}{p_{B}(i)} - \frac{1-p_{B}(i')}{p_{B}(i')}\right]p_{D}t_{D} = \frac{p_{A}(i)p_{D}t_{A}(i)}{p_{B}(i)} - \frac{p_{A}(i')p_{D}t_{A}(i')}{p_{B}(i')} + p_{F}(j)[t_{B}(i) - t_{B}(i')] + \left[\frac{1}{p_{B}(i)} - \frac{1}{p_{B}(i')}\right]p_{D}t_{D}$$
(12)

Also, according to Equation (11), we have

$$\begin{aligned} &\frac{p(i,j)t(i,j)}{p_{B}(i)} - \frac{p(i',j)t(i',j)}{p_{B}(i')} \\ &= c\frac{p_{A}^{*}(i)p_{D}[t_{A}^{*}(i)+t_{D}^{*}]}{p_{B}^{*}(i)} + cp_{F}^{*}(j)[t_{B}^{*}(i) + t_{F}^{*}(j)] - c\frac{p_{A}^{*}(i')p_{D}[t_{A}^{*}(i)+t_{D}^{*}]}{p_{B}^{*}(i')} - cp_{F}^{*}(j)[t_{B}^{*}(i') + t_{F}^{*}(j)] \\ &= c\{\frac{p_{A}^{*}(i)p_{D}[t_{A}^{*}(i)+t_{D}^{*}]}{p_{B}^{*}(i)} + p_{F}^{*}(j)[t_{B}^{*}(i) + t_{F}^{*}(j)] - \frac{p_{A}^{*}(i')p_{D}[t_{A}^{*}(i)+t_{D}^{*}]}{p_{B}^{*}(i)} - p_{F}^{*}(j)[t_{B}^{*}(i) + t_{F}^{*}(j)]\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} + \frac{p_{A}^{*}(i)p_{D}t_{D}^{*}}{p_{B}^{*}(i)} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{D}^{*}}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{D}^{*}}{p_{B}^{*}(i)} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + \frac{p_{A}^{*}(i)p_{D}t_{D}}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{D}}{p_{B}^{*}(i')}\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i)} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + [\frac{1-p_{B}^{*}(i)}{p_{B}^{*}(i)} - \frac{1-p_{B}^{*}(i')}{p_{B}^{*}(i')}]p_{D}t_{D}^{*}\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i')} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + [\frac{1-p_{B}^{*}(i)}{p_{B}^{*}(i)} - \frac{1-p_{B}^{*}(i')}{p_{B}^{*}(i')}]p_{D}t_{D}^{*}\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i')} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + [\frac{1-p_{B}^{*}(i)}{p_{B}^{*}(i')}]p_{D}t_{D}^{*}\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i')} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + [\frac{1-p_{B}^{*}(i)}{p_{B}^{*}(i')}]p_{D}t_{D}^{*}\} \\ &= c\{\frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}^{*}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}^{*}(i')} + p_{F}^{*}(j)[t_{B}^{*}(i) - t_{B}^{*}(i')] + [\frac{1-p_{B}^{*}(i)}{p_{B}^{*}(i')}]p_{D}t_{D}^{*}\}$$

Substitute the following values into Equation (13),

$$p_{B}^{*}(i) = cp_{B}(i)$$

$$p_{F}^{*}(j) = p_{F}(j)/c + (c - 1) p_{D}/c$$

$$t_{B}^{*}(i) = t_{B}(i) + f.$$

Then

$$\frac{p(i,j)t(i,j)}{p_{B}(i)} - \frac{p(i',j)t(i',j)}{p_{B}(i')} = c\{\frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{cp_{B}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{cp_{B}(i')} + \left[\frac{p_{F}(j)}{c} + \frac{(c-1)p_{D}}{c}\right][t_{B}(i) - t_{B}(i')] + \left[\frac{1}{cp_{B}(i)} - \frac{1}{cp_{B}(i')}\right]p_{D}t_{D}^{*}\}$$

$$= \frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}(i)} - \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}(i')} + [p_{F}(j) + (c-1)p_{D}][t_{B}(i) - t_{B}(i')] + \left[\frac{1}{p_{B}(i)} - \frac{1}{p_{B}(i')}\right]p_{D}t_{D}^{*}$$
(14)

Because Equations (12) and (14) have the same left side, the right sides are the same, that is,

$$\frac{p_A(i)p_Dt_A(i)}{p_B(i)} - \frac{p_A(i')p_Dt_A(i')}{p_B(i')} + p_F(j)[t_B(i) - t_B(i')] + \left[\frac{1}{p_B(i)} - \frac{1}{p_B(i')}\right]p_Dt_D = \frac{p_A^*(i)p_Dt_A^*(i)}{p_B(i)} - \frac{p_A^*(i')p_Dt_A^*(i')}{p_B(i')} + [p_F(j) + (c-1)p_D][t_B(i) - t_B(i')] + \left[\frac{1}{p_B(i)} - \frac{1}{p_B(i')}\right]p_Dt_D^*$$

Equivalently,

$$\frac{p_{A}(i)p_{D}t_{A}(i)}{p_{B}(i)} - \frac{p_{A}(i')p_{D}t_{A}(i')}{p_{B}(i')} + p_{F}(j)[t_{B}(i) - t_{B}(i')] - \frac{p_{A}^{*}(i)p_{D}t_{A}^{*}(i)}{p_{B}(i)} + \frac{p_{A}^{*}(i')p_{D}t_{A}^{*}(i')}{p_{B}(i')} - [p_{F}(j) + (c-1)p_{D}][t_{B}(i) - t_{B}(i')] = [\frac{1}{p_{B}(i)} - \frac{1}{p_{B}(i')}]p_{D}t_{D}^{*} - [\frac{1}{p_{B}(i)} - \frac{1}{p_{B}(i')}]p_{D}t_{D}$$

The left side can be simplified into

$$\frac{p_D}{p_B(i)} [p_A(i)t_A(i) - p_A^*(i)t_A^*(i)] - \frac{p_D}{p_B(i')} \{p_A(i')t_A(i') - p_A^*(i')t_A^*(i')\} + (1-c)p_D][t_B(i) - t_B(i')]$$

Meanwhile, the right side can be simplified into

$$[\frac{1}{p_B(i)} - \frac{1}{p_B(i')}]p_D[t_D^* - t_D]$$

Then we have,

$$\frac{p_D}{p_B(i)} \left[p_A(i)t_A(i) - p_A^*(i)t_A^*(i) \right] - \frac{p_D}{p_B(i')} \left\{ p_A(i')t_A(i') - p_A^*(i')t_A^*(i') \right\} + (1-c)p_D \left[t_B(i) - t_B(i') \right] \\ = \left[\frac{1}{p_B(i)} - \frac{1}{p_B(i')} \right] p_D \left[t_D^* - t_D \right]$$

So for any *i* such that $p_B(i) \neq p_B(i')$, there is,

$$t_{D}^{*} - t_{D} = \frac{\frac{p_{D}}{p_{B}(i)} \left[p_{A}(i)t_{A}(i) - p_{A}^{*}(i)t_{A}^{*}(i) \right] - \frac{p_{D}}{p_{B}(i')} \left\{ p_{A}(i')t_{A}(i') - p_{A}^{*}(i')t_{A}^{*}(i') \right\} + (1 - c)p_{D} \left[t_{B}(i) - t_{B}(i') \right]}{\left[\frac{1}{p_{B}(i)} - \frac{1}{p_{B}(i')} \right] p_{D}}$$

The left hand side of the above equation cannot change when *i* changes. So the left hand side must be a constant, denote it as *e*.

Hence, $t^*_D = t_D + e$.

Because

$$\begin{split} p(i, j)t(i, j) &= p_A(i)p_D \left[t_A(i) + t_D \right] + p_B(i)p_F(j) \left[t_B(i) + t_F(j) \right] \\ p(i, j)t(i, j) &= p_A^*(i)p_D^* \left[t_A^*(i) + t_D^* \right] + p_B^*(i)p_F^*(j) \left[t_B^*(i) + t_F^*(j) \right] \end{split}$$

There is

$$p_A(i)p_D [t_A(i) + t_D] + p_B(i)p_F(j) [t_B(i) + t_F(j)] = p_A^*(i)p_D^* [t_A^*(i) + t_D^*] + p_B^*(i)p_F^*(j) [t_B^*(i) + t_F^*(j)]$$

Substitute

$$p_{B}^{*}(i) = cp_{B}(i)$$

$$p_{F}^{*}(j) = p_{F}(j)/c + (c - 1) p_{D}/c$$

$$t_{B}^{*}(i) = t_{B}(i) + f$$

$$t_{F}^{*}(j) = \frac{p_{F}(j)t_{F}(j) + f * [p_{F}(j') - p_{F}(j)] + t_{F}^{*}(j')[p_{F}(j') + (c - 1)p_{D}] - p_{F}(j')t_{F}(j')}{p_{F}(j) - (c - 1)p_{D}}$$

and $t^*_D = t_D + e$. We get

$$p_{A}(i)p_{D}[t_{A}(i) + t_{D}] + p_{B}(i)p_{F}(j)[t_{B}(i) + t_{F}(j)] \\= [1 - cp_{B}(i)]p_{D}[t_{A}^{*}(i) + t_{D} + e] + cp_{B}(i)\left[\frac{p_{F}(j)}{c} + \frac{(c-1)p_{D}}{c}\right] * \{t_{B}(i) + f(j) + \frac{p_{F}(j)t_{F}(j) + f[p_{F}(j') - p_{F}(j)] + t_{F}^{*}(j')[p_{F}(j') + (c-1)p_{D}] - p_{F}(j')t_{F}(j')}{p_{F}(j) + (c-1)p_{D}}\}$$

Equivalently,

$$\begin{aligned} & [1 - cp_B(i)]p_D[t_A^*(i) + t_D + e] - [1 - p_B(i)]p_D[t_A(i) + t_D] \\ &= p_B(i)p_F(j)[t_B(i) + t_F(j)] - cp_B(i) \Big[\frac{p_F(j)}{c} + \frac{(c-1)p_D}{c} \Big] * \{t_B(i) + f(j) + \frac{p_F(j)t_F(j) + f[p_F(j') - p_F(j)] + t_F^*(j')[p_F(j') + (c-1)p_D] - p_F(j')t_F(j')}{p_F(j) + (c-1)p_D} \Big\} \end{aligned}$$

The left side can be simplified into

$$p_{D}\{[1-cp_{B}(i)]t_{A}^{*}(i)-p_{A}(i)t_{A}(i)+(1-c)p_{B}(i)t_{D}-cep_{B}(i)+e\}$$

The right side can be simplified into

$$p_B(i)\{t_B(i)(1-c)p_D + [f + t_F^*(j')][(1-c)p_D - p_F(j')] + p_F(j')t_F(j')\}$$

Hence,

$$p_D\{[1-cp_B(i)]t_A^*(i) - p_A(i)t_A(i) + (1-c)p_B(i)t_D - cep_B(i) + e\}$$

= $p_B(i)\{t_B(i)(1-c)p_D + [f + t_F^*(j')][(1-c)p_D - p_F(j')] + p_F(j')t_F(j')\}$

Then we have,

$$\begin{split} [1-cp_B(i)]t_A^*(i) \\ &= p_B(i)\{t_B(i)(1-c) + [f+t_F^*(j')]\Big[(1-c) - \frac{p_F(j')}{p_D}\Big] + \frac{p_F(j')t_F(j')}{p_D}\} \\ &+ p_A(i)t_A(i) - (1-c)p_B(i)t_D + cep_B(i) - e\} \\ &= p_A(i)t_A(i) + p_B(i)\{[t_B(i) - t_D](1-c) + [f+t_F^*(j')]\Big[(1-c) - \frac{p_F(j')}{p_D}\Big] + \frac{p_F(j')t_F(j')}{p_D} + ce\} - e \\ &\text{So we get,} \end{split}$$

$$t_{A}^{*}(i) = \frac{p_{A}(i)t_{A}(i) + p_{B}(i)\{[t_{B}(i) - t_{D}](1 - c) + [f + t_{F}^{*}(j')]\left[(1 - c) - \frac{p_{F}(j')}{p_{D}}\right] + \frac{p_{F}(j')t_{F}(j')}{p_{D}} + ce\} - e}{1 - cp_{B}(i)}$$

II. Conversely, suppose for all $1 \le i \le I$, and $1 \le j \le J$, there exist $p^*_A(i)$, p^*_D , $p^*_B(i)$, $p^*_F(j)$, $t^*_A(i)$, t^*_D , $t^*_B(i)$, and $t^*_F(j)$ such that

$$0 \le p_A^*(i), p_D^*, p_B^*(i), p_F^*(j) \le 1$$
$$t_A^*(i), t_D^*, t_B^*(i), t_F^*(j) \ge 0$$

with the following equations,

$$p_{B}^{*}(i) = cp_{B}(i),$$

$$p_{D}^{*} = p_{D},$$

$$p_{F}^{*}(j) = p_{F}(j)/c + (c - 1) p_{D}/c,$$

$$t_{B}^{*}(i) = t_{B}(i) + f,$$

$$t_{F}^{*}(j) = \frac{p_{F}(j)t_{F}(j) + f * [p_{F}(j') - p_{F}(j)] + t_{F}^{*}(j')[p_{F}(j') + (c - 1)p_{D}] - p_{F}(j')t_{F}(j')}{p_{F}(j) - (c - 1)p_{D}},$$

$$t_{D}^{*} = t_{D} + e,$$

and

$$t_A^*(i) = \frac{p_A(i)t_A(i) + p_B(i)\left\{ [t_B(i) - t_D](1 - c) + [f + t_F^*(j')][1 - c - \frac{p_F(j')}{p_D}] + \frac{p_F(j')t_F(j')}{p_D} + ce \right\} - e}{1 - cp_B(i)}$$

As well as the bounds for *c* in Table 3. Then,

$$\begin{aligned} p_A^*(i) p_D^*[t_A^*(i) t_D^*] \\ &= [1 - p_A^*] p_D^*[p_A^*(i) + t_D^*] \\ &= p_A(i) t_A(i) p_D + p_B(i) \{ [t_B(i) - t_D](1 - c) p_D + [f + t_F^*(j')][(1 - c) p_D - p_F(j')] \\ &+ p_F(j') t_F(j') + cep_D \} - ep_D + (t_D + e)[1 - cp_B(i)] p_D \\ &= p_A(i) p_D[t_D + t_A(i)] + p_B(i) \{ t_B(i)(1 - c) p_D + [f + t_F^*(j')][(1 - c) p_D - p_F(j')] + p_F(j') t_F(j') \} \end{aligned}$$

Meanwhile,

$$\begin{split} p_B^*(i) p_F^*[t_B^*(i) \ t_F^*(j)] \\ &= c p_B(i) \left[\frac{p_F(j)}{c} + \frac{(c-1)p_D}{c} \right] \{ t_B(i) + f \\ &+ \frac{p_F(j)t_F(j) + f[p_F(j') - p_F(j)] + t_F^*(j')[p_F(j') + (c-1)p_D] - p_F(j')t_F(j')}{p_F(j) + (c-1)p_D} \} \\ &= p_B(i) [p_F(j) + (c-1)p_D] * \{ t_B(i) + f \\ &+ \frac{p_F(j)t_F(j) + f[p_F(j') - p_F(j)] + t_F^*(j')[p_F(j') + (c-1)p_D] - p_F(j')t_F(j')}{p_F(j) + (c-1)p_D} \} \\ &= p_B(i) \{ [t_B(i) + f] [p_F(j) + (c-1)p_D] + p_F(j)t_F(j) + f[p_F(j') - p_F(j)] \\ &+ t_F^*(j') [p_F(j') + (c-1)p_D] - p_F(j')t_F(j') \} \\ &= p_B(i) \{ t_B(i) [p_F(j) + (c-1)p_D] + p_F(j)t_F(j) \\ &+ [t_F^*(j') + f] [p_F(j') + (c-1)p_D] - p_F(j')t_F(j') \} \end{split}$$

Putting these two parts together, we get

 $p_{A}^{*}(i)p_{D}^{*}[t_{A}^{*}(i)+t_{D}^{*}]+p_{B}^{*}(i)p_{F}^{*}(j)[t_{B}^{*}(i)+t_{F}^{*}(j)]$ $= p_A(i)p_D[t_D + t_A(i)] + p_B(i)\{t_B(i)(1-c)p_D + [f + t_F^*(j')][(1-c)p_D - p_F(j')] + p_F(j')t_F(j')\}$ $+p_B(i)\{t_B(i)[p_F(j) + (c-1)p_D] + p_F(j)t_F(j) + [t_F^*(j') + f][p_F(j') + (c-1)p_D] - p_F(j')t_F(j')\}$ $= p_A(i)p_D[t_D + t_A(i)] + p_B(i)p_F(j)[t_B(i) + t_F(j)]$ = p(i, j)t(i, j)

2.2.3. Remarks on Nonnegative Measure Values

In some applications the measure associated with an arc may be positive or negative. For example, the measure in a decision tree is a payoff, which could be positive (a gain) or negative (a loss). For application to response time, we assume the measure associated with an arc is a time, a nonnegative quantity. In such an application, an admissible transformation of a measure must transform a nonnegative quantity to another nonnegative quantity. The bounds in Table 4 on measure scaling parameters achieve this.

Table 4. Bounds on Scaling Parameters in Admissible Transformations For the Standard Binary Tree for Ordered Processes For Measure Parameters to be Nonnegative.

$\max\{-t_B(i)\} \le f \le \min\{t_B^*(i)\}$	
$\max\{-t_D\} \le e \le \min\{t_{DD}^*\}$	
$\max\{-t_C\} \le k \le \min\{t_C^*\}$	

From Table 2, consider the admissible transformation $t_B^*(i) = t_B(i) + f$. For both $t_B(i)$ and $t_B^*(i)$ to be nonnegative, *f* has boundaries as

$$t_B^*(i) = t_B(i) + f \ge 0$$

 $t_B(i) = t_B^*(i) - f \ge 0$

So for *f*, there is $\max\{-t_B(i)\} \le f \le \min\{t_B^*(i)\}.$

From Table 2, consider the admissible transformation $t^*_D = t_D + e$. For both t^*_D and t_D to be nonnegative, e has boundaries as

$$t_D^* = t_D + e \ge 0$$
$$t_D = t_D^* - e \ge 0$$

·*

So for *f*, there is $\max\{-t_D\} \le e \le \min\{t_D^*\}$.

2.3. Degrees of Freedom

It is useful to know the degrees of freedom. When designing an experiment to test the model, the number of levels of the factors must be large enough to provide at least one degree of freedom; otherwise parameter values can always be found that fit the data. After an experiment is done, degrees of freedom are needed to evaluation of a goodness of fit statistic such as G². Using the admissible transformations of parameters allows us to calculate the degrees of freedom for the Standard Binary Tree for Ordered processes in the following corollary.

Corollary 1. Suppose probability matrix $\mathbf{P} = (p(i, j))$, correct-response-measure matrix $\mathbf{T} = (t(i, j))$ and incorrect-response-measure-matrix $\mathbf{T}_{\mathbf{w}} = (tw(i, j))$ are produced by Factors Φ and Ψ selectively influencing two vertices ordered by the factors in the Standard Binary Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , with probability parameters $p_A(i)$, $p_B(i)$, p_C , p_D , $p_E(j)$, and $p_F(j)$, and measure parameters $t_A(i)$, $t_B(i)$, t_C , t_D , $t_E(j)$, and $t_F(j)$.

Suppose Factor Φ has I levels and Factor Ψ has J levels. Then the degrees of freedom are 3IJ - 3I - 3J + 3.

Proof. For each combination of a level *i* of Factor Φ and a level *j* of Factor Ψ we have an observed probability of a correct response, an observed measure for a correct response and an observed measure for an incorrect response. The probability of an incorrect response is determined by the probability of a correct response. Hence the total number of observations is *3IJ*. The arc probabilities to be estimated are $p_A(i)$, $p_B(i)$, p_C , p_D , $p_E(j)$, and $p_F(j)$. But for every *i*

and for every *j*,

 $p_E(j) + p_F(j) = 1.$

 $p_A(i) + p_B(i) = 1$

Further,

$$p_C + p_D = 1$$

Hence the number of independent arc probabilities to be estimated is I + J + 1. The arc measure parameters to be estimated are $t_A(i)$, $t_B(i)$, t_C , t_D , $t_E(j)$, and $t_F(j)$; their number is 2I + 2 + 2J. The scaling parameters to be freely selected are *c* for arc probabilities, *e*, *f*, and *k* for measures, also $t^*_E(j')$ and $t^*_F(j')$ so there are 6 scaling parameters. Hence, there are 3IJ - (I + J + 1) - (2 + 2I + 2J) + 6 = 3(IJ - I - J + 1) = 3(I - 1)(J - 1) degrees of freedom. A large experiment is not needed to test the MPT; a design with I = 2 levels of one factor and J = 2 levels of the other factor produces 3 degrees of freedom.

3. Discussion

Multinomial Processing Trees are widely used as models of phenomena in psychology. One reason is straightforwardness and relative simplicity, but the major reason is their ability to often account for data. Agreement with data is usually evaluated by goodness of fit. Additional support is sometimes provided by a factorial experiment, with tests of whether factors selectively influence vertices in an MPT (reviewed by [20]).

In a binary MPT exactly two arcs descend from each nonterminal vertex. MPTs with more than two arcs descending from a vertex are occasionally inferred from data. But many MPTs currently in use are binary; most in the applications discussed in the reviews cited above are binary trees.

Here we provide necessary and sufficient conditions, which can be tested by bootstrapping with data from factorial experiments, for selective influence of the factors on vertices in a particular binary MPT, the Standard Binary Tree for Ordered Processes. This MPT has a special role, because under certain conditions if each of two factors selectively influences a different vertex in an arbitrary MPT, that MPT is equivalent for the factors to this one. Methods of testing are beyond the scope of this paper, but described in [12].

19 of 20

Parameter values are not unique. Admissible transformations are given that allow one set of parameter values to be transformed to another. Degrees of freedom for an experiment with the two factors are calculated from the number of observations and the number of parameters to be estimated, both depend on the number of levels of the factors.

Goodness of fit test are general. If a goodness of fit test fails, one learns from the failure about data points that are far from their predicted values. From this information, it may be hard to diagnose what went wrong. Conditions to be tested here are focused, each focusing on a particular parameter. The first two are about probability parameters, the last is about time (or another measure). If the first two hold but the third does not, an investigator leans that response probability is accounted for by the probability parameters, but something is wrong with the way time (or the other measure) is incorporated in the model. Tests of conditions developed here are intended to be more diagnostic than tests of overall goodness of fit.

Author Contributions: Conceptualization, R.S.; formal analysis, R.S. and X.Z.; writing—original draft preparation, R.S. and X.Z.; writing—review and editing, R.S. and X.Z.; funding acquisition, R.S. and X.Z. All authors have read and agreed to the published version of the manuscript.

Funding: This research was funded by a scholarship from China Scholarship Council (CSC) to Zheng and by AFOSR, United States Grant FA9550-09-01-0252 to Dzhafarov and Schweickert.

Conflicts of Interest: The authors declare no conflict of interest.

Appendix A

Theorem 5 of [22].

Suppose for all *i* and *j*, 0 < p(i, j) < 1. Probability matrix $\mathbf{P} = (p(i, j))$, correct-response measure matrix $\mathbf{T} = (t(i, j))$, and incorrect-response-measure matrix $\mathbf{T}_{\mathbf{w}} = (t_w(i, j))$ are produced by Factor Φ and Factor Ψ selectively influencing two different vertices ordered by the factors in the Standard Tree for Ordered Processes, with the vertex selectively influenced by Factor Φ preceding the vertex selectively influenced by Factor Ψ , if and only if there is a level *n* of Factor Ψ and for every level *i* of Factor Φ there are numbers $r_i \ge 0$ and s_i such that the following three conditions are true.

1. The columns of **P** can be numbered so $j \ge j'$ implies that for every $i, p(i, j) \ge p(i, j')$.

2. There are levels *i** and *j** such that for every *i* and *j*,

$$p(i, j) - p(i, j^*) = r_i[p(i^*, j) - p(i^*, j^*)].$$

3. Let $max\{r_i\} = r_h$. For every *j*,

$$r_h r_i s_i [p(h, j) - p(h, n)]$$

= $r_h [p(i, j)t(i, j) - p(i, n)t(i, n)] - r_i [p(h, j)t(h, j) - p(h, n)t(h, n)]$
= $-r_h \{ [1 - p(i, j)]t_w(i, j) - [1 - p(i, n)]t_w(i, n) \} + r_i [1 - p(h, j)]t_w(h, j) - [1 - p(h, n)]t_w(h, n) \}.$

References

- Batchelder, W.H.; Riefer, D.M. Theoretical and empirical review of multinomial process tree modeling. *Psychon. Bull. Rev.* 1999, 6, 57–86. [CrossRef]
- Calanchini, J.; Rivers, A.M.; Klauer, K.C. Multinomial processing trees as theoretical bridges between cognitive and social psychology. *Psychol. Learn. Motiv.* 2018, 63, 39–64.
- Erdfelder, E.; Auer, T.S.; Hilbig, B.E.; Aβfalg, A.; Moshagen, M.; Nadarevic, L. Multinomial processing tree models. Z. Psychol. J. Psychol. 2009, 217, 108–124. [CrossRef]
- 4. Hütter, M.; Klauer, K.C. Applying processing trees in social psychology. Eur. Rev. Soc. Psychol. 2016, 27, 116–159. [CrossRef]
- 5. Jacoby, J.L. A process dissociation framework: Separating automatic from intentional uses of memory. *J. Mem. Lang.* **1991**, *30*, 513–541. [CrossRef]
- 6. Conrey, F.R.; Sherman, J.W.; Gawronski, B.; Hugenberg, K. Separating multiple processes in implicit social cognition: The Quad Model of implicit task performance. *J. Personal. Soc. Psychol.* **2005**, *89*, 469–487. [CrossRef]
- 7. Hu, X.; Batchelder, W.H. The statistical analysis of general processing tree models with the EM algorithm. *Psychometrika* **1994**, *59*, 21–47. [CrossRef]

- 8. Knapp, B.R.; Batchelder, W.H. Representing parametric order constraints in multi-trial applications of multinomial processing tree models. *J. Math. Psychol.* **2004**, *48*, 215–229. [CrossRef]
- Purdy, B.P.; Batchelder, W.H. A context-free language for binary multinomial processing tree models. J. Math. Psychol. 2009, 53, 547–561. [CrossRef]
- 10. Pitt, M.A.; Myung, I.J.; Zhang, S. Toward a method of selecting among computational models of cognition. *Psychol. Rev.* 2002, 109, 472–491. [CrossRef]
- 11. Roberts, S.; Pashler, H. How persuasive is a good fit? A comment on theory testing. *Psychol. Rev.* 2000, *107*, 358–367. [CrossRef] [PubMed]
- 12. Schweickert, R.; Zheng, X. Tree Inference: Selective influence in multinomial processing trees with supplementary measures such as response time. *J. Math. Psychol.* **2018**, *86*, 10–29. [CrossRef]
- 13. Brainerd, C.J.; Nakamura, K.; Lee, W.F. Recollection is fast and slow. J. Exp. Psychol. Learn. Mem. Cogn. 2019, 45, 302–319. [CrossRef] [PubMed]
- 14. Heck, D.W.; Erdfelder, E. Extending multinomial processing tree models to measure the relative speed of cognitive processes. *Psychon. Bull. Rev.* **2016**, *23*, 1440–1465. [CrossRef]
- 15. Hu, X. Extending General Processing Tree models to analyze reaction time experiments. *J. Math. Psychol.* **2001**, *45*, 603–634. [CrossRef] [PubMed]
- 16. Klauer, K.C.; Kellen, D. RT-MPTs: Process models for response-time distributions based on multinomial processing trees with applications to recognition memory. *J. Math. Psychol.* **2018**, *82*, 111–130. [CrossRef]
- 17. Link, S.W. Correcting response measures for guessing and partial information. Psychol. Bull. 1982, 92, 469–486. [CrossRef]
- Wollschläger, L.M.; Diederich, A. The 2N-ary choice tree model for N alternative preferential choice. *Front. Psychol.* 2012, 3, 189.
 [CrossRef]
- 19. Brainerd, C.J.; Gomes, C.F.A.; Nakamura, K. Dual recollection in episodic memory. J. Exp. Psychol. Gen. 2015, 144, 816–843. [CrossRef]
- 20. Schweickert, R.; Fisher, D.L.; Sung, K. Discovering Cognitive Architecture by Selectively Influencing Mental Processes; World Scientific: Singapore, 2012.
- Schweickert, R.; Zheng, X. Tree Inference: Uniqueness of Multinomial Processing Trees representing response time when two factors selectively influence processes. J. Math. Psychol. 2019, 88, 58–77. [CrossRef]
- Schweickert, R.; Zheng, X. Tree Inference: Response time in Multinomial Processing Trees, representation and uniqueness of parameters. J. Math. Psychol. 2019, 91, 51–69. [CrossRef]
- Schweickert, R.; Chen, S. Tree inference with factors selectively influencing processes in a processing tree. J. Math. Psychol. 2008, 52, 158–183. [CrossRef]