

Article

Empirical Model of Human Thermal Comfort in Subtropical Climates: A First Approach to the Brazilian Subtropical Index (BSI)

João Paulo Assis Gobo^{1,*}, Marlon Resende Faria², Emerson Galvani¹, Fabio Luiz Teixeira Goncalves³ and Leonardo Marques Monteiro⁴

- ¹ Faculty of Philosophy, Literature and Humanities (FFLCH), University of São Paulo (USP), São Paulo 05508-080, Brazil; egalvani@usp.br
- ² Institute of Physics (IF), University of São Paulo (USP), São Paulo 05508-090, Brazil; marlon.faria@usp.br
- ³ Institute of Astronomy, Geophysics and Atmospheric Sciences (IAG), University of São Paulo (USP), São Paulo 05508-090, Brazil; fabio.goncalves@iag.usp.br
- ⁴ School of Architecture and Urbanism (FAU), University of São Paulo (USP), São Paulo 05508-080, Brazil; leo4mm@usp.br
- * Correspondence: jpgobo@usp.br; Tel.: +55-11-999664168

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Abstract: The bioclimatic well-being of individuals is associated with the environmental characteristics of where they live. Knowing the relationships between local and regional climatic variables as well as the physical characteristics of a given region and their implications on thermal comfort is important for identifying aspects of thermal sensation in the population. The aim of this study is to develop an empirical model of human thermal comfort based on subjective and individual environmental patterns observed in the city of Santa Maria, located in the state of Rio Grande do Sul, Brazil (Subtropical climate). Meteorological data were collected by means of an automatic meteorological station installed in the city center, which contained sensors measuring global solar radiation, air temperature, globe temperature (via a grey globe thermometer), relative humidity and wind speed and direction. A total of 1720 people were also interviewed using a questionnaire adapted from the model recommended by ISO 10551. Linear regressions were performed to obtain the predictive model. The observed results proposed a new empirical model for subtropical climate, the Brazilian Subtropical Index (BSI), which was verified to be more than 79% accurate, with a coefficient of determination of 0.926 and an adjusted R^2 value of 0.924.

Keywords: thermal comfort index; linear regression; predictive model

1. Introduction

Human thermal comfort ranges have been established or adapted in the last few years from a thermal sensitivity scale of seven or nine points for the evaluation of the average perception of people regarding weather conditions in open spaces [1–6]. New ways to predict the thermal sensations of people in their typical environments, based on personal, environmental and physiological variables, have been explored for decades [7]. Accordingly, mathematical models that predict the thermal responses of individuals in their natural environments have been developed [7–10].

These studies have been developed for the determination of thermal comfort in open spaces under uncontrolled thermal conditions [2–4,11,12] and use modelling and evaluation methods from a thermophysiological perspective. Examples include the studies of Gulyás et al. [13] and Höppe [1], while others are based on the perspective of the relationships between the climatic parameters that determine the level of thermal comfort of humans in open spaces [14,15].



To assess the thermal comfort and sensation patterns of the population during the different seasons of the year, for temperate climates [5,6,16–19], or hot and humid climates [20–23], many studies have attempted to predict comfort conditions by calibrating the interpretive ranges of different indices to the climatic and thermal perception conditions of certain places.

Studies show differences in outdoor thermal comfort between distinct climatic zones [24,25] and suggest a need for additional field surveys on subjective human perception in those environments [22].

However, the calibration of the thermal comfort ranges of a given index does not always answer all of the questions raised about the thermal comfort of a given location. Thus, the determination of a predictive model based on the environmental, physiological, and subjective aspects of the reference environment offers a solution for better interpretations of the biometeorological aspects of the place of study [24].

Several indices used to evaluate outdoor thermal comfort were originally developed for indoor spaces [26,27]. Some examples are the Standard Effective Temperature (SET) index [28], which was adapted as OUT_SET [29] for outdoor environments, and the Predicted Mean Vote (PMV) [30], the use of which is suggested by [31] and by [32], adapted for outdoor environments and taking into account the influence of shortwave radiation.

Thus, according to Salata et al. [33], one of the main solutions found for considering the influence of personal adaptions and expectations of the population of a certain place on thermal sensation is the proposal of an empirical index based on multiple regression analyses [8–10,33]. The correlation of multiple variables is commonly performed by means of linear regression similar to many previous studies [6,8–10,34–36], and the resulting groups of indices define human comfort as a function of the thermal environment. These indices are suitable for long term studies such as historical bioclimatic analyses [37].

Empirical indices accurately describe the thermal sensations of pedestrians and the environmental factors that most affect their thermal behavior [38]. These indices are adequate for the cities in which they were analyzed, whereas the simplified models are suitable for large-scale studies in which urban microclimates can be neglected [37,38].

A wide variety of studies have been conducted in temperate and subtropical countries [5,6,16–18,24,39–42], each of which are characterized by specific climatic and cultural conditions, using a cross-sectional approach. However, there is a lack of research on thermal comfort in subtropical climates, which are characteristic of the geographic areas South of the Tropic of Capricorn and North of the Tropic of Cancer [43,44], especially in the Brazilian subtropical climate zone.

The present study proposes a definition for a new simplified empirical index based on cross-sectional interviews and on meteorological data, including air temperature, relative air humidity, and wind speed, observed simultaneously to the interviews through multiple linear regressions in the city of Santa Maria, state of Rio Grande do Sul, Brazil. This is a subtropical climate region, classified as Cfa according to Alvares et al. [44] and Koppen [45]. However, the objective of this study is not to propose a global index with a wide application, but to pave the way for future studies that promote the improvement and adjustment of this model in order to meet a wider application demand in subtropical regions.

2. Experiments

The meteorological and thermal sensation data were collected during 2015 and 2016 in the city of Santa Maria, Rio Grande do Sul, Brazil, which is located in the geographic center of the state and has an estimated population of 278,445 inhabitants [46]. According to Köppen [45], in terms of general climatic classification, the city is classified as type Cfa, with a hot and rainy temperate climate, no dry season, and a hot summer. The hottest month has a mean temperature higher than 22 °C and a mean air temperature in the hottest four months above 10 °C, and the coldest month has a mean temperature above 3 °C.

For this purpose, a Campbell CR100 Automatic Weather Station (AWS) was used, at a maximum height of 2.0 m, with a mobile aluminum tripod containing the following sensors (Table 1):

Sensor Model	Resolution
Rain gage model TE525 Tipping Bucket Rain Gage	0.10 mm
Global radiation sensor model LI200X Pyranometer	$0.2 \text{ kW} \cdot \text{m}^{-2}$
Temperature and humidity sensor model HMP35C Temperature and Relative Humidity Probe	0.1 °C and 0.6%
Speed and wind direction sensor models 03001 R.M. Young Wind Setry Set 03101 R.M. Young Wind Sentry Anemometer 03301 R. M. Young Wind Sentry Vane	$0.5\ m{\cdot}s^{-1}$ and 5°
TMCx-HD gray globe thermometer	0.03 °C

Table	 Sensor 	model and	resol	lution
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Primary data on air temperature, gray globe temperature—because the station was set up in an open space exposed to direct solar radiation [47]—relative humidity, wind speed, wind gusts, global solar radiation, and rain were collected. The station was positioned in a paved area in the center of Santa Maria (Figure 1).

Field data collection was performed during the following periods: 5–7 August 2015, 17–19 January 2016, and 6–8 July 2016. These periods were selected as representative of summer (January 2016) and winter (July 2016). The period of August 2015 was used for the validation of the data.

The choice of the representative periods of summer and winter was aimed at demonstrating the seasonality of the data and evaluating its influence on the final results. The objective is to incorporate the region's climatic amplitudes into the proposed models, because methodological and operational evaluations are not necessarily possible over the course of the entire year. The seasonal aspect is incorporated into the model to make observations in representative periods of summer and winter. Collection of meteorological data and interviews with the local population were performed from 9:00 a.m. to 5:00 p.m. local time during the field experiments described above. The model can also be applied to the night period, simply by including the input variables of this period. However, it is understood that the flow of people in this environmental profile is greater in the daytime and in the early hours of dusk.



(**a**)



Figure 1. Location of the study area (a) and of the automatic weather station (b).

On the days of field research, the 5–7 August 2015, period was characterized by a persistent high-pressure system anomaly (anticyclone positioned at approximately 30 degrees of latitude), with relatively slow displacement of high pressures, which persisted for several days. During the January 2016 field research days, it was possible to identify a pattern compatible with Santa Maria's normal climatological averages for that month, which presented high temperatures with maxima above 32 °C. In the analysis of the next winter, during July 2016, above-average temperatures for Santa Maria were observed mainly in the first day of analysis, but after that day, temperatures were within the range expected for that season.

The mean temperature and relative humidity patterns observed in all periods of data collection are shown in Figure 2. The patterns of average velocity and maximum wind speed are observed in Figure 3 and the predominant direction of wind for the period of analysis is shown in Figure 4.



Figure 2. Mean air temperature and relative air humidity for the period of analysis.



Figure 3. Patterns of average velocity and maximum wind speed for the period of analysis.



Figure 4. Predominant direction of wind for the period of analysis.

The mean patterns of the main climatic attributes observed in the field are shown in Table 2.

Days	Average Temperature (°C)	Minimum Temperature (°C)	Maximum Temperature (°C)	Average Relative Humidity (%)	Minimum Relative Humidity (%)	Maximum Relative Humidity (%)	Average Wind Speed (m/s)	Maximum Wind Speed (m/s)
5 August 2015	22.9	18.0	25.7	73.1	65.0	88.2	0.0	1.1
6 August 2015	28.4	24.4	30.7	37.6	32.0	47.2	1.4	8.8
7 August 2015	29.3	26.4	31.4	34.4	30.9	38.2	1.3	8.3
17 January 2016	30.0	21.1	35.2	52.7	23.9	86.7	0.1	2.5
18 January 2016	31.5	23.6	36.6	36.3	11.4	59.8	0.1	2.3
19 January 2016	30.5	23.5	36.7	46.9	23.6	70.8	0.1	2.2
6 July 2016	10.5	8.1	12.9	79.7	70.3	89.8	0.5	3.1
7 July 2016	16.0	11.4	18.4	51.2	38.5	74.3	0.2	2.3
8 July 2016	17.2	11.3	19.7	52.9	44.1	71.9	0.0	0.7

Table 2. Average values of climatic attributes collected during the periods of analysis.

In the present study, only those individuals who had resided in the town for more than one year were interviewed to derive a function of the individual thermal history and environmental memory, as observed by Nikolopoulou et al. [48], in a total of 1720 interviews (Table 3). Interviewees also had to exhibit 0.3 to 1.5 clo of clothing insulation, which corresponds to wearing jeans and T-shirt or a suit [49], and 300 W of physical activity, because only people in motion (walking) were included [50]. The questionnaire used was an adaptation of the one included in the standard ISO 10551 (1995) (Figure 5).

Days of Analysis	Gender	Ν	Average Age (Years)	Average Weight (Kg)	Average Height (m)
5 August 2015	F	159	33	62.90	1.630
5 August 2015	М	128	35	78.00	1.745
Total		287	34	69.63	1.681
6 August 2015	F	146	31	62.97	1.639
0 August 2015	М	136	34	76.75	1.742
Total		282	32	69.61	1.688
7 August 2015	F	149	31	63.30	1.641
7 August 2015	М	139	32	77.23	1.755
Total		288	32	70.02	1.696
17 January 2016	F	63	30	65.00	1.640
17 January 2010	М	81	37	75.20	1.722
Total		144	34	70.74	1.686
18 January 2016	F	81	34	64.64	1.642
10 January 2010	М	63	33	81.30	1.766
Total		144	33	71.93	1.696
10 January 2016	F	85	30	64.00	1.644
19 January 2016	Μ	59	33	79.20	1.743
Total		144	31	70.23	1.685
6 July 2016	F	72	35	65.54	1.618
0 July 2010	М	72	32	77.10	1.745
Total		144	34	71.32	1.681
7 July 2016	F	89	33	69.70	1.641
7 July 2010	М	54	33	79.61	1.751
Total		143	33	73.44	1.683
8 July 2016	F	92	33	68.12	1.648
6 July 2010	М	52	36	78.52	1.747
Total		144	35	71.88	1.684
Grand Tota	1	1720	33	70.68	1.687

Table 3. Description of the analyzed population.

Data of the interviewee: age (___) Sex: (___) M (___) F Weight (___) Height (___) With regard to the individual's dress, he is wearing: 0,3 clo (___); 0,5 clo (___); 0,8 clo (___); 1,0 clo (___); 1,5 clo (___) 1. At this very moment, I'm feeling: 4. With regard to air temperature, I would prefer it to be: () Cold -3 () Cool -2 () lowest -1 () Slightly cool -1 () as is 0 () Neither cold nor hot 0 () highest 1 () I do not know how to say X. () Slightly warm 1 () Warm 2 () Hot 3 5. With regard to air humidity, I would prefer the air to be: () drier -1 2. At this very moment, regarding the () as is 0 weather, I am: () wettest 1 () comfortable 0 () I do not know how to say X. () a little uncomfortable 1 6. Regarding the wind, I would prefer it to () uncomfortable 2 () very uncomfortable 3 be: () weaker -1 () as is 0 3. Right now, I'd rather be feeling: () stronger 1 () Much colder -3 () I do not know how to say X. () Colder -2 7. With regard to solar radiation, I would () A little more cold -1 () No changes 0 prefer it to be: () A little more heat 1 () softer -1 () More heat 2 () as is 0

() Much more heat 3

() more intense 1 () I can not say X.

Figure 5. The questionnaire adapted from ISO 10551 (1995).

From the correlation between the environmental variables (air temperature, relative air humidity and wind speed) observed in the field and the subjective answers, a new empirical index was proposed. Only the data referring to the collection days in January and July 2016 were used because the August 2015 data were later used to validate the results of the index by means of an uncertainty test for the samples and a Student's *t*-test.

All statistical procedures were performed using the statistical software R, and the multiple linear regression method is described by the following mathematical expression:

$$Y = \alpha_1 \cdot x_1 + \alpha_2 \cdot x_2 + \alpha_3 \cdot x_3 + \dots + \alpha_n \cdot x_n + c \tag{1}$$

where *n* is the number of terms in the equation, α_n are the parameters (constants) obtained by the regression, *c* is the constant that corresponds to the point where the adjusted line intersects the y-axis, *Y* is the dependent variable, and x_n are the independent variables.

In the present study, what we call the mean index is, in fact, the mean of the responses observed in the field for a given climatic variable value. Since *Y* is a random discrete variable whose possible values are finite, the best way to estimate the expected mean index value in a given time interval is to use the formula of the expected value for discrete random variables, or a finite case, since the responses are limited to the [32] 7-point class interval:

$$E[y] = \sum_{k=1}^{n} y_k p_k \tag{2}$$

where y_k is the possible value, and p_k is its respective probability in some independent tests.

In the case of the present study,

$$I_{mean} = \sum_{n=1}^{n} I_k p_k \tag{3}$$

where I_k is the mean index, and p_k is the probability estimated by dividing the number of responses to the index by the total number of interviews performed in a given interval.

Subsequently, more than 100 regressions were performed with multiple variables to obtain the appropriate parameters for the various situations examined. Regressions were also tested using absolute humidity rather than relative humidity. This is because collinearity was observed in the results, and the variance inflation factor (VIF) test was used, which, according to Marquardt and Marquardt and Snee [51,52], is used to determine the variables with VIF values exceeding 10, which should be excluded to prevent compromising the model.

Finally, the index was validated with the data obtained in the first field survey conducted in August 2015, and through the uncertainty test for the samples,

$$\sigma_I = \sum_{n=1}^{n} (y_k - I)^2 p_k$$
(4)

where σ_i is the uncertainty of the mean of the votes obtained through interviews, y_k is the k index (k = -3 to 3), and p_k is the estimated probability of obtaining the y_k index.

After the uncertainty calculation, the *t*-test or Student's test was performed as given by the following equation:

$$T = \frac{I_{mean} - I_{model}}{\frac{\sigma}{\sqrt{n}}}$$
(5)

where I_{mean} is the comfort index obtained by the mean of the votes obtained in the interviews, I_{model} is the index calculated by the model, σ is the uncertainty of the mean of the votes obtained through interviews, and n is the number of samples (number of interviews to obtain the calculated mean).

The multiple linear regression method is simple and robust, since the linearity assumptions of the data, such as normality (obtained by the measure of obliquity and kurtosis of the distribution) were verified, which were, finally, executed in the process method validation [53].

In order to obtain an even more efficient validation for the developed model, 33% of the data were used for the comparison between this and the already traditional models in the literature, including Physiologically Equivalent Temperature (PET) [54], Standard Effective Temperature (SET) *, and Predicted Mean Vote (PMV) [30], which had their classes adjusted by Gobo, Galvani and Wollmann [55] for the same climatic situation in Santa Maria.

3. Results

We chose to develop a simplified thermal comfort index for winter and summer conditions based on the independent variables including air temperature, relative air humidity, and wind speed. These variables are easily measurable and available in databases from meteorological institutes.

The assumptions of the linear models, as if the predictors are normally distributed and have equal variances, are verified by Asymmetry, Curtosis, and Heteroscedasticity measurements obtained by the GVLMA (Global Validation of Linear Model Assumption) package of the R programming language (Table 4).

Table 4. Validation test of the method.

Test	Value	<i>p</i> -Value
Asymmetry	0.09089	0.76305
Curtosis	0.59938	0.43881
Heteroscedasticity	1.46280	0.22648

The correlation between logistic and linear regression models was \pm 98%, proving unnecessary the use of two different methods, and therefore, the multiple linear regression model was used. The values of Table 3 show acceptable results, which increase the confidence in the model.

The results revealed a strong positive correlation between the means of the thermal sensation votes and air temperature, with an R^2 value equal to 0.96 (Figure 6).



Figure 6. Correlation between thermal sensation vote and air temperature.

The correlation analysis of air temperature with the other variables indicated a strong negative correlation with relative air humidity for both July and January. However, the strong correlation between the variables air temperature and relative air humidity may impair the model results due to internal collinearity.

According to Monteiro et al. [56], when the relative air humidity has a significant negative correlation with air temperature, direct consideration of this information may lead to the false interpretation that higher relative humidity leads to more intense cold sensations. However, relative

humidity has a strong negative correlation with air temperature, approaching 1 in the most restricted data set observed by the author. Taking into account the fact that the absolute humidity is more or less constant during a given period, higher air temperature results in lower relative humidity. Thus, the correlation obtained for air humidity is largely due to the variation in air temperature [56].

Monteiro et al. [56] add that when considering a larger data set that includes hotter and colder thermal conditions, a longer series of days will have different absolute humidities, leading to different correlations between air temperature and relative humidity. Nevertheless, relative humidity is dependent on air temperature, and therefore, as is observed in the subsequent analyses, it is necessary to consider the absolute humidity for effective testing of the effects of humidity on the subjective thermal sensation responses.

Thus, a new regression was performed for the same parameters, but with the variable absolute humidity (*g* of water vapor per m³ of air) replacing relative humidity. However, the results obtained for the absolute humidity variable did not show a considerable statistical improvement, as shown in Table 5, where the significance value of the variable absolute humidity decreased relative to relative humidity, becoming 0.395 (Table 5) compared to the significance level of 0.711 for relative humidity (Table 6).

	Coefficient	Standard Error	t	<i>p</i> -Value	
Constant	-2.78739	0.12286	-22.688	$<\!\!2 \times 10^{-16}$	
Air temperature	0.15143	0.00484	31.290	$<\!\!2 imes 10^{-16}$	
Absolute humidity	0.01018	0.01192	0.854	0.395	
Wind speed	-0.41469	0.16482	-2.516	0.013	
Residual standard error		0.3889			
Multiple R ²	0.9264				
Adjusted R ²	0.9248				
F-Statistic	587.5				
Level of significance— <i>p</i>	$<\!\!2.2 imes 10^{-16}$				

Table 5. Results of the regression performed with the variable absolute air humidity.

Thus, the VIF was calculated for the previously selected variables. As stated by Marquardt [52], variables with VIF values greater than 10 should be excluded; however, the collinear variables do not add any relevant value to the model (Table 5), and values above the threshold established by Marquardt [52] were not observed, so it was decided to keep relative humidity in the construction of the index.

The relative air humidity and the means of thermal sensation votes showed a strong negative correlation, with an R^2 value of 0.736 (Figure 7). However, the relative humidity had a weak positive correlation with wind speed, with an R^2 value of 0.209. Wind speed had a weak correlation with all the other variables, with the correlation between wind and mean thermal sensation votes being negative, which was also the case between wind speed and air temperature (Figure 8).

Correlation between thermal sensation vote and relative humidity



Figure 7. Correlation between the pedestrians' thermal sensation votes and mean relative air humidity in the two analysis periods (January and July 2016).



Figure 8. Correlation between the pedestrians' thermal sensation votes and mean wind speed in the two analysis periods (January and July 2016).

The probability density function (PDF)—or density of a continuous random variable—was calculated for the means of the thermal sensation votes of the individuals interviewed during the entire study (winter and summer) as well as for mean air temperature, mean relative humidity and mean wind speed for that period. The PDF is a function that describes the relative probability of these random variables when taking a given value.

Table 5 shows the results of the linear regressions performed for the mean values reached for the conditions surveyed during the study (summer and winter) to obtain the model.

	Coefficient	Standard Error	Т	Level of Significance	Variance Inflation Factor (VIF)
Constant	-2.642548	0.269316	-9.812	$<2 \times 10^{-16}$	-
T _{air}	0.152299	0.005831	26.118	$< 2 \times 10^{-16}$	2.437
RH	-0.001043	0.002811	-0.371	0.71119	2.379
S	-0.430040	0.164042	-2.622	0.00972	1.071
Residual standard error				0.389	
Multiple R ²				0.926	
Adjusted R ²				0.924	
F-Statistic				584.9	
	Level of signi	ificance—p		<2.2	

Table 6. Statistical analysis of the constant, three independent variables and variance for the regression performed with the complete series (January and July 2016).

Thus, the model had high multiple R^2 and adjusted R^2 values of 0.926 and 0.924, respectively, whereas the statistical test F showed a high value of 584.9, which confirms that together the variables contribute to the prediction of the independent variable.

Next, an equation was defined that considers the three independent variables (Tair, RH and S) correlated for the situation survey and the mean value of perceived thermal sensation in each (based on the results of the questionnaires applied in January and July 2016):

$$BSI = 0.1523 \times Tair - 0.0010 \times RH - 0.4300 \times S - 2.6425$$
(6)

where BSI = Brazilian Subtropical Index, Tair = air temperature (°C), RH = relative humidity (%), and S = wind speed (m/s).

The Brazilian Subtropical Index (BSI) is the proposed model, which is defined as a thermal sensation scale based on the mean vote of individuals interviewed during winter and summer using the 7-point scale (Table 7).

_		
	≤ -3	Very cold
	-2.9 to -2	Cold
	-1.9 to -1	Slightly cold
	-0.9 to 0.9	Neither cold nor hot
	1 to 1.9	Slightly hot
	2 to 2.9	Hot
	≥ 3	Very hot

Table 7. Interpretive bands for the Brazilian Subtropical Index (BSI).

Validation of the Proposed Empirical Index

The model was validated using a Student's *t*-test to evaluate the relationship between the comfort index predicted by the BSI model (obtained with data from January and July) and the thermal sensation votes of the interviewees. For this purpose, 857 interviews were obtained from the August 2015 survey and compared to the BSI results for the same period.

The results of the validations are shown in Table 8 and Figure 9.

Table 8. Accuracy score by *t*-test (*p*-value = 0.05).

Days Tested	BSI
5 August 2015	0.875
6 August 2015	0.812
7 August 2015	0.791

Correlation between thermal sensation vote and ISB



Figure 9. Correlation between the empirical data from interviews collected in August and the model made by the data obtained in January and July: 0.96.

The hit rate obtained via the validation of the index is 88% for the first test, 81% for the second test and 79% for the third test. These values are higher than those observed by Gobo et al. [54] in a study calibrating the interpretative ranges of PET, SET and PMV indices for Santa Maria, where the authors identified hit rates after the calibration of 69.3% for the PET, 64.9% for the SET and 58.7% for the PMV index.

As observed in the study of Salata et al. [36], the Mediterranean Outdoor Comfort Index (MOCI) presented an adjusted R^2 value of 0.395, an R^2 value of 0.398, and a Pearson coefficient of 0.631, whereas the BSI presented an R^2 value of 0.926 and an adjusted R^2 value of 0.924 and a higher Pearson coefficient, 0.790.

Comparing the results obtained during the development and validation of the BSI with those observed by Ruiz and Correa [8], the adjusted R^2 value of the latter is closer (0.719), with a predictive capacity of 73%, which is very close to that observed in the validation of the BSI.

When comparing the results of the Global Outdoor Comfort Index (GOCI) developed by Golasi et al. [10], an adjusted R^2 value of 0.379 and a Pearson coefficient of 0.616 were observed

for the GOCI, whereas the BSI presented an adjusted R² value of 0.924 and Pearson coefficient higher than 0.790, further indicating the efficiency of the BSI.

The BSI hit rates are higher than 50%, and it should be noted that Nikolopoulou and Steemers [57], when analyzing psychological aspects related to thermal sensation, found that climatic variables have a strong influence on thermal sensation, but these explain approximately 50% of the variation between objective and subjective assessments of comfort.

Of the approximately 165 thermal indices developed to date, only 4 (PET, PMV, Universal Thermal Climate Index (UTCI) and SET *) are widely used in studies of outdoor thermal perception [58]. Thus, a correlation test was performed between the respondents' thermal preference responses in the period used for the BSI validation (August 2015) and the PET, SET * and PMV models calibrated by Gobo, Galvani and Wollmann [54] to Santa Maria. The results are described in Table 9.

Table 9. Correlation between the PET, SET * and PMV models calibrated by Gobo, Galvani and Wollmann (2018) for Santa Maria and the thermal preference of the interviewees.

Indexes	Correlation
Calibrated PET	0.23
Calibrated SET *	0.14
Calibrated PMV	0.38

It is important to note that the correlations made between the PET, SET *, PMV, and the thermal preference responses of the interviewees (Table 9) were only for the BSI validation period of August 2015, differing from the period used by Gobo, Galvani and Wollman [55] in the calibration of the mentioned models, where the whole series of August of 2015, January and July of 2016 was used.

The low efficiency of the PET, SET * and PMV indices for the analyzed period can be explained in part by the large size of the comfort range of these indices as calibrated by Gobo, Galvani and Wollman [54] for the study area, featuring 16 °C–24 °C for PET, 17 °C–23 °C for the SET and –1–0.8 for the PMV. Potchter et al. [58], when analyzing a work done with PET in the Cfa climate [45], signals for the acceptance of the comfort range of 87% of the case studies between 24 °C and 27 °C and up to 94% of the cases between 25 °C–26 °C, which considerably limits the comfort range for these indexes. Therefore, the validation of the BSI presented greater efficiency when compared to the other indexes commonly used in studies appearing in the literature.

4. Conclusions

The index obtained by means of multiple linear regressions presented high statistic values that did not reveal any anomalous behaviors that indicated inadequacy of the chosen model. The multiple linear regression model presented considerable results when compared to the other regression models, which increased confidence in the use of this model. The VIF demonstrated that the collinear variables do not add any relevant value to the model, with no values above the limit established by Marquardt [52] being found.

The proposal of the Brazilian Subtropical Index (BSI) provides a simple and easy-to-apply model with temperature, relative air humidity, and wind velocity as input variables. Therefore, it accounts for meteorological attributes commonly measurable in conventional and automatic surface meteorological stations.

The results expressed the perception of the interviewees regarding comfort and thermal discomfort for the locality of Santa Maria, a subtropical climate region in southern Brazil. A high rate of model accuracy, with a multiple R^2 and adjusted R^2 of 0.926 and 0.924 respectively, and a high statistical F with value of 584.9, confirmed that, together, the modeled variables contributed to the prediction of the independent variable. Thus, like any statistical model, the index proposed here has limitations, and its effectiveness in situations different from those analyzed should be explored based on more detailed studies and with a longer observation times.

Finally, BSI validation has proved to be effective. With hit rates higher than 80%, these values are higher than those observed for the PET, SET and PMV indices for Santa Maria, which indicates the efficiency and reliability of the model.

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Nomenclature

BSI	Brazilian Subtropical Index	PMV	Predicted Mean Vote
SET	Standard Effective Temperature	T _{air}	Air Temperature
AWS	Automatic Weather Station	RH	Relative Humidity
VIF	Variance Inflation Factor	S	Wind Speed
PDF	Probability Density Function	UTCI	Universal Thermal Climate Index
GVLMA	Global Validation of Linear Model Assumption	GOCI	Global Outdoor Comfort Index
PET	Physiologically Equivalent Temperature	MOCI	Mediterranean Outdoor Comfort Index
Cfa	Humid subtropical climate		

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